

Inference for a mean from a Gaussian dist.

$y_1, \dots, y_n \mid \theta \sim \text{i.i.d. } \mathcal{N}(\theta, \sigma^2)$   
 $\sigma^2$  known

$$p(y \mid \theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2\sigma^2}(y_i - \theta)^2\right\}$$

$$= (2\pi)^{-n/2} (\sigma^2)^{-n/2} \exp\left\{-\frac{1}{2\sigma^2} \sum (y_i - \theta)^2\right\}$$

$$\propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \theta)^2\right\}$$

$$\propto \exp\left\{-\frac{1}{2\sigma^2} \left( \sum y_i^2 - 2\theta \sum y_i + n\theta^2 \right)\right\}$$

$$p(y|\theta) \propto \exp\left\{-\frac{1}{2\sigma^2} (-2\sigma n \bar{y} + n\sigma^2)\right\}$$

is exponential form  $t(y)$  sufficient

$$\propto \underbrace{\exp\left\{-\frac{n\sigma^2}{2\sigma^2}\right\}}_{g(\theta)^n} \exp\left\{+\frac{\sigma}{\sigma^2} n \bar{y}\right\}$$

$$g(\theta)^n \exp\left\{\phi(\theta) t(y)\right\}$$

$$\pi(\theta) \propto \exp\left\{-\frac{n\sigma^2}{2\sigma^2}\right\} \exp\left\{\frac{\sigma}{\sigma^2} n\right\}$$

Prior for  $\theta$

$$\pi(\theta) \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$p(y|\theta) \propto \exp\left\{-\frac{n}{2\sigma^2}(\bar{y}-\theta)^2\right\}$$

dist<sup>n</sup> of sufficient statistic.

$$\bar{y}|\theta \sim \mathcal{N}\left(\theta, \frac{\sigma^2}{n}\right)$$

(usual result for the sample mean)

# Posterior

$$\pi(\theta | y) \propto \pi(\theta) p(y | \theta)$$

$$\propto \exp \left\{ -\frac{1}{2\sigma_0^2} (\theta - \mu_0)^2 \right\} \exp \left\{ -\frac{n}{2\sigma^2} (\bar{y} - \theta)^2 \right\}$$

$$\propto \exp \left\{ -\frac{1}{2\sigma_0^2} (\theta^2 - 2\mu_0\theta) - \frac{n}{2\sigma^2} (\theta^2 - 2\theta\bar{y}) \right\}$$

$$\propto \exp \left\{ -\frac{1}{2} \left( \theta^2 \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) - 2\theta \left( \frac{\mu_0}{\sigma_0^2} + \frac{\bar{y}n}{\sigma^2} \right) \right) \right\}$$

$$\propto \exp \left\{ -\frac{1}{2} \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \left( \theta^2 - 2\theta \frac{\frac{\mu_0}{\sigma_0^2} + \frac{\bar{y}n}{\sigma^2}}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}} \right) \right\}$$

$$\pi(\theta|y) \propto \exp \left\{ -\frac{1}{2} \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \left( \theta - \frac{\frac{\mu_0}{\sigma_0^2} + \bar{y} \frac{n}{\sigma^2}}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}} \right)^2 \right\}$$

$$\theta|y \sim \mathcal{N} \left( \underbrace{\frac{\frac{\mu_0}{\sigma_0^2} + \bar{y} \frac{n}{\sigma^2}}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}}}_{\mu_n}, \underbrace{\frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}}}_{\sigma_n^2} \right)$$

$$\begin{aligned} \mu_n &\xrightarrow{n \rightarrow \infty} \bar{y} \\ \mu_n &\xrightarrow{\sigma_0^2 \rightarrow 0} \mu_0 \\ \mu_n &\xrightarrow{\sigma_0^2 \rightarrow \infty} \bar{y} \end{aligned}$$

$$\begin{aligned} \sigma_n^2 &\xrightarrow{n \rightarrow \infty} 0 \\ \sigma_n^2 &\xrightarrow{\sigma_0^2 \rightarrow 0} 0 \\ \sigma_n^2 &\xrightarrow{\sigma_0^2 \rightarrow \infty} \frac{\sigma^2}{n} \end{aligned}$$

Predictive distribution  $\tilde{y} | \sigma \sim \mathcal{N}(\sigma, \sigma^2)$   
 exchange.  $\rightarrow$  indep. of  $y_1, \dots, y_n | \sigma$

$$p(\tilde{y} | y) \propto \int p(\tilde{y} | \sigma, X) \pi(\sigma | y) d\sigma$$

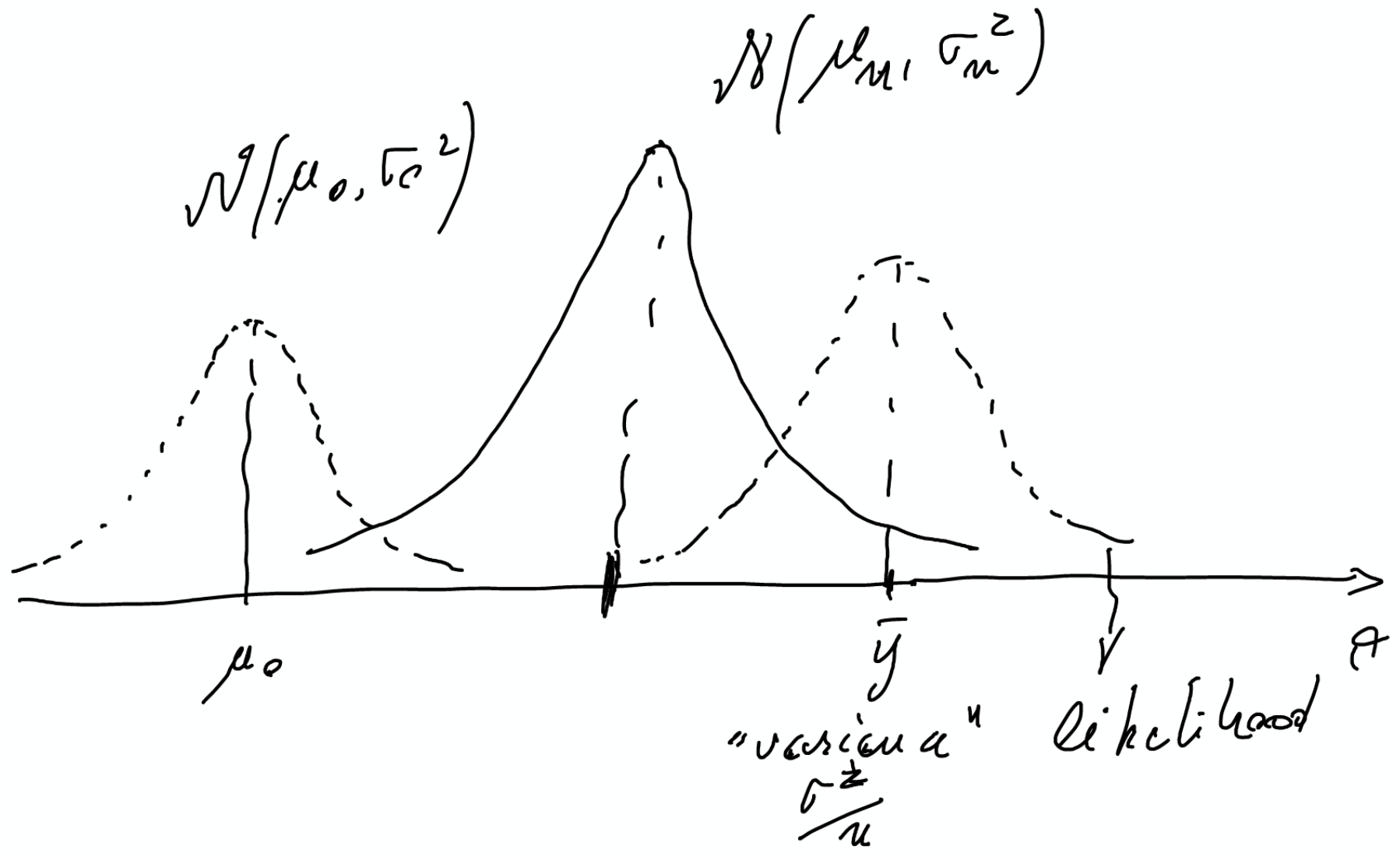
$$\propto \int \exp\left\{-\frac{1}{2\sigma^2} (\tilde{y} - \sigma)^2\right\} \exp\left\{-\frac{1}{2\tau_n^2} (\sigma - \mu_n)^2\right\} d\sigma$$

$$\int \exp\left\{ \dots \text{quadratic in } \tilde{y}, \sigma \dots \right\} d\sigma$$

joint dist. of  $(\tilde{y}, \sigma)$  is  
 Gaussian  $\Rightarrow$  marginal for  
 $\tilde{y}$  is Gaussian

$$E(\tilde{y} | y) = E \left( \underbrace{E(\tilde{y} | \tilde{y}, \theta)}_{\theta} \mid y \right) = \underline{\mu_n}$$

$$\begin{aligned} V(\tilde{y} | y) &= E \left( \underbrace{V(\tilde{y} | \tilde{y}, \theta)}_{= \sigma^2} \mid y \right) + V \left( \underbrace{E(\tilde{y} | \theta)}_{\theta} \mid y \right) \\ &= \sigma^2 + \sigma_n^2 \end{aligned}$$





# Interval estimate for the mean $\theta$

posterior intervals (quantiles)

→ high posterior density regions

$$\theta | y \sim \mathcal{N}(\mu_u, \sigma_u^2)$$

Posterior interval prob.  $1-\alpha$

$$\mu_u \pm z_{1-\frac{\alpha}{2}} \sigma_u$$

if  $\sigma_u \rightarrow \infty$  then approx.

$$\bar{y} \pm z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$$

$z_\alpha = \Phi^{-1}(\alpha)$   
 $\alpha$  quantile  
of standard  
normal  
dist

