

Sampling Random Variates

- X random variable.

Generate SAMPLES/OBSERVATIONS/REALIZATIONS

x_1, \dots, x_N of R.V. X

EXAMPLES: THROW A COIN
ROLL A DIE

DO IT ON A COMPUTER

1) ADD A RANDOMNESS SOURCE (EXTERNAL)

2) PSEUDO-RANDOM NUMBER GENERATOR

generate a sequence of numbers which looks random. \rightarrow passes some statistical test.

ALL PRNGs generate samples from UNIFORM DISTRIBUTION in $[0,1]$

$$X \sim U(0,1) \quad \text{iff} \quad p(x) = \begin{cases} 1, & x \in [0,1] \\ 0, & \text{otherwise} \end{cases}$$

•) PRNGs in pills

•) SAMPLING OTHER R.V.

⋮

•) MCMC

PSEUDO-RANDOM NUMBER GENERATOR (PRNG)

def: $(\mathcal{S}, \mathcal{U}, s_0, f, g)$

- \mathcal{S} STATE SPACE
- \mathcal{U} OUTPUT SPACE
- $s_0 \in \mathcal{S}$ INITIAL SEED
- $f: \mathcal{S} \rightarrow \mathcal{S}$
- $g: \mathcal{S} \rightarrow \mathcal{U}$

$$\mathcal{S} = \mathbb{Z} \bmod m = \{0, \dots, m-1\}$$

m LARGE

$$\mathcal{U} = [0, 1]$$

$$g(s) = \frac{s}{m}$$

PRNG WORKS by producing a sequence

$$s_0, s_1 = f(s_0), s_2 = f(s_1), \dots, s_i = f(s_{i-1}), \dots$$

$$u_1 = g(s_1), u_2 = g(s_2), \dots, u_i = g(s_i), \dots$$

- m IS VERY BIG
- \mathcal{S} IS FINITE
- $S = \{s_i\}_{i \in \mathbb{N}}$ $s_i = f(s_{i-1})$

$$\exists K, p \gg 0 : \forall i \gg K$$

$$\underline{s_{i+p} = s_i}$$

- we want $u_1, u_2, \dots, u_i, \dots$ looks RANDOM.

LINEAR CONGRUENTIAL GENERATORS (LCG)

$$f(s_i) = (\underset{\substack{\text{multiplier} \\ a}}{a} \cdot s_i + \underset{\substack{\text{increment} \\ c}}{c}) \bmod m$$

$$s = Z \bmod m$$

ADDITIVE LCG

if $c=0$, MULTIPLICATIVE LCG

• A LCG has FULL PERIOD ($p=m$) iff

1) c and m are RELATIVELY PRIME

2) q is prime and $q|m \Rightarrow q|a-1$
 q divides m

3) $4|m \Rightarrow 4|a-1$

Example: $a=5, c=3, m=2^9$ ✓

$m=2^{63}$ for 64-bit precision for double.

COMPOSITE LCG

$$1) \mathcal{S} = \sum_{\mathbb{Z}_m}^{q+1} = \{0, \dots, m-1\}^{q+1}, |\mathcal{S}| = m^{q+1}$$

$$(s_{i+1}, s_i, \dots, s_{i-q+1}) = f(s_i, \dots, s_{i-q}) \quad \boxed{\begin{array}{l} a_0 s_i + \dots + a_q s_{i-q} \pmod m \\ \vdots \\ s_{i+1} \end{array}}$$

2) q different LCG

$s_{1,i}, \dots, s_{q,i}$ are states of q different generators.

$$s_i = (\delta_1 s_{1,i} + \dots + \delta_q s_{q,i}) \pmod m$$

We can have very large p , e.g. $p = 2^{191} \approx 3 \cdot 10^{57}$ (rand.c)

$$s_{j,i} \in \mathbb{Z}_{m_j} \quad \mathcal{S} = \prod_j \mathbb{Z}_{m_j} \quad |\mathcal{S}| = m_1 \times \dots \times m_q$$

TESTING PRNGs

Consider x_1, \dots, x_n of samples

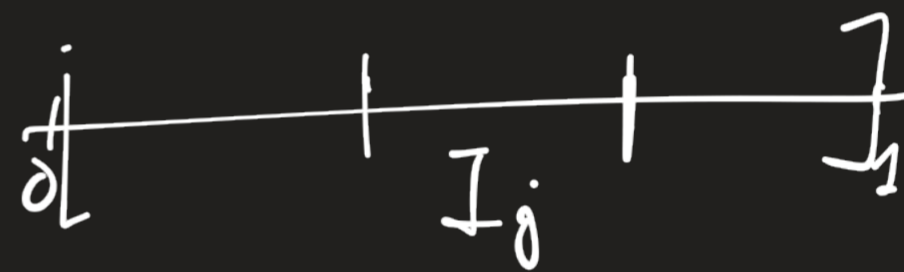
$H_0: x_1, \dots, x_n \sim U(0,1)$ and are independent

STATISTICAL TEST: χ^2 -TEST

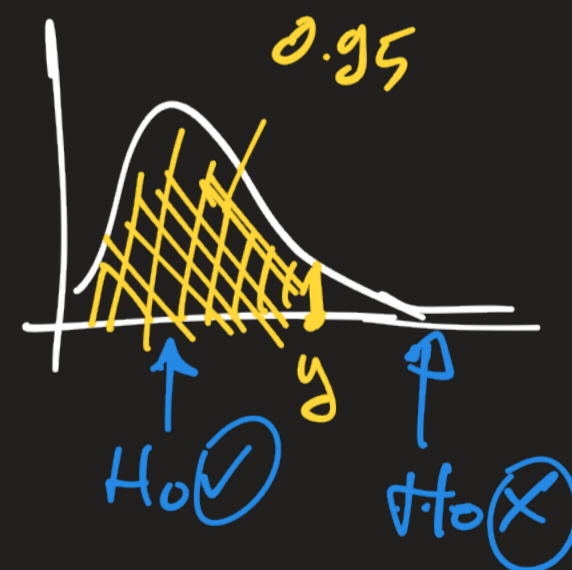
1) DIVIDE $[0,1]$ in K sub-intervals of equal length $1/K$. I_j

2) Count: $N_j = \sum_{i=1}^n \mathbb{1}(x_i \in I_j)$

$\mathbb{1}$ if $x_i \in I_j$
0 otherwise



$$3) \text{CHI2} = \sum_{j=1}^K (N_j - \frac{n}{K})^2 / \frac{n}{K} = \frac{K}{n} \cdot \sum_{j=1}^K (N_j - \frac{n}{K})^2$$



FOR LARGE n , CHI2 IS APPROX. χ^2 -distributed with $K-1$

$[Y \sim \chi^2(K-1) \text{ iff } Y = \sum_{j=1}^{K-1} Z_j^2, Z_j \sim N(0,1)]$ degrees of freedom

FIND y s.t. $P(Y \leq y) = 0.95$ ^{any confidence level} ACCEPT H_0 iff $\text{CHI2} \leq y$

STRONGER TEST
BY CONSIDERING BLOCKS OF
↓ SAMPLES AND TESTING
THAT THEY COME FROM
 $U([0,1]^d)$

INVERSION METHOD

X real valued R.V. (\mathbb{R}) with p density $p(x)$

CUMULATIVE DISTRIBUTION FUNCTION $F(x) = P(X \leq x) = \int_{-\infty}^x p(y) dy$

F IS MONOTONE, $F(-\infty) = 0$ ($\lim_{x \rightarrow -\infty} F(x) = 0$)

$$F(+\infty) = 1$$

X is abs. continuous, then F is continuous and strictly monotonic

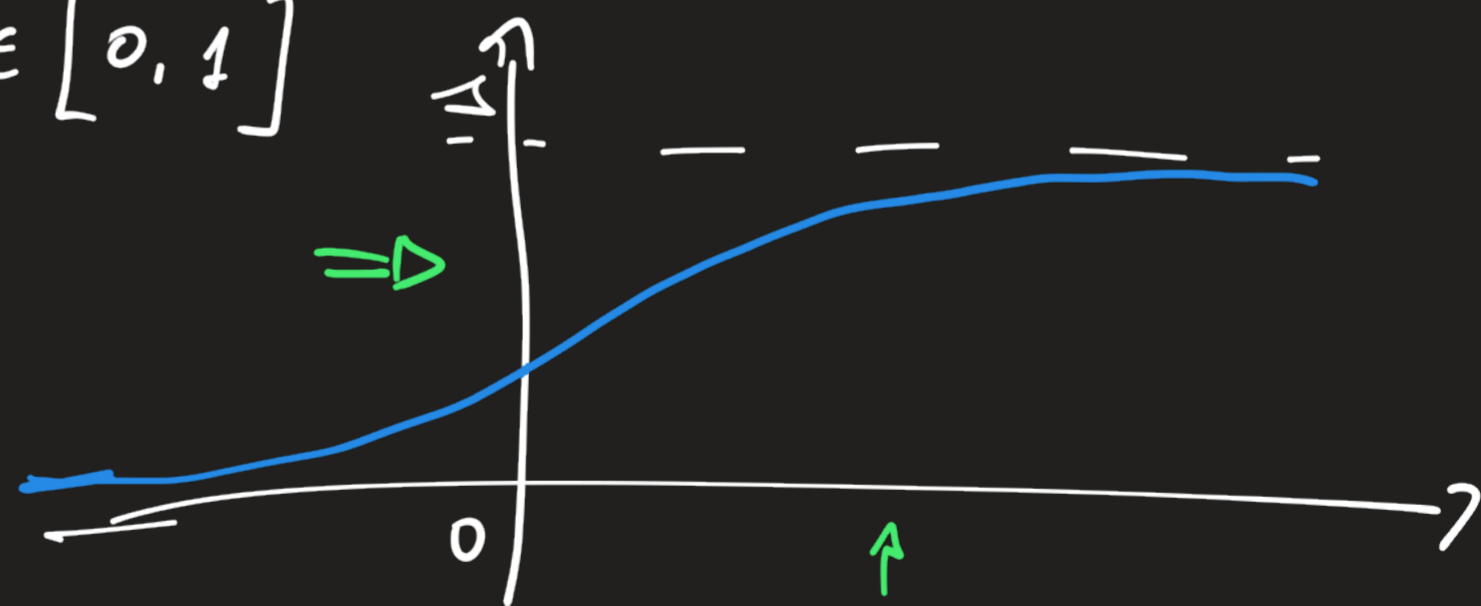
$\Rightarrow F$ is invertible, $F^{-1}(u)$ $u \in [0, 1]$

$F(X) = U$, then $X = F^{-1}(U)$

1) Compute F^{-1} , maybe analytically

2) Sample $u \in U(0, 1)$

3) Compute $x = F^{-1}(u)$, then $x \sim X$



$$\begin{aligned} X &\sim \text{Exp}(\lambda), p(x) = \lambda e^{-\lambda x} \\ &\quad x \in [0, +\infty) \\ F(x) &= 1 - \exp(-\lambda x) \\ F(x) = U &\Leftrightarrow \begin{matrix} u \sim U \\ 1-u \sim U \end{matrix} \\ x &= -\frac{1}{\lambda} \log(U) \end{aligned}$$

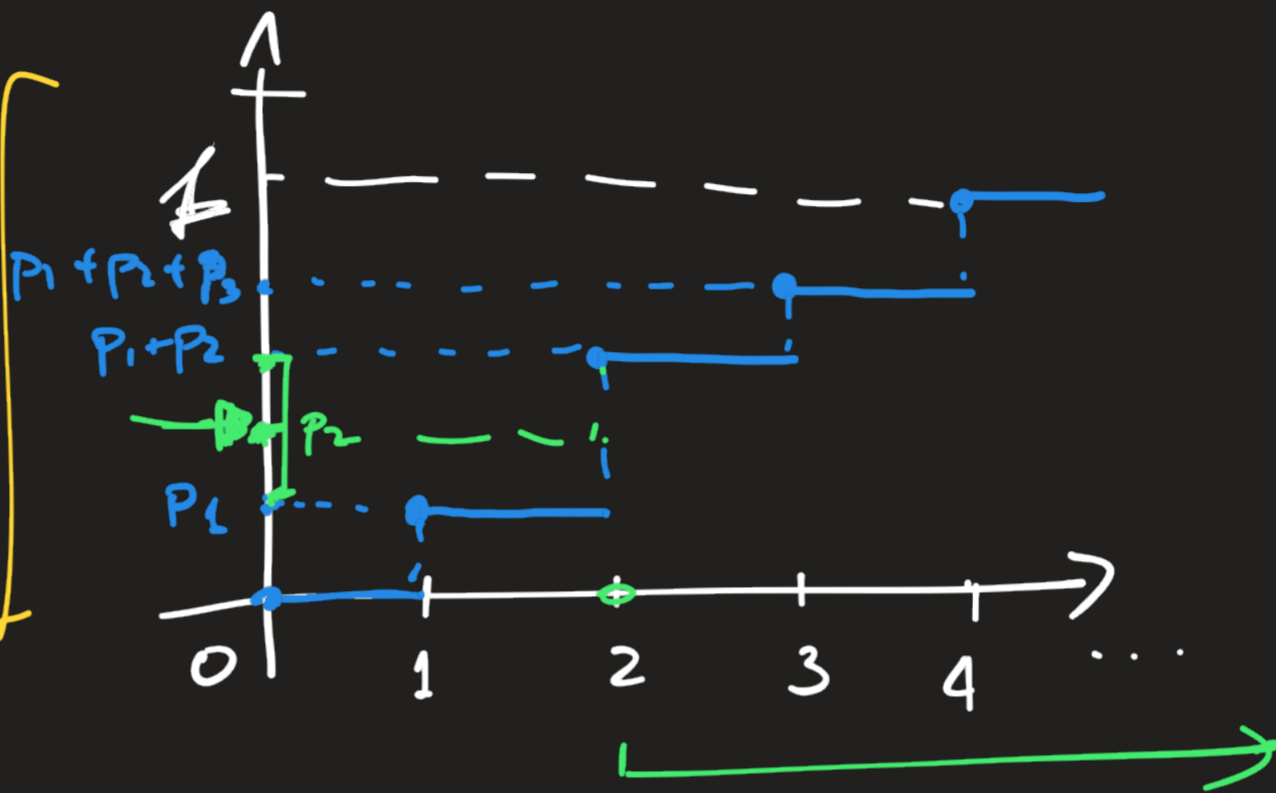
INVERSION METHOD FOR DISCRETE RANDOM VARIABLES

$$X = \{1, \dots, K\} \quad X \sim p: X \rightarrow [0, 1] \quad p(i) = p_i \quad \sum_{i=1}^K p_i = 1$$

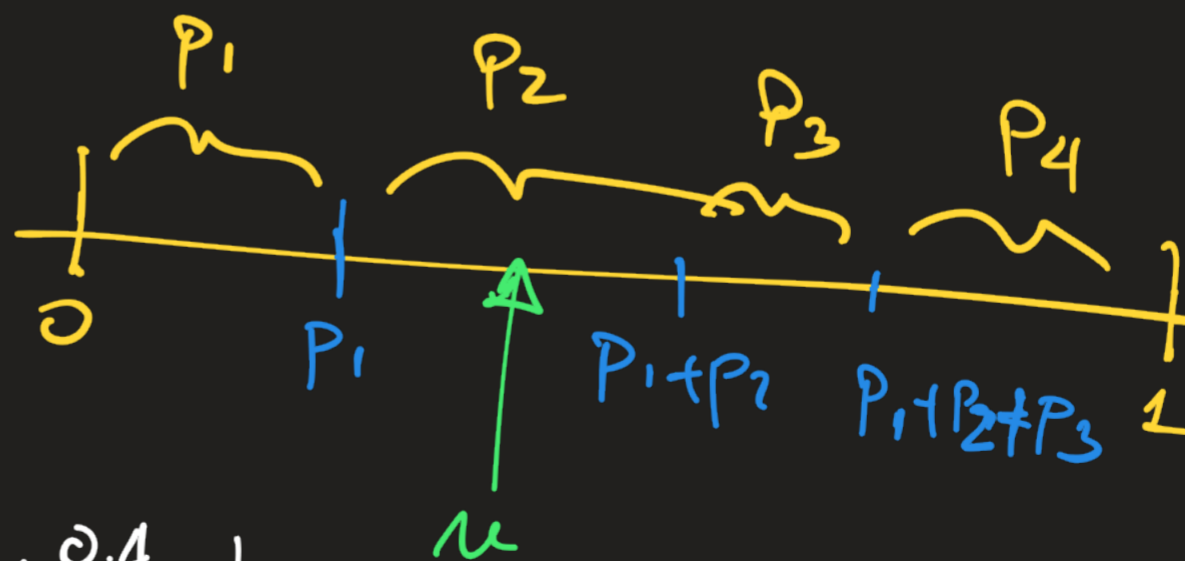
$$F(x) = P(X \leq i) = \sum_{j \leq i} p_j$$

PSEUDO-INVERSE

$$F^{-1}(u) = \inf_{x \in \mathbb{R}} F(x) \geq u$$

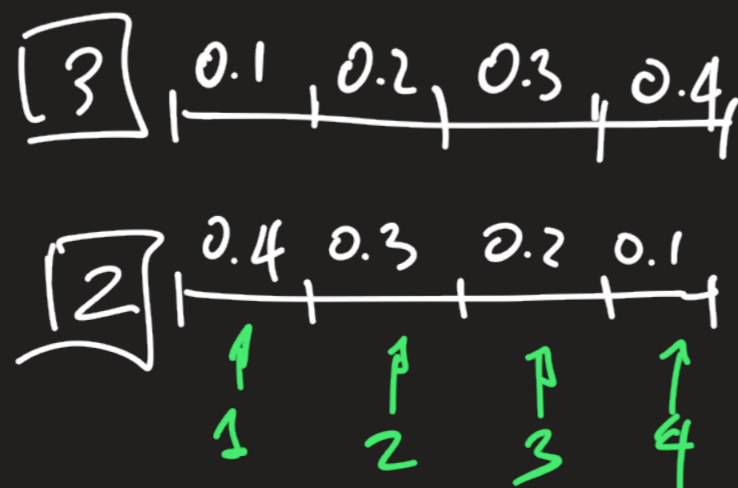


- 1) sample $u \sim U(0, 1)$
- 2) $F^{-1}(u)$ and return it.



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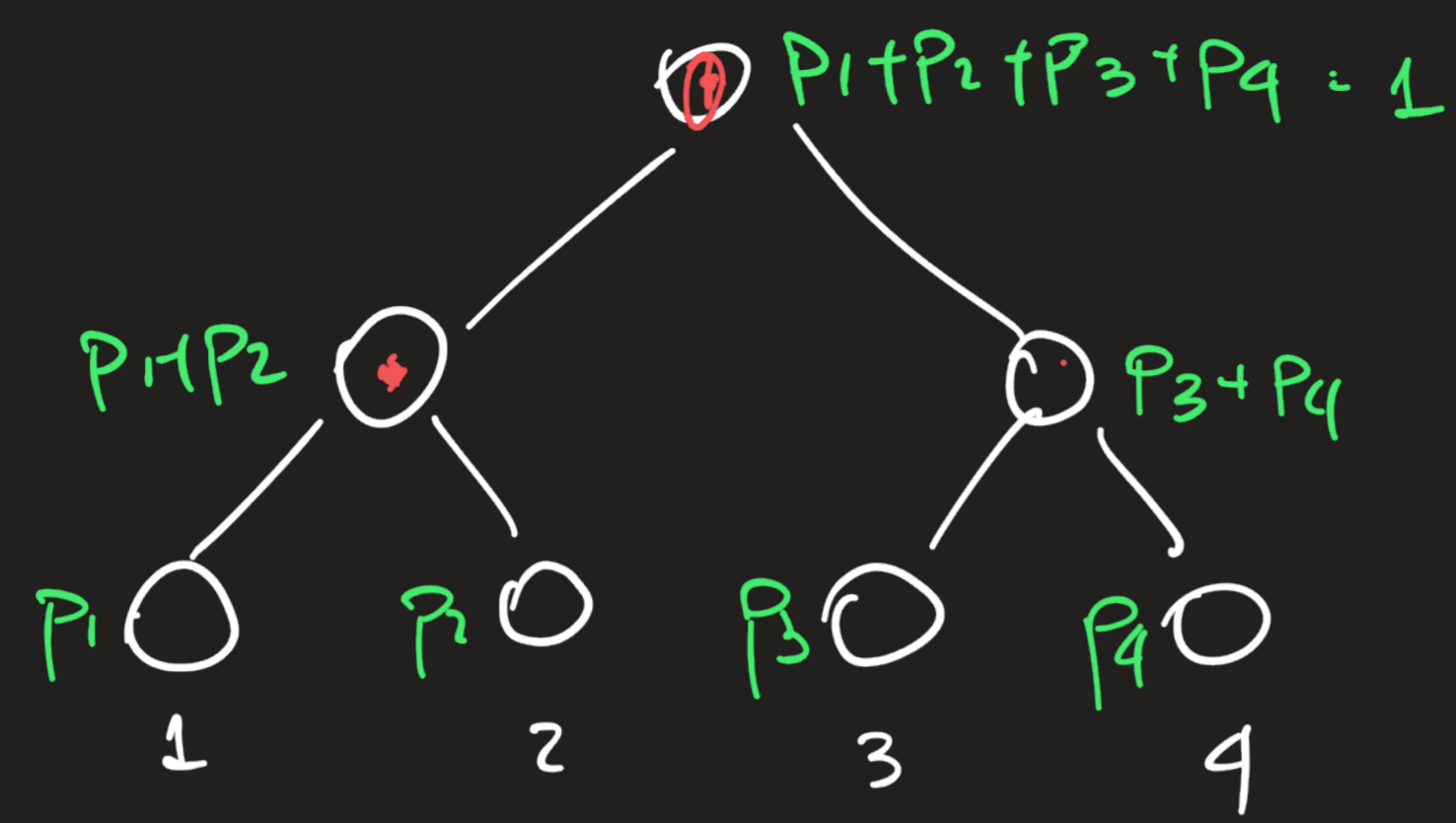
S ← 0
FOR i ← 1 TO K
  S + = P[i]
  IF u ≤ S
    RETURN i
  
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two scenarios. Which is the expected number of comparisons / complexity of the code.

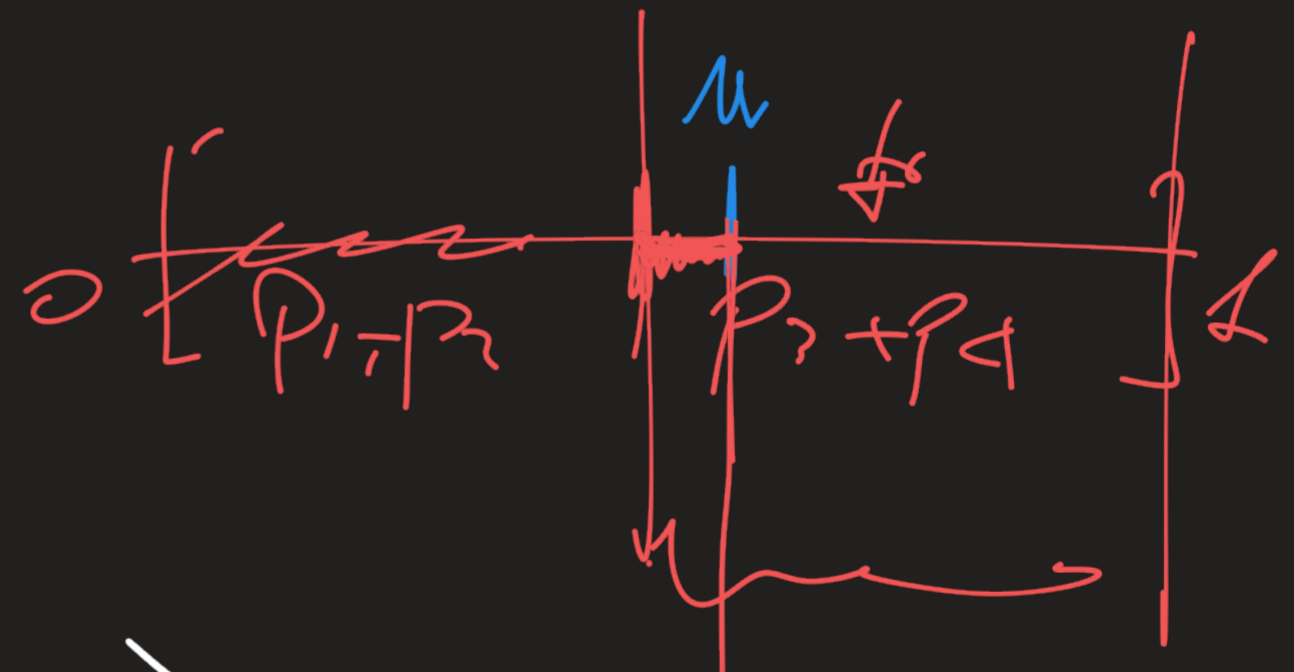
• Alternative speed up, BINARY SEARCH TREE

BINARY SEARCH TREE u ∈ [0, 1]



① IF $u \leq p_1 + p_2$
 ↓
 go left
 else
 $u = u - (p_1 + p_2)$
 go right

$u \in \text{UNIFORM}(0, 1)$
 BT-SAMPLE (node, u)
 IF node is a leaf
 return node.id
 IF $u \leq \text{node.left.value}$
 BT-SAMPLE (node.left)
 ELSE
 $u = u - \text{node.left.value}$
 BT-SAMPLE (node.right)



$O(\log_2 k)$

BOX MULLER TRANSFORM

$X \sim \mathcal{N}(\mu, \sigma^2)$, then $F(x)$ is not analytically tractable

$X, Y \sim \mathcal{N}(0, 1)$ independent

$$p(x, y) = p(x)p(y) = \frac{1}{2\pi} \exp\left(-\frac{x^2 + y^2}{2}\right) \quad \text{joint density}$$

CHANGE INTO POLAR COORDINATES

$$\begin{cases} R = \sqrt{X^2 + Y^2} \\ \theta = \tan^{-1}(Y/X) \end{cases}$$

$$(r, \theta) \sim (X, Y)$$

$$\begin{aligned} R &= \sqrt{x^2 + y^2} \\ \theta &= \tan^{-1}(y/x) \end{aligned}$$

$$\frac{\partial \tan^{-1}(x)}{\partial x} = \frac{1}{1+x^2}$$

$$p(x, y) \sim \underbrace{p(\theta, r)} \cdot |\det(J)| \quad J = \text{jacobian}$$

$$J = \begin{pmatrix} \frac{\partial \theta}{\partial x} & \frac{\partial \theta}{\partial y} \\ \frac{\partial r}{\partial x} & \frac{\partial r}{\partial y} \end{pmatrix} = \begin{pmatrix} -y/(x^2+y^2) & x/(x^2+y^2) \\ x & y \end{pmatrix}$$

$$\det J = -2$$

$$P(\theta, r) = \frac{1}{4\pi} e^{-r/2} = \underbrace{\frac{1}{2\pi}}_{\text{unif. } [0, 2\pi]} \cdot \underbrace{\frac{1}{2} e^{-r/2}}_{\text{exponential rate } \frac{1}{2}}$$

$$p(\theta, r) = p(\theta)p(r)$$

$$p(\theta) = \frac{1}{2\pi} \sim \text{Unif}(0, 2\pi)$$

$$p(r) = \frac{1}{2} e^{-r/2} \sim \text{Exp}\left(\frac{1}{2}\right)$$

Sample $\mu_1, \mu_2 \sim \mathcal{U}(0, 1)$

$$\begin{cases} R = -2 \log \mu_1 \\ \theta = 2\pi \cdot \mu_2 \end{cases}$$

$$\begin{cases} X = \sqrt{R} \cos \theta \\ Y = \sqrt{R} \sin \theta \end{cases}$$

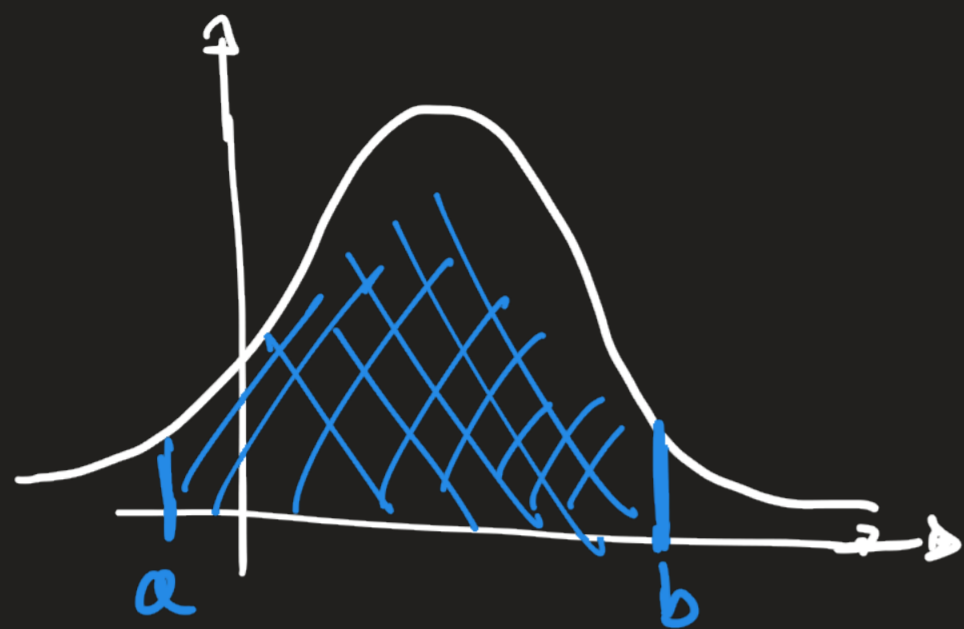
or Box Muller transform

$$X \sim \mathcal{N}(0, 1)$$

$$X' \sim \underline{\mathcal{N}(\mu, \sigma^2)}$$

$$X' = \sigma X + \mu \rightarrow$$

TRUNCATED DISTRIBUTIONS INVERSION METHOD



$$x = F^{-1}(u)$$

CONSTRAINT
 $x \in [a, b]$

$$F(b) - F(a) = P\{X \in [a, b]\}$$

$$p^*(x) = \begin{cases} \frac{P(x)}{F(b) - F(a)} & , x \in [a, b] \\ 0 & , \text{otherwise} \end{cases}$$

$$F^*(x) = \begin{cases} \frac{F(x) - F(a)}{F(b) - F(a)} & , x \in [a, b] \\ 0 & , x < a \\ 1 & , x > b \end{cases}$$

$$u \sim \text{Unif}(0, 1)$$

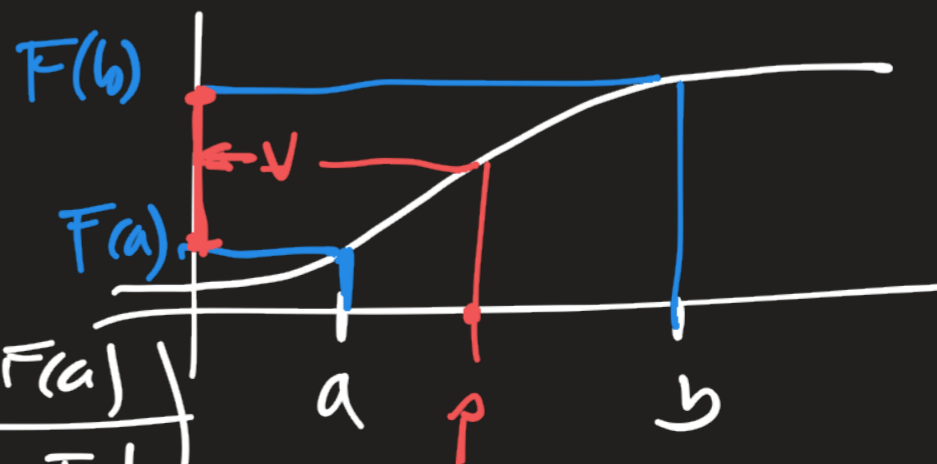
$$V = F(a) + (F(b) - F(a)) \cdot u$$

RETURN $F^{-1}(V)$

$$P(X \leq x) = P(F^{-1}(V) \leq x) = P(V \leq F(x)) =$$

$$= P(F(a) + (F(b) - F(a))u \leq F(x)) = P(u \leq \frac{F(x) - F(a)}{F(b) - F(a)})$$

$$= P(u \leq F^*(x)) = F^*(x)$$



MIXTURE OF R.V.

$$p(x) = \sum_{j=1}^K \pi_j p_j(x)$$

$$\pi_j \in [0, 1] \quad \sum_{j=1}^K \pi_j = 1$$

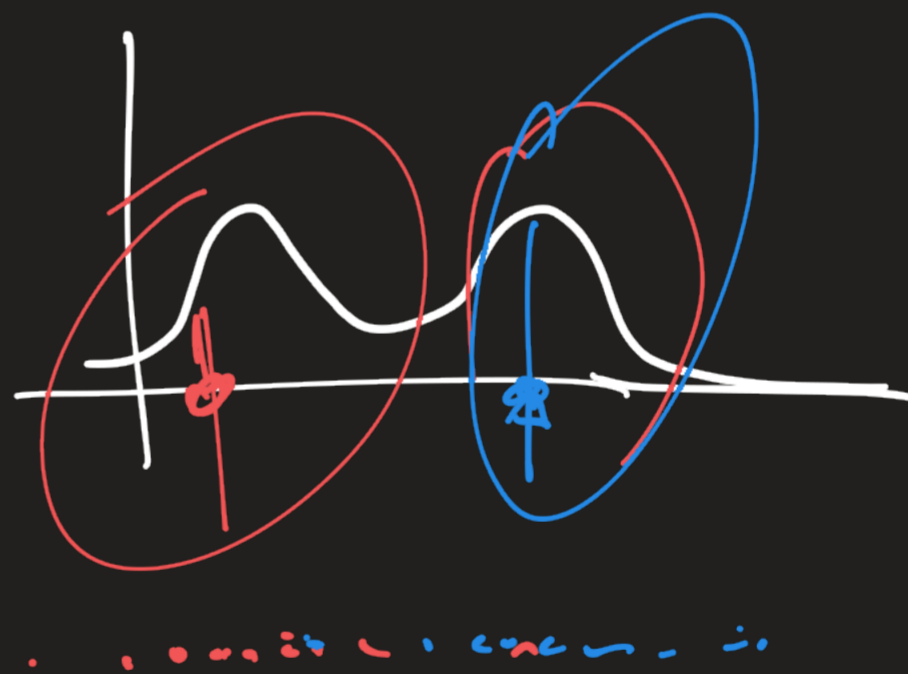
$p_j(x)$ density and so is $p(x)$

Suppose we can sample from p_j
How to sample from $p(x)$?

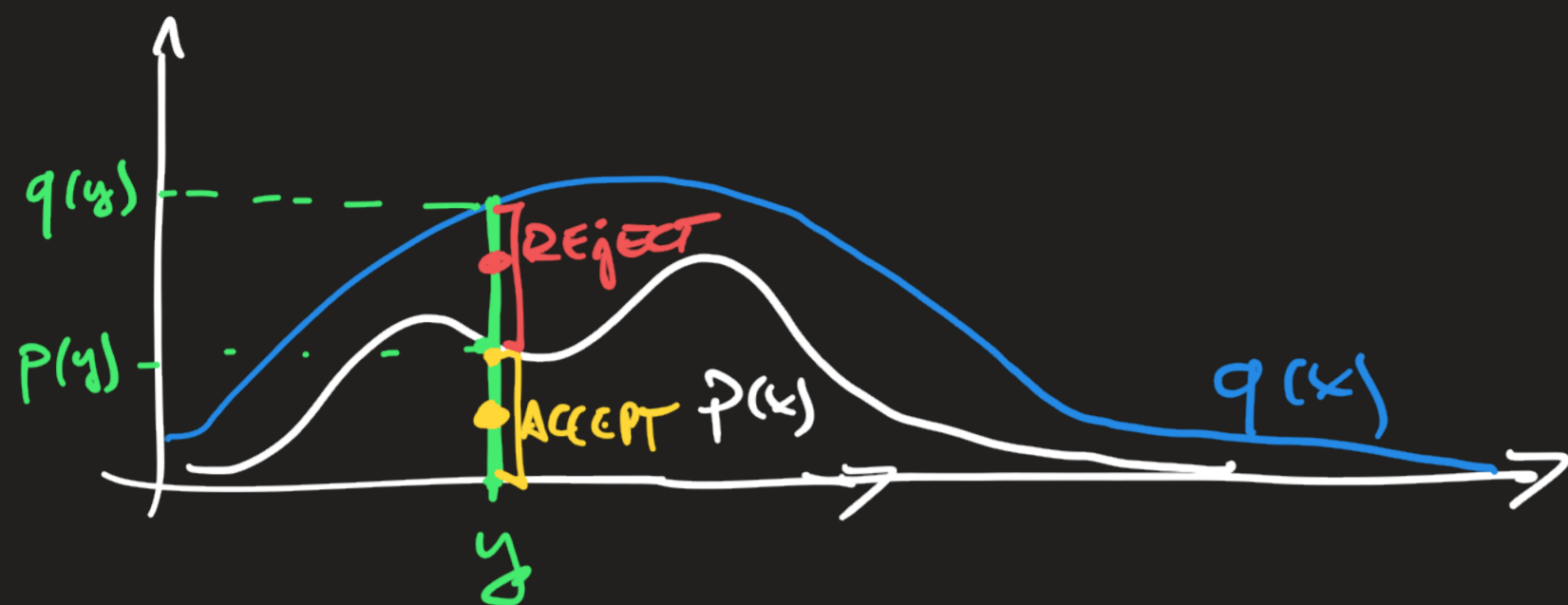
$$p_j(x) = \mathcal{N}(x | \mu_j, \sigma_j^2)$$

ANCESTRAL SAMPLING.

- 1) Sample from $\pi_j = \bar{j}$
- 2) Sample x from $p_{\bar{j}}(x)$



REJECTION SAMPLING.



$$p(x) \leq q(x)$$

$$c = \int_{\mathcal{X}} q(x) \geq \int_{\mathcal{X}} p(x) = 1$$

$$c \geq 1$$

$\pi(x) = \frac{q(x)}{c}$ is a density, a nice density

1) sample y from $\pi(x)$

2) sample $u \sim U(0,1)$ indep. from y

3) IF $u \leq \frac{p(y)}{q(y)}$ RETURN y , OTHERWISE REJECT AND REPEAT

$$q(y) \cdot u \leq p(y)$$

AVERAGE ACCEPTANCE PROB.

$$\int \underbrace{p(\text{accept} | y)}_{p(y)/q(y)} \cdot \pi(y) dy = \int \frac{p(y)}{q(y)} \cdot \pi(y) dy = \frac{1}{c}$$