Statistical Machine Learning
Gaussian Processes

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Outline

BAYESIAN LINEAR REGRESSION REVISITED

- Bayesian linear regression places a (Gaussian) prior over the weights vector, and computes the (Gaussian) posterior distribution over weights.
- What does this mean? Consider linear basis functions. In this case, the regression line is a *random line*, with the property that the output prediction at any point is a Gaussian random variable
- This concept can be generalised: taking linear combinations of basis functions with (Gaussian) random coefficients leads to a (Gaussian) random function

RANDOM FUNCTIONS TERMINOLOGY

- A random function is an infinite collection of random variables indexed by the argument of the function
- A popular alternative name is a *stochastic process*
- When considering the random function evaluated at a (finite) set of points, we get a random vector
- The distribution of this random vector is called *finite dimensional marginal*

IMPORTANT EXERCISE

Let $\phi_0(x), \ldots, \phi_{M-1}(x)$ be a fixed set of functions, and let $f(x) = \sum w_i \phi_i(x)$. If $\mathbf{w} \sim \mathcal{N}(0, I)$, compute:

1. The single-point marginal distribution of $f(x)$
2. The two-point marginal distribution of $f(x_1), f(x_2)$

THE GRAM MATRIX

- Generalising the exercise to more than two points, we get that *any* finite dimensional marginal of this process is multivariate Gaussian
- The covariance matrix of this function is given by evaluating a function of two variables at all possible pairs
- The function is defined by the set of basis functions

$$k(x_i, x_j) = \boldsymbol{\phi}(x_i)^T \boldsymbol{\phi}(x_j)$$

- The covariance matrix is often called *Gram matrix* and is (necessarily) symmetric and positive definite
- Bayesian prediction in regression then is essentially the same as computing conditionals for Gaussians (more later)

MAIN LIMITATION OF BAYESIAN REGRESSION

- Choice of basis functions inevitably impacts what can be predicted
- Suppose one wishes the basis functions to tend to zero as $x \to \infty$
- Then, necessarily, very large input values will have predicted outputs near zero with high confidence!
- Ideally, one would want a prior over functions which would have the same uncertainty everywhere

FUNCTION SPACE VIEW

- In order to construct such priors, one possibility would be to construct a countable sequence of basis functions. We can partition the full $\mathbb{R}^n$ in compact sets, and define a finite number of basis functions supported in each compact set so that the variance in each point of the state space is a constant (partition of unity).
- This approach, called the *weights space view*, is unpractical, but it demonstrates the existence of truly infinite dimensional Gaussian Processes.
- In general, it is more useful to take the dual point of view, and work with kernels rather than with basis functions.

Outline

GP DEFINITION

- A Gaussian Process (GP) is a stochastic process indexed by a continuous variable x s.t. all finite dimensional marginals are multivariate Gaussian
- A GP is uniquely defined by its *mean* and *covariance* functions, denoted by $\mu(x)$ and $k(x, x')$:

$$f \sim \mathcal{GP}(\mu, k) \leftrightarrow \mathbf{f} = (f(x_1), \ldots, f(x_N)) \sim \mathcal{N}(\boldsymbol{\mu}, K),$$

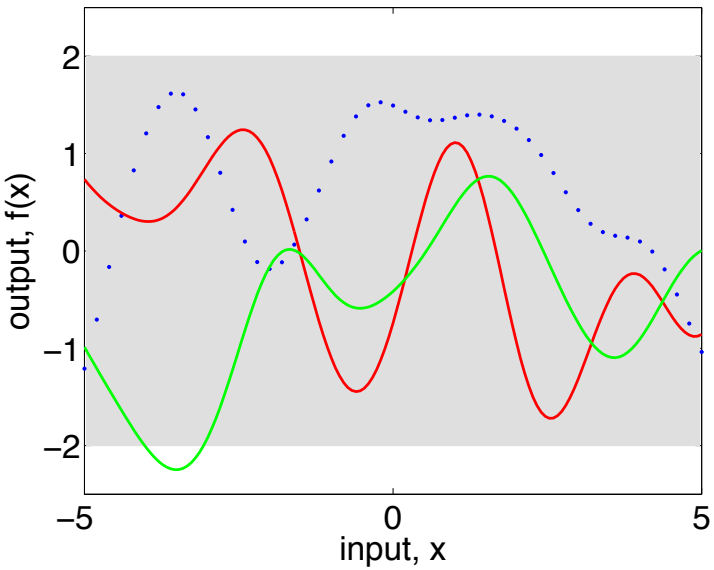
$$\boldsymbol{\mu} = (\mu(x_1), \ldots, \mu(x_N)), \quad K = (k(x_i, x_j))_{i,j}$$

- The covariance function must satisfy some conditions (Mercer's theorem), essentially it needs to evaluate to a symmetric positive definite function for all sets of input points

AN EXAMPLE

Consider a 1-dimensional GP with mean function $\mu(x) \equiv 0$, and with Gaussian covariance function:

$$k(x, x') = \exp\left[-\frac{1}{2}|x - x'|^2\right]$$



The variance at each point x is $k(x, x) = 1$. If we consider a test set $X^* = x_1, \ldots x_n$, then the joint distribution of $\mathbf{f}^* = (f(x_1), \ldots, f(x_n))$ is

$$\mathbf{f}^* \sim \mathcal{N}(\mathbf{0}, K(X^*, X^*))$$

where $K(X^*, X^*)$ is the Gram matrix, $K_{ij} = k(x_i, x_j)$, which is symmetric and positive definite.

NOISE-FREE PREDICTION

- Suppose now to observe the exact value of the GP at N different points, $X = x_1, \ldots, x_N$, with observations $\mathbf{f} = f(x_1), \ldots, f(x_N)$.
- Consider also the test points $X^* = x_1, \ldots x_n$, with function values $\mathbf{f}^* = (f(x_1), \ldots, f(x_n))$ (unobserved, to be estimated).
- The joint prior distribution of f on inputs X and test points X^* is

$$\begin{bmatrix} \mathbf{f} \\ \mathbf{f}_* \end{bmatrix} \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} K(X,X) & K(X,X_*) \\ K(X_*,X) & K(X_*,X_*) \end{bmatrix}\right). \tag{2.18}$$

- If we observe the values at X, then we need to condition on these values. Hence the conditional $\mathbf{f}^*|\mathbf{f}$ is

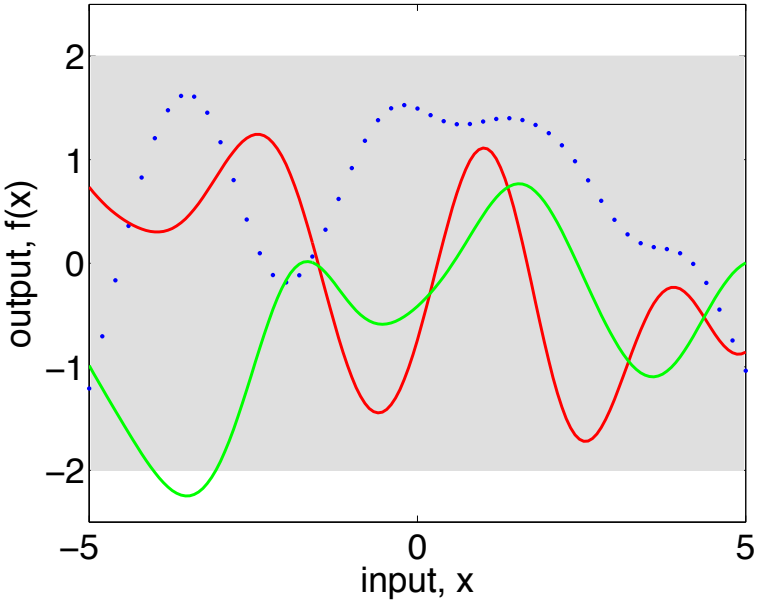
$$\begin{aligned} \mathbf{f}_*|X_*, X, \mathbf{f} \sim \mathcal{N}\big(&K(X_*,X)K(X,X)^{-1}\mathbf{f}, \\ &K(X_*,X_*) - K(X_*,X)K(X,X)^{-1}K(X,X_*)\big). \end{aligned} \tag{2.19}$$

which is obtained by the standard formula for the conditional of a Gaussian.

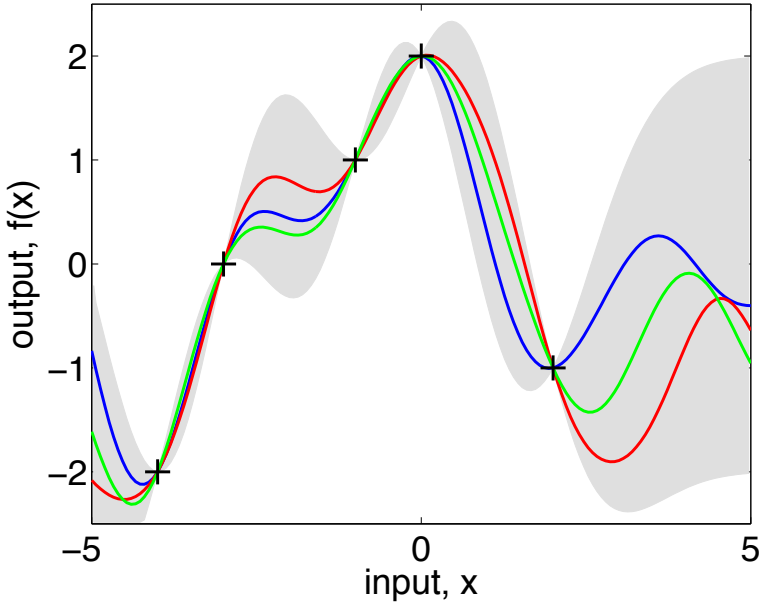
AN EXAMPLE

Consider again the 1-dimensional GP with mean function $\mu(x) \equiv 0$, and with Gaussian covariance function:

$$k(x, x') = \exp\left[-\frac{1}{2}|x - x'|^2\right]$$



(a), prior



(b), posterior

NOISY PREDICTIONS

- Suppose we cannot observe the values $\mathbf{f}$ of a GP at points X, but a perturbed version of them:

$$y(\mathbf{x}) = f(\mathbf{x}) + \varepsilon,$$

  where $\varepsilon \sim \mathcal{N}(0, \sigma^2)$

- Then the covariance of observations is $cov(\mathbf{y}) = K(X, X) + \sigma^2 I$
- The prior between observations X and test points X^* is then

$$\begin{bmatrix} \mathbf{y} \\ \mathbf{f}_* \end{bmatrix} \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} K(X,X) + \sigma_n^2 I & K(X, X_*) \\ K(X_*, X) & K(X_*, X_*) \end{bmatrix}\right). \tag{2.21}$$

- Conditioning on observations $\mathbf{y}$, we get

$$\mathbf{f}_* | X, \mathbf{y}, X_* \sim \mathcal{N}(\bar{\mathbf{f}}_*, \operatorname{cov}(\mathbf{f}_*)), \quad \text{where} \tag{2.22}$$

$$\bar{\mathbf{f}}_* \triangleq \mathbb{E}[\mathbf{f}_* | X, \mathbf{y}, X_*] = K(X_*, X)[K(X,X) + \sigma_n^2 I]^{-1}\mathbf{y}, \tag{2.23}$$

$$\operatorname{cov}(\mathbf{f}_*) = K(X_*, X_*) - K(X_*, X)[K(X,X) + \sigma_n^2 I]^{-1} K(X, X_*). \tag{2.24}$$

COMMENTS: LINEAR PREDICTOR

- For a single point $\mathbf{x}^*$, the predictive distribution reads

$$\bar{f}_* = \mathbf{k}_*^\top (K + \sigma_n^2 I)^{-1} \mathbf{y}, \tag{2.25}$$

$$\mathbb{V}[f_*] = k(\mathbf{x}_*, \mathbf{x}_*) - \mathbf{k}_*^\top (K + \sigma_n^2 I)^{-1} \mathbf{k}_*. \tag{2.26}$$

  where $\mathbf{k}_* = (k(\mathbf{x}^*, \mathbf{x_1}), \ldots, k(\mathbf{x}^*, \mathbf{x_N}))$

- It can be seen that the average prediction is a linear combination of the kernels evaluated on the input points:

$$\bar{f}(\mathbf{x}^*) = \sum_{i=1}^{N} \alpha_i k(\mathbf{x}^*, \mathbf{x_i})$$

  where $\alpha = (K + \sigma^2 I)^{-1} \mathbf{y}$.

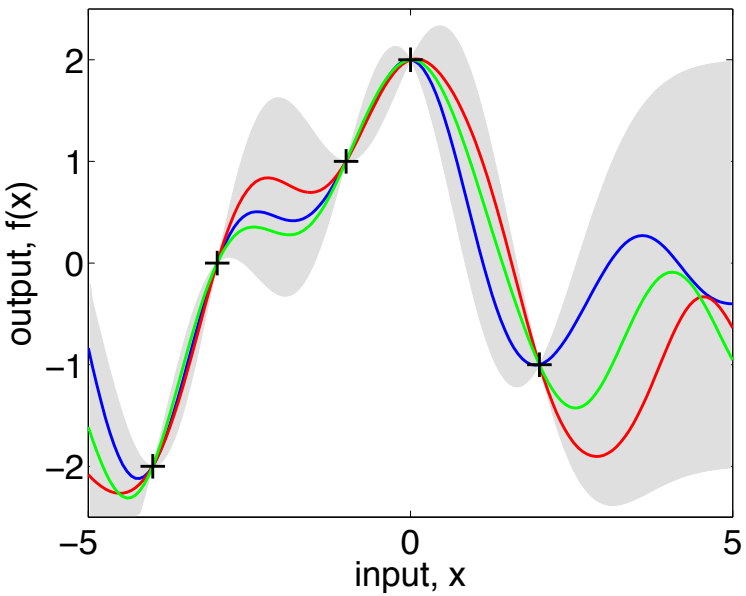
COMMENTS: POSTERIOR GP

- It is easy to see that the posterior process $f|\mathbf{y}$ is again a Gaussian process, with mean

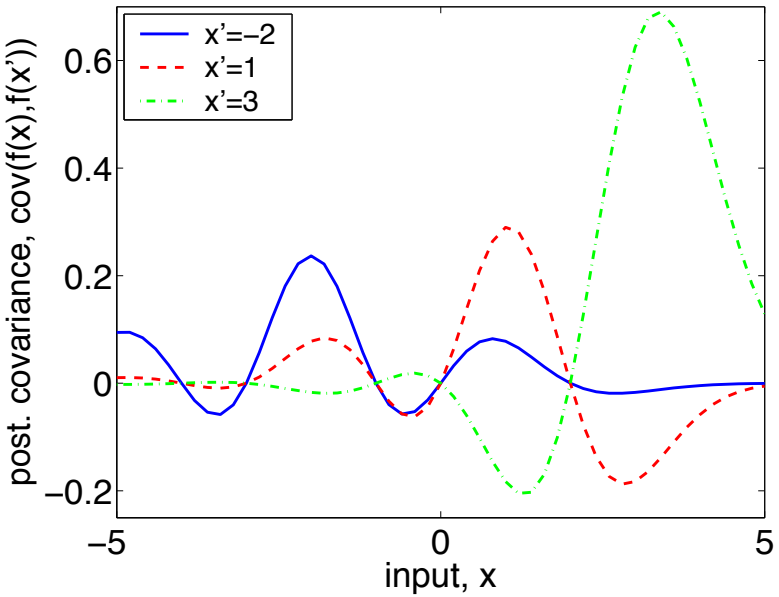
$$\mathbb{E}[f(\mathbf{x})|\mathbf{y}] = K(\mathbf{x}, X)(K + \sigma^2 I)^{-1}\mathbf{y}$$

and covariance

$$k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}, \mathbf{x}') - K(\mathbf{x}, X)(K + \sigma^2 I)^{-1}K(X, \mathbf{x}')$$



(a), posterior



(b), posterior covariance

Outline

KERNELS

- The notion of kernel comes from the theory of integral operators on a space $\mathcal{X}$ with measure μ. A real kernel $k : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ defines an integral operator T_k (applied to integrable f) as:

$$(T_k f)(\mathbf{x}) = \int_{\mathcal{X}} k(\mathbf{x}, \mathbf{y}) f(\mathbf{y}) d\mu(\mathbf{y})$$

- A kernel is positive semidefinite if, for all $f \in L_2(\mathcal{X}, \mu)$:

$$\int_{\mathcal{X}\times\mathcal{X}} k(\mathbf{x}, \mathbf{y}) f(\mathbf{x}) f(\mathbf{y}) d\mu(\mathbf{x}) d\mu(\mathbf{y}) \geq 0$$

- Equivalently, a kernel is positive (semi)definite if for any collection of n points $\{\mathbf{x_i} \mid i = 1, \ldots, n\}$, the Gram matrix K, $K_{ij} = I(\mathbf{x_i}, \mathbf{x_j})$ is positive (semi)definite (Mercer's theorem).
- The Gram matrix of a symmetric kernel, $k(\mathbf{x}, \mathbf{y}) = k(\mathbf{y}, \mathbf{x})$, is symmetric.

EIGENFUNCTIONS

- An eigenfunction ϕ with eigenvalue λ of k satisfies

$$\int k(\mathbf{x}, \mathbf{y})\phi(\mathbf{x})d\mu(\mathbf{x}) = \lambda\phi(\mathbf{y})$$

- There can be an infinite number of eigenfunctions, which can be ordered w.r.t. decreasing eigenvalues, and they can be chosen orthogonal, i.e. such that $\int \phi_i(\mathbf{x})\phi_j(\mathbf{x})d\mu(\mathbf{x}) = \delta_{ij}$
- A kernel can be decomposed using eigenfunctions:

Theorem 4.2 *(Mercer's theorem). Let $(\mathcal{X}, \mu)$ be a finite measure space and $k \in L_\infty(\mathcal{X}^2, \mu^2)$ be a kernel such that $T_k : L_2(\mathcal{X}, \mu) \to L_2(\mathcal{X}, \mu)$ is positive definite (see eq. (4.2)). Let $\phi_i \in L_2(\mathcal{X}, \mu)$ be the normalized eigenfunctions of T_k associated with the eigenvalues $\lambda_i > 0$. Then:*

1. *the eigenvalues $\{\lambda_i\}_{i=1}^\infty$ are absolutely summable*
2. $$k(\mathbf{x}, \mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x})\phi_i^*(\mathbf{x}'), \tag{4.37}$$

holds μ^2 almost everywhere, where the series converges absolutely and uniformly μ^2 almost everywhere. □

REPRODUCING KERNEL HILBERT SPACES

Definition 6.1 *(Reproducing kernel Hilbert space). Let $\mathcal{H}$ be a Hilbert space of real functions f defined on an index set $\mathcal{X}$. Then $\mathcal{H}$ is called a reproducing kernel Hilbert space endowed with an inner product $\langle \cdot, \cdot \rangle_{\mathcal{H}}$ (and norm $\|f\|_{\mathcal{H}} = \sqrt{\langle f, f \rangle_{\mathcal{H}}}$) if there exists a function $k : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ with the following properties:*

1. *for every $\mathbf{x}$, $k(\mathbf{x}, \mathbf{x}')$ as a function of $\mathbf{x}'$ belongs to $\mathcal{H}$, and*
2. *k has the* reproducing property *$\langle f(\cdot), k(\cdot, \mathbf{x}) \rangle_{\mathcal{H}} = f(\mathbf{x})$.* □

See e.g. Schölkopf and Smola [2002] and Wegman [1982]. Note also that as $k(\mathbf{x}, \cdot)$ and $k(\mathbf{x}', \cdot)$ are in $\mathcal{H}$ we have that $\langle k(\mathbf{x}, \cdot), k(\mathbf{x}', \cdot) \rangle_{\mathcal{H}} = k(\mathbf{x}, \mathbf{x}')$.

The RKHS uniquely determines k, and vice versa, as stated in the following theorem:

Theorem 6.1 *(Moore-Aronszajn theorem, Aronszajn [1950]). Let $\mathcal{X}$ be an index set. Then for every positive definite function $k(\cdot, \cdot)$ on $\mathcal{X} \times \mathcal{X}$ there exists a unique RKHS, and vice versa.* □

RKHS AND EIGENFUNCTIONS

- The functions belonging to the RKHS associated with a kernel k can be written as a linear combination of the eigenfunctions ϕ_j of k: $f(\mathbf{x}) = \sum_j f_j \phi_j(\mathbf{x})$, with $\sum_j f_j^2/\lambda_j < \infty$ (this is a smoothness constraint).
- Such functions define an Hilbert space H with inner product $\langle f, g \rangle_H = \sum_j \frac{f_j g_j}{\lambda_j}$
- This Hilbert space is the RKHS corresponding to kernel k:

$$\langle f(\cdot), k(\cdot, \mathbf{x}) \rangle_{\mathcal{H}} = \sum_{i=1}^{N} \frac{f_i \lambda_i \phi_i(\mathbf{x})}{\lambda_i} = f(\mathbf{x}). \tag{6.2}$$

Similarly

$$\langle k(\mathbf{x}, \cdot), k(\mathbf{x}', \cdot) \rangle_{\mathcal{H}} = \sum_{i=1}^{N} \frac{\lambda_i \phi_i(\mathbf{x}) \lambda_i \phi_i(\mathbf{x}')}{\lambda_i} = k(\mathbf{x}, \mathbf{x}'). \tag{6.3}$$

- Furthermore, the norm of $k(\mathbf{x}, \cdot)$ is $k(\mathbf{x}, \mathbf{x}) < \infty$: it belongs to H.

KERNEL FUNCTIONS: CLASSIFICATION

A kernel $k(\mathbf{x}, \mathbf{y})$ can be classified w.r.t dependence on $\mathbf{x}$ and $\mathbf{y}$.

- Stationary kernel: it is a function of $\mathbf{x} - \mathbf{y}$ (invariant to translations).
- Isotropic kernel: it is a function of $\|\mathbf{x} - \mathbf{y}\|$ (invariant to rigid motions).
- Dot-product kernel: it is a function of $\mathbf{x}^T\mathbf{y}$ (invariant w.r.t. rotations with respect to the origin).

Continuity properties of the GPs and kernels k.

- Continuity in mean square of a process f at $\mathbf{x}$: for each $\mathbf{x_k} \to \mathbf{x}$, it holds that $\mathbb{E}[\|f(\mathbf{x_k}) - f(\mathbf{x})\|^2] \to 0$.
- A process is continuous in m.s. at $\mathbf{x}$ iff k is continuous at $k(\mathbf{x}, \mathbf{x})$. For stationary kernels, k must be continuous at zero.
- If k is $2k$th differentiable, than f is kth differentiable (in m.s.).

GAUSSIAN KERNEL

- The Gaussian or Squared Exponential kernel is defined by

$$k(\mathbf{x}, \mathbf{y}) = \alpha \exp\left[-\frac{\|\mathbf{x} - \mathbf{y}\|^2}{\lambda^2}\right]$$

- α is called the amplitude, it regulates the magnitude of variance at each point $\mathbf{x}$. λ, instead, is the characteristic length-scale, which regulates the speed of decay of the correlation between points.
- The Gaussian kernel is isotropic and among the most used in computational statistics, and its RKHS is dense in the space of continuous functions over a compact set in $\mathbb{R}^n$.
- The Automatic-Relevance Detection Gaussian Kernel generalises the GK as

$$k(\mathbf{x}, \mathbf{y}) = \alpha \exp\left[-\sum_j \frac{|x_j - y_j|^2}{\lambda_j^2}\right]$$

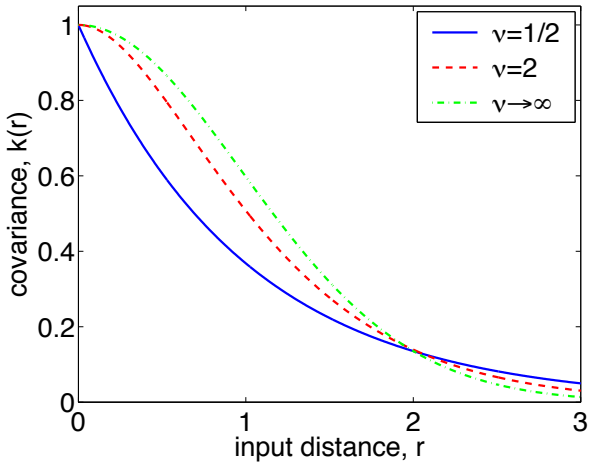
MATÉRN KERNEL

- The Matérn kernel is defined by

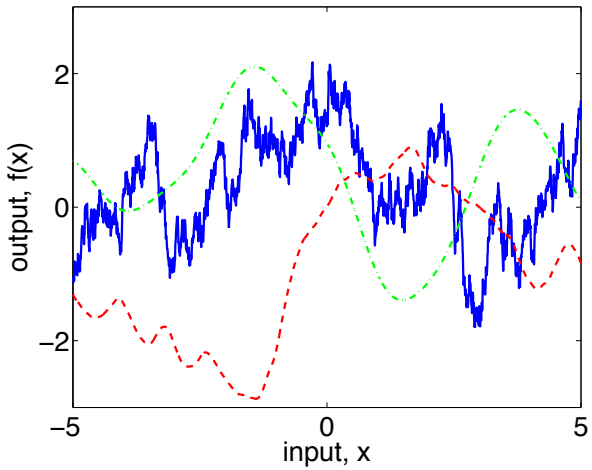
$$k_{\text{Matern}}(r) = \frac{2^{1-\nu}}{\Gamma(\nu)}\left(\frac{\sqrt{2\nu}r}{\ell}\right)^{\nu} K_\nu\left(\frac{\sqrt{2\nu}r}{\ell}\right), \tag{4.14}$$

with positive parameters ν and ℓ, where K_ν is a modified Bessel function

- If $\nu > h$, then the process with Matérn kernel is h times differentiable (in m.s.) For $\nu \to \infty$, then the MK becomes the GK.
- Examples of Matern Kernel:



(a)



(b)

MATÉRN AND EXPONENTIAL KERNEL

- Typical choice for MK is $\nu = p + 1/2$, giving

$$k_{\nu=p+1/2}(r) = \exp\left(-\frac{\sqrt{2\nu}r}{\ell}\right)\frac{\Gamma(p+1)}{\Gamma(2p+1)}\sum_{i=0}^{p}\frac{(p+i)!}{i!(p-i)!}\left(\frac{\sqrt{8\nu}r}{\ell}\right)^{p-i}. \quad (4.16)$$

It is possible that the most interesting cases for machine learning are $\nu = 3/2$ and $\nu = 5/2$, for which

$$\begin{aligned} k_{\nu=3/2}(r) &= \left(1+\frac{\sqrt{3}r}{\ell}\right)\exp\left(-\frac{\sqrt{3}r}{\ell}\right), \\ k_{\nu=5/2}(r) &= \left(1+\frac{\sqrt{5}r}{\ell}+\frac{5r^2}{3\ell^2}\right)\exp\left(-\frac{\sqrt{5}r}{\ell}\right), \end{aligned} \quad (4.17)$$

- for $\nu = 1/2$, we get the Exponential Kernel

$$k(\mathbf{x}, \mathbf{y}) = \exp\left[\|\mathbf{x} - \mathbf{y}\|/\lambda\right]$$

which in one dimension corresponds to the Ornstein-Ulembeck process (the model of velocity of a particle undergoing Brownian motion), which is continuous but nowhere differentiable.

POLYNOMIAL KERNEL

- Simple dot-products kernels are the polynomial kernel, for p integer:

$$k(\mathbf{x}, \mathbf{x}') = (\mathbf{x}^T \mathbf{x}')^p$$

- This corresponds to a kernel obtained by a set of polynomial basis functions:

$$\begin{aligned} k(\mathbf{x}, \mathbf{x}') &= (\mathbf{x} \cdot \mathbf{x}')^p = \Big(\sum_{d=1}^{D} x_d x'_d \Big)^p = \Big(\sum_{d_1=1}^{D} x_{d_1} x'_{d_1} \Big) \cdots \Big(\sum_{d_p=1}^{D} x_{d_p} x'_{d_p} \Big) \\ &= \sum_{d_1=1}^{D} \cdots \sum_{d_p=1}^{D} (x_{d_1} \cdots x_{d_p})(x'_{d_1} \cdots x'_{d_p}) \triangleq \boldsymbol{\phi}(\mathbf{x}) \cdot \boldsymbol{\phi}(\mathbf{x}'). \end{aligned} \tag{4.23}$$

- The basis functions ϕ_m are given by all monomials of degree p, i.e. $\sum m_j = p$:

$$\phi_{\mathbf{m}}(\mathbf{x}) = \sqrt{\frac{p!}{m_1! \cdots m_D!}} x_1^{m_1} \cdots x_D^{m_D}. \tag{4.24}$$

COMPOSITION OF KERNELS

Kernels can be composed according to certain rules, giving rise to new kernels.

Techniques for Constructing New Kernels.

Given valid kernels $k_1(\mathbf{x}, \mathbf{x}')$ and $k_2(\mathbf{x}, \mathbf{x}')$, the following new kernels will also be valid:

$$
\begin{aligned}
k(\mathbf{x}, \mathbf{x}') &= c k_1(\mathbf{x}, \mathbf{x}') && (6.13)\\
k(\mathbf{x}, \mathbf{x}') &= f(\mathbf{x}) k_1(\mathbf{x}, \mathbf{x}') f(\mathbf{x}') && (6.14)\\
k(\mathbf{x}, \mathbf{x}') &= q\left(k_1(\mathbf{x}, \mathbf{x}')\right) && (6.15)\\
k(\mathbf{x}, \mathbf{x}') &= \exp\left(k_1(\mathbf{x}, \mathbf{x}')\right) && (6.16)\\
k(\mathbf{x}, \mathbf{x}') &= k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}') && (6.17)\\
k(\mathbf{x}, \mathbf{x}') &= k_1(\mathbf{x}, \mathbf{x}') k_2(\mathbf{x}, \mathbf{x}') && (6.18)\\
k(\mathbf{x}, \mathbf{x}') &= k_3\left(\boldsymbol{\phi}(\mathbf{x}), \boldsymbol{\phi}(\mathbf{x}')\right) && (6.19)\\
k(\mathbf{x}, \mathbf{x}') &= \mathbf{x}^{\mathrm{T}} \mathbf{A} \mathbf{x}' && (6.20)\\
k(\mathbf{x}, \mathbf{x}') &= k_a(\mathbf{x}_a, \mathbf{x}'_a) + k_b(\mathbf{x}_b, \mathbf{x}'_b) && (6.21)\\
k(\mathbf{x}, \mathbf{x}') &= k_a(\mathbf{x}_a, \mathbf{x}'_a) k_b(\mathbf{x}_b, \mathbf{x}'_b) && (6.22)
\end{aligned}
$$

where $c > 0$ is a constant, $f(\cdot)$ is any function, $q(\cdot)$ is a polynomial with nonnegative coefficients, $\boldsymbol{\phi}(\mathbf{x})$ is a function from $\mathbf{x}$ to $\mathbb{R}^M$, $k_3(\cdot, \cdot)$ is a valid kernel in $\mathbb{R}^M$, $\mathbf{A}$ is a symmetric positive semidefinite matrix, $\mathbf{x}_a$ and $\mathbf{x}_b$ are variables (not necessarily disjoint) with $\mathbf{x} = (\mathbf{x}_a, \mathbf{x}_b)$, and k_a and k_b are valid kernel functions over their respective spaces.

Outline

MARGINAL LIKELIHOOD

- In order to do model selection (e.g. between different kernels) we can use the marginal likelihood.
- This can be used also to set hyperparameters of the kernel functions, like the amplitude or the lengthscale of the Gaussian kernel.
- For GP, we can compute the marginal likelihood analytically:

$$\mathcal{L} = \log p(\mathbf{y}|X) = \log \int p(\mathbf{f}|X)p(\mathbf{y}|\mathbf{f}, X)d\mathbf{f}$$

which gives

$$\mathcal{L} = -\frac{1}{2}\mathbf{y}^T(K + \sigma^2 I)^{-1}\mathbf{y} - \frac{1}{2}\log|(K + \sigma^2 I)| - \frac{N}{2}\log 2\pi$$

- This follows also by observing that $\mathbf{y} \sim \mathcal{N}(\mathbf{0}, K + \sigma^2 I)$.

MARGINAL LIKELIHOOD

The log marginal likelihood

$$\mathcal{L} = -\frac{1}{2}\mathbf{y}^T(K + \sigma^2 I)^{-1}\mathbf{y} - \frac{1}{2}\log|(K + \sigma^2 I)| - \frac{N}{2}\log 2\pi$$

has three terms

- $-\frac{1}{2}\mathbf{y}^T(K + \sigma^2 I)^{-1}\mathbf{y}$ is the data fit.
- $-\frac{1}{2}\log|(K + \sigma^2 I)|$ is a complexity penalty.
- $-\frac{N}{2}\log 2\pi$ is a constant.

MARGINAL LIKELIHOOD - HYPERPARAMETERS

Data from 1dim example with Gaussian kernels

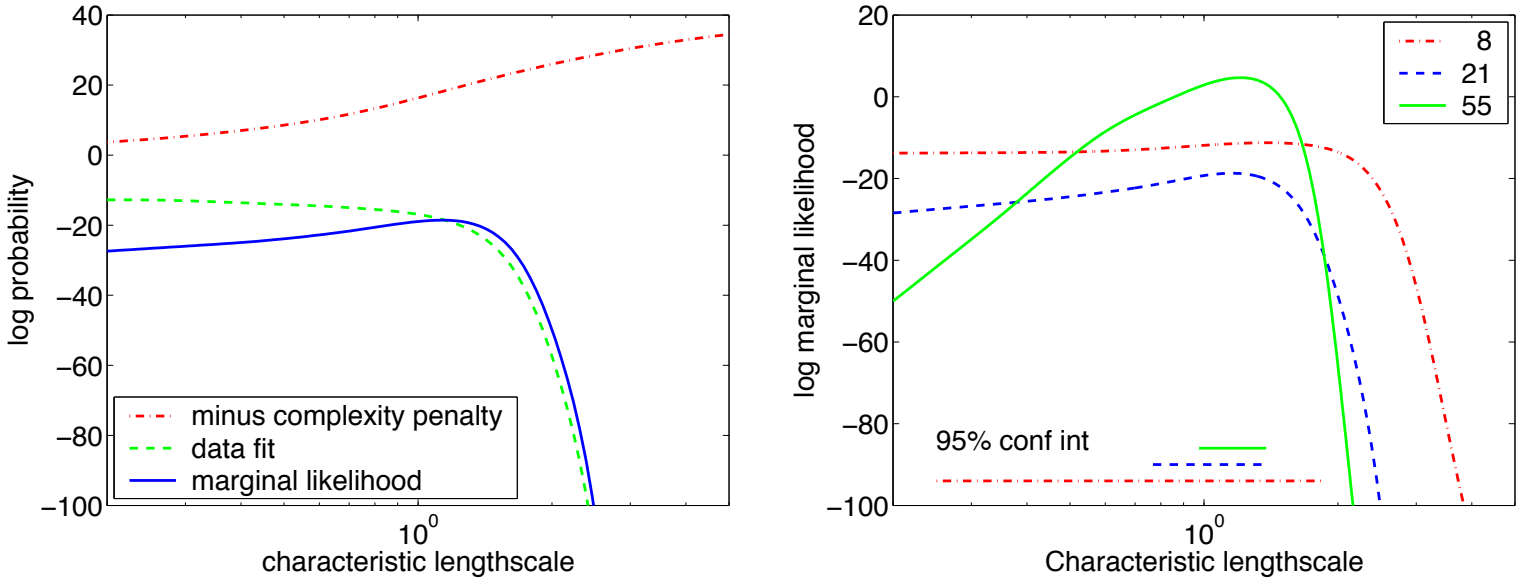


Figure 5.3: Panel (a) shows a decomposition of the log marginal likelihood into its constituents: data-fit and complexity penalty, as a function of the characteristic length-scale. The training data is drawn from a Gaussian process with SE covariance function and parameters $(\ell, \sigma_f, \sigma_n) = (1, 1, 0.1)$, the same as in Figure 2.5, and we are fitting only the length-scale parameter ℓ (the two other parameters have been set in accordance with the generating process). Panel (b) shows the log marginal likelihood as a function of the characteristic length-scale for different sizes of training sets. Also shown, are the 95% confidence intervals for the posterior length-scales.

MARGINAL LIKELIHOOD - HYPERPARAMETERS

Data from 1dim example with Gaussian kernels

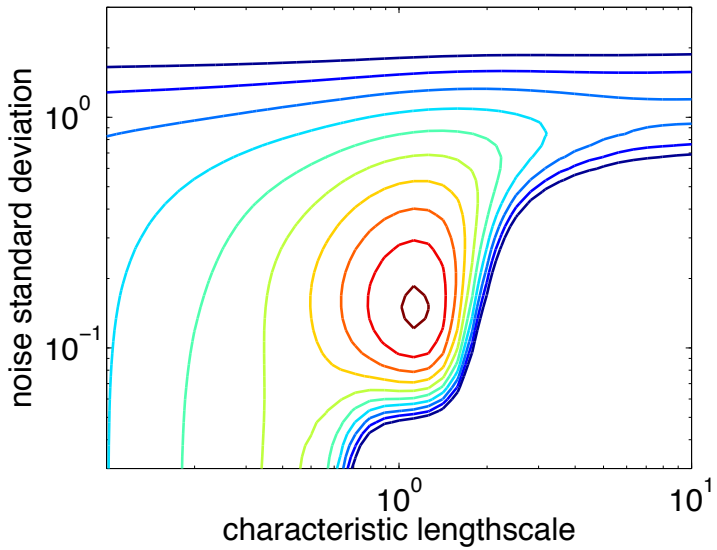
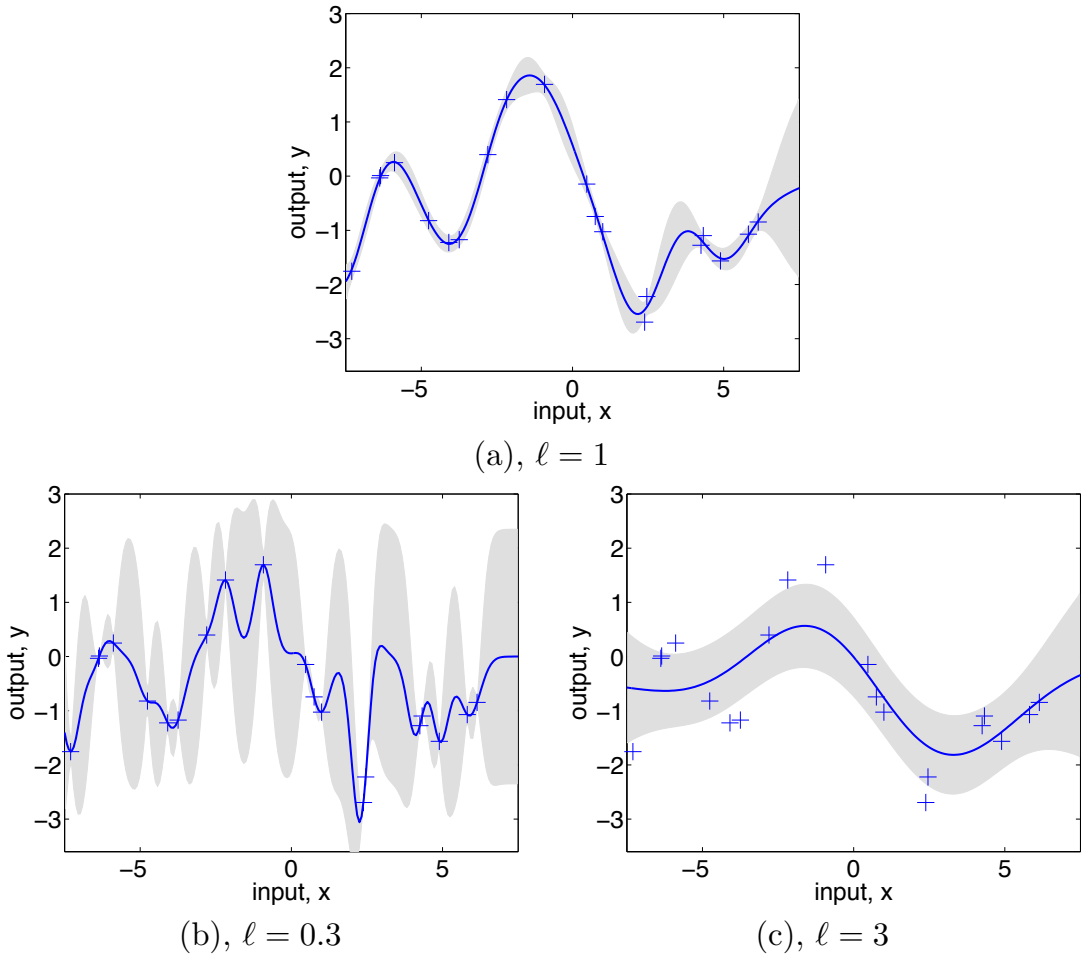


Figure 5.4: Contour plot showing the log marginal likelihood as a function of the characteristic length-scale and the noise level, for the same data as in Figure 2.5 and Figure 5.3. The signal variance hyperparameter was set to $\sigma_f^2 = 1$. The optimum is close to the parameters used when generating the data. Note, the two ridges, one for small noise and length-scale $\ell = 0.4$ and another for long length-scale and noise $\sigma_n^2 = 1$. The contour lines spaced 2 units apart in log probability density.

MARGINAL LIKELIHOOD - HYPERPARAMETERS

Data coming from a sample of a 1dim GP with Gaussian kernel and hyperparameters $\lambda = 1$, $\alpha = 1$, $\sigma = 0.1$.



(a), $\ell = 1$

(b), $\ell = 0.3$

(c), $\ell = 3$

Figure 2.5: (a) Data is generated from a GP with hyperparameters $(\ell, \sigma_f, \sigma_n) = (1, 1, 0.1)$, as shown by the + symbols. Using Gaussian process prediction with these hyperparameters we obtain a 95% confidence region for the underlying function f (shown in grey). Panels (b) and (c) again show the 95% confidence region, but this time for hyperparameter values $(0.3, 1.08, 0.00005)$ and $(3.0, 1.16, 0.89)$ respectively.

HYPERPARAMETER OPTIMISATION

- In order to set the hyperparameters, we can maximise the log marginal likelihood:

$$\mathcal{L} = -\frac{1}{2}\mathbf{y}^T(K + \sigma^2 I)^{-1}\mathbf{y} - \frac{1}{2}\log|(K + \sigma^2 I)| - \frac{N}{2}\log 2\pi$$

- Its derivative w.r.t. an hyperparameter θ is

$$\begin{aligned}\frac{\partial}{\partial\theta_j}\log p(\mathbf{y}|X,\boldsymbol{\theta}) &= \frac{1}{2}\mathbf{y}^\top K^{-1}\frac{\partial K}{\partial\theta_j}K^{-1}\mathbf{y} - \frac{1}{2}\operatorname{tr}\left(K^{-1}\frac{\partial K}{\partial\theta_j}\right)\\ &= \frac{1}{2}\operatorname{tr}\left((\boldsymbol{\alpha}\boldsymbol{\alpha}^\top - K^{-1})\frac{\partial K}{\partial\theta_j}\right) \quad \text{where} \quad \boldsymbol{\alpha} = K^{-1}\mathbf{y}.\end{aligned} \tag{5.9}$$

- The derivative is relatively cheap to compute, once we invert the matrix K. Hence we can use gradient methods to optimise $\mathcal{L}$.
- Purely Bayesian methods (giving a prior on hyperparameters) are complicated by the in general complex functional form (no conjugate prior).

NON-CONSTANT PRIOR MEAN

- The typical choice for the prior mean is the zero function. Data is processed by subtracting the sample mean from the observations.
- As an alternative, one can either use a deterministic function for the priori mean (and subtract if from data, adding it back to predictions), or use a generalised linear model for the prior mean:

$$g(\mathbf{x}) \;=\; f(\mathbf{x}) + \mathbf{h}(\mathbf{x})^\top \boldsymbol{\beta}, \quad \text{where} \quad f(\mathbf{x}) \;\sim\; \mathcal{GP}\big(0, k(\mathbf{x}, \mathbf{x}')\big), \tag{2.39}$$

- If we put a Gaussian prior over coefficients β, we can treat them in a Bayesian way, and get a GP:

$$g(\mathbf{x}) \;\sim\; \mathcal{GP}\big(\mathbf{h}(\mathbf{x})^\top \mathbf{b},\; k(\mathbf{x}, \mathbf{x}') + \mathbf{h}(\mathbf{x})^\top B \mathbf{h}(\mathbf{x}')\big), \tag{2.40}$$

NON-CONSTANT PRIOR MEAN

- In this way, we obtain the following predictive distribution at a point $\mathbf{x}^*$:

$$\begin{aligned} \bar{\mathbf{g}}(X_*) &= H_*^\top \bar{\boldsymbol{\beta}} + K_*^\top K_y^{-1}(\mathbf{y} - H^\top \bar{\boldsymbol{\beta}}) = \bar{\mathbf{f}}(X_*) + R^\top \bar{\boldsymbol{\beta}}, \\ \mathrm{cov}(\mathbf{g}_*) &= \mathrm{cov}(\mathbf{f}_*) + R^\top (B^{-1} + H K_y^{-1} H^\top)^{-1} R, \end{aligned} \tag{2.41}$$

where the H matrix collects the $\mathbf{h}(\mathbf{x})$ vectors for all training (and H_* all test) cases, $\bar{\boldsymbol{\beta}} = (B^{-1} + H K_y^{-1} H^\top)^{-1}(H K_y^{-1}\mathbf{y} + B^{-1}\mathbf{b})$, and $R = H_* - H K_y^{-1} K_*$.

- The new predictive distribution has mean $H_*^T \bar{\beta}$ (from the linear model) plus a term coming from the GP model of residuals.
- Taking a flat prior (limit for $B^-1 \rightarrow$ matrix of zeros):

$$\begin{aligned} \bar{\mathbf{g}}(X_*) &= \bar{\mathbf{f}}(X_*) + R^\top \bar{\boldsymbol{\beta}}, \\ \mathrm{cov}(\mathbf{g}_*) &= \mathrm{cov}(\mathbf{f}_*) + R^\top (H K_y^{-1} H^\top)^{-1} R, \end{aligned} \tag{2.42}$$

where the limiting $\bar{\boldsymbol{\beta}} = (H K_y^{-1} H^\top)^{-1} H K_y^{-1}\mathbf{y}$. Notice that predictions under

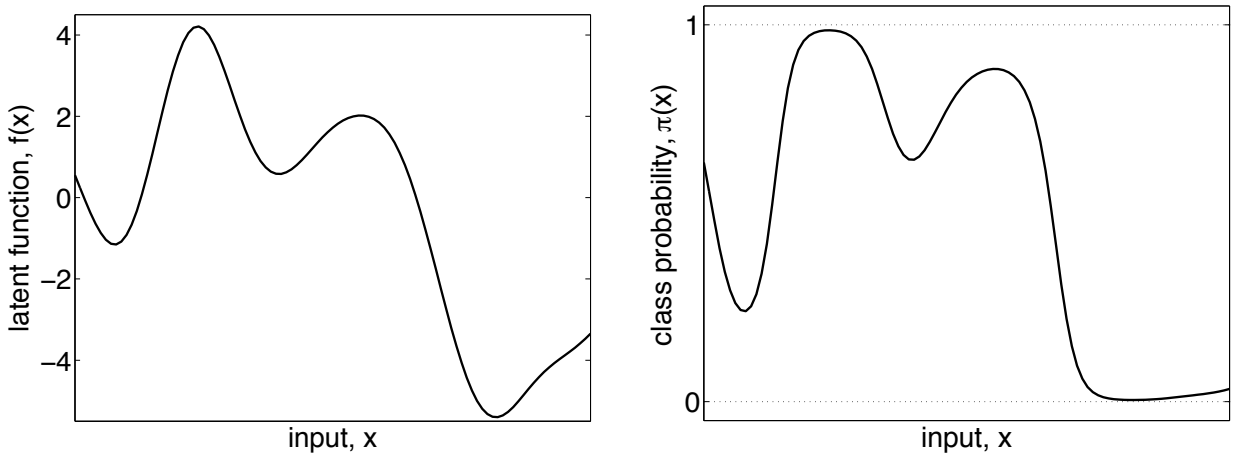
Outline

FROM LOGISTIC REGRESSION TO GP CLASSIFICATION

- The idea behind GP classification is to extend logistic (or probit) regression, by assuming the following model for the class conditionals:

$$\pi(\mathbf{x}) = p(C_1|\mathbf{x}) = \sigma(f(\mathbf{x})) \text{ where } f \sim GP(\mu, k)$$

- f is often call latent function. Note that π is a random function, as f is.



(a) (b)

Figure 3.2: Panel (a) shows a sample latent function $f(x)$ drawn from a Gaussian process as a function of x. Panel (b) shows the result of squashing this sample function through the logistic logit function, $\lambda(z) = (1 + \exp(-z))^{-1}$ to obtain the class probability $\pi(x) = \lambda(f(x))$.

GP CLASSIFICATION

- f is often call latent or nuisance function. It is not observed directly. We only observe at a point $\mathbf{x}$ the realisation of a Bernoulli random variable with probability $\pi(\mathbf{x})$.
- Inference at a test point $\mathbf{x}^*$ is done, as usual in a Bayesian setting, in two steps:
  1. Compute the posterior f^* of f at the prediction point $\mathbf{x}^*$.

$$p(f_*|X,\mathbf{y},\mathbf{x}_*) = \int p(f_*|X,\mathbf{x}_*,\mathbf{f})p(\mathbf{f}|X,\mathbf{y})\,d\mathbf{f}, \tag{3.9}$$

  with $p(\mathbf{f}|X,\mathbf{y}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|X)/p(\mathbf{y}/X)$ by Bayes theorem.
  2. Compute the predictive distribution at $\mathbf{x}^*$

$$\bar{\pi}_* \triangleq p(y_*=+1|X,\mathbf{y},\mathbf{x}_*) = \int \sigma(f_*)p(f_*|X,\mathbf{y},\mathbf{x}_*)\,df_*. \tag{3.10}$$

LAPLACE APPROXIMATION

- As in Bayesian logistic regression, the computation of the posterior $p(\mathbf{f}|X, \mathbf{y})$ cannot be carried out analytically.
- However, we can do a Laplace approximation of the posterior around the MAP $\hat{f}$. The unnormalised log posterior is:

$$\begin{aligned} \Psi(\mathbf{f}) &\triangleq \log p(\mathbf{y}|\mathbf{f}) + \log p(\mathbf{f}|X) \\ &= \log p(\mathbf{y}|\mathbf{f}) - \frac{1}{2}\mathbf{f}^\top K^{-1}\mathbf{f} - \frac{1}{2}\log|K| - \frac{n}{2}\log 2\pi. \end{aligned} \tag{3.12}$$

Differentiating eq. (3.12) w.r.t. $\mathbf{f}$ we obtain

$$\nabla\Psi(\mathbf{f}) = \nabla \log p(\mathbf{y}|\mathbf{f}) - K^{-1}\mathbf{f}, \tag{3.13}$$

$$\nabla\nabla\Psi(\mathbf{f}) = \nabla\nabla \log p(\mathbf{y}|\mathbf{f}) - K^{-1} = -W - K^{-1}, \tag{3.14}$$

where W is diagonal, as observations are i.i.d.

- It can be optimised with a Newton-Rapson scheme:

$$\begin{aligned} \mathbf{f}^{\text{new}} = \mathbf{f} - (\nabla\nabla\Psi)^{-1}\nabla\Psi &= \mathbf{f} + (K^{-1} + W)^{-1}(\nabla \log p(\mathbf{y}|\mathbf{f}) - K^{-1}\mathbf{f}) \\ &= (K^{-1} + W)^{-1}\big(W\mathbf{f} + \nabla \log p(\mathbf{y}|\mathbf{f})\big). \end{aligned} \tag{3.18}$$

LAPLACE APPROXIMATION

- The Laplace approximation around the MAP $\hat{f}$ is a Gaussian q with mean

$$\mathbb{E}_q[f_*|X,\mathbf{y},\mathbf{x}_*] \;=\; \mathbf{k}(\mathbf{x}_*)^\top K^{-1}\hat{\mathbf{f}} \;=\; \mathbf{k}(\mathbf{x}_*)^\top \nabla \log p(\mathbf{y}|\hat{\mathbf{f}}). \tag{3.21}$$

  and variance

$$\begin{aligned}\mathbb{V}_q[f_*|X,\mathbf{y},\mathbf{x}_*] \;&=\; k(\mathbf{x}_*,\mathbf{x}_*) - \mathbf{k}_*^\top K^{-1}\mathbf{k}_* + \mathbf{k}_*^\top K^{-1}(K^{-1}+W)^{-1}K^{-1}\mathbf{k}_* \\ &=\; k(\mathbf{x}_*,\mathbf{x}_*) - \mathbf{k}_*^\top (K+W^{-1})^{-1}\mathbf{k}_*, \end{aligned} \tag{3.24}$$

- The prediction π^* can be computed by the integral

$$\bar{\pi}_* \;\simeq\; \mathbb{E}_q[\pi_*|X,\mathbf{y},\mathbf{x}_*] \;=\; \int \sigma(f_*) q(f_*|X,\mathbf{y},\mathbf{x}_*)\, df_*, \tag{3.25}$$

  which can be approximated with the same logit-probit-logit trick used for Bayesian logistic regression.

EXPECTATION PROPAGATION

- A (better) alternative to Laplace approximation is to use a variational method, typically for the probit activation function.
- A first option is to approximate the posterior distribution by a Gaussian q, minimising the (reversed) KL divergence $KL(q(\mathbf{f}|X,\mathbf{y}), p(\mathbf{f}|X,\mathbf{y}))$ (the minimisation of the KL divergence $KL(p(\mathbf{f}|X,\mathbf{y}), q(\mathbf{f}|X,\mathbf{y}))$ is intractable).
- Alternatively, one can use the Expectation Propagation algorithm, which constructs iteratively (over obs i, until convergence) a Gaussian approximation of the posterior by
  1. taking the current Gaussian approximation and factoring out the term for the i-th likelihood $p(y_i|f_i)$, obtaining a distribution for all observations but the i-th one.
  2. multiplying the cavity by the exact likelihood of the i-th observation, and finding a Gaussian approximation by moment matching of such a (non-Gaussian) distribution.
- EP is more accurate than Laplace approximation, and provides also an approximation of the Marginal likelihood.

PITFALLS OF GP PREDICTION

- Addition of a new observation *always* reduces uncertainty at all points → vulnerable to outliers
- Optimisation of hyperparameters often tricky: works well if σ^2 is known, otherwise it can be seriously multimodal
- **MAIN PROBLEM: GP prediction relies on a matrix inversion which scales cubically with the number of points!**
- Sparsification methods have been proposed but in high dimension GP regression is likely to be tricky nevertheless