

# INVARIANT MEASURES AND STEADY STATE

## INVARIANT MEASURE

Consider a CTMC with rate matrix  $Q$  and **finite** state space  $S$ . An invariant measure for the CTMC is a probability distribution  $\pi$  satisfying

$$\pi Q = 0.$$

If  $Q$  is **irreducible** (has a strongly connected graph), then **it has a unique invariant measure**.

## STEADY STATE BEHAVIOUR

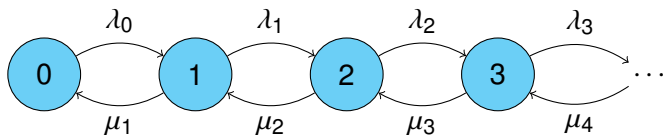
Consider an irreducible CTMC with rate matrix  $Q$  and finite state space  $S$ , and let  $\pi$  be its invariant probability distribution. Then, for each  $s_i, s_j \in S$ ,

$$\lim_{t \rightarrow \infty} P_{ij}(t) = \pi_j.$$

Notice that aperiodicity is not required. Why?

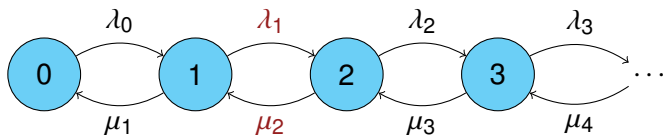
## EXAMPLE: BIRTH-DEATH PROCESS

A birth-death process is a CTMC on  $S = \mathbb{N}$  with birth rate  $\lambda_i$  (from  $i$  to  $i + 1$ ) and death rate  $\mu_i$  (from  $i$  to  $i - 1$ ).



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To derive the steady state  $\pi$ , we can use the fact that the net flow along each **cut** must be zero (why?):

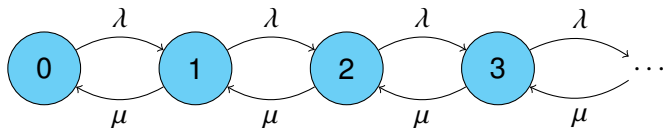
$$\pi_i \lambda_i = \pi_{i+1} \mu_{i+1}$$

Hence we get:

$$\pi_k = \prod_{i=0}^{k-1} \frac{\lambda_i}{\mu_{i+1}} \pi_0; \quad \pi_0 = \left( 1 + \sum_{k=1}^{\infty} \prod_{i=0}^{k-1} \frac{\lambda_i}{\mu_{i+1}} \right)^{-1}$$

## EXAMPLE: BIRTH-DEATH PROCESS

Consider a birth-death process with constant birth rate  $\lambda$  and constant death rate  $\mu$ . It is the model of an **M/M/ $\infty$  queue**.



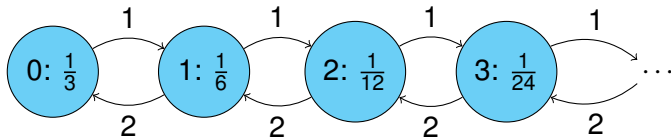
$$\pi_k = \left(\frac{\lambda}{\mu}\right)^k \pi_0; \quad \pi_0 = \left(1 + \sum_{k=1}^{\infty} \left(\frac{\lambda}{\mu}\right)^k\right)^{-1}$$

- If  $\lambda \geq \mu$ , then  $\pi_0 = 0 = \pi_k$ . No state is positive recurrent, there is no invariant measure. The chain escapes to infinity.
- If  $\lambda < \mu$ , then  $\pi_0 = \frac{1-\lambda/\mu}{2-\lambda/\mu}$  and  $\pi_k = \left(\frac{\lambda}{\mu}\right)^k \frac{1-\lambda/\mu}{2-\lambda/\mu}$

## EXAMPLE: BIRTH-DEATH PROCESS

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Assume  $\lambda = 1, \mu = 2$ .



# MATRIX EXPONENTIAL

The solution of the forward Kolmogorov equation  $\frac{dP(t)}{dt} = P(t)Q$ , for a generic CTMC, can be given in terms of the **matrix exponential**

$$P(t) = e^{Qt} = \sum_{n=0}^{\infty} \frac{t^n Q^n}{n!}.$$

However, numerical computation of the series expansion is **numerically unstable**.

## UNIFORMIZATION

A more efficient strategy is to solve the **uniformized CTMC**.

Let  $\lambda \geq \max_i \{-q_{ii}\}$ .

Then one considers a CTMC with jump chain  $Y(n)$  with matrix

$$\Pi = I + \frac{1}{\lambda} Q,$$

and uniform exit rate  $\lambda$ .

The number of fires of this CTMC up to time  $t$  is a Poisson process  $N_\lambda(0, t)$ , and so

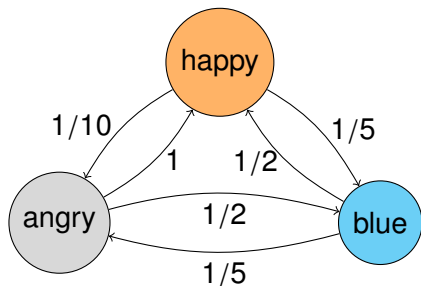
$$X(t) = Y_{N(0,t)} = Y_{y(\lambda t)}.$$

It follows that

$$P(t) = \sum_{n=0}^{\infty} \frac{e^{-\lambda t} (\lambda t)^n}{n!} \Pi^n,$$

which can be truncated above (and below) by bounding the Poisson r.v.

## A SIMPLE EXAMPLE: THE MOOD CHAIN



Upper bound on exit rate: 2

$$P(t) = \sum_{n=0}^{\infty} \frac{e^{-2t}(2t)^n}{n!} \Pi^n$$

$$\Pi = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix} + \frac{1}{2} \begin{pmatrix} -\frac{3}{10} & \frac{1}{5} & \frac{1}{10} \\ \frac{1}{2} & -\frac{7}{10} & \frac{1}{5} \\ 1 & \frac{1}{2} & -\frac{3}{2} \end{pmatrix} = \begin{pmatrix} \frac{17}{20} & \frac{2}{20} & \frac{1}{20} \\ \frac{5}{20} & \frac{13}{20} & \frac{2}{20} \\ \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \end{pmatrix}$$



# OUTLINE

- 1 PRELIMINARIES
  - Exponential Distribution
- 2 CONTINUOUS TIME MARKOV CHAINS
  - Main concepts
  - Poisson Process
  - Time-inhomogeneous rates
- 3 POPULATION CONTINUOUS TIME MARKOV CHAINS
- 4 SIMULATION
  - SSA
  - Next Reaction Method
  - $\tau$ -leaping

# TIME-INHOMOGENEOUS EXPONENTIAL

## DEFINITION

A exponential random variable  $T \sim \text{Exp}(\lambda)$  has time inhomogeneous rate iff  $\lambda = \lambda(t)$  is a function  $\lambda : [0, \infty[ \rightarrow \mathbb{R}^+$ .

- **Cumulative rate** is  $\Lambda(t) = \int_0^t \lambda(s) ds$
- Cdf is  $\mathbb{P}(T < t) = 1 - e^{-\Lambda(t)}$
- Survival probability is  $\mathbb{P}(T > t) = e^{-\Lambda(t)}$

## INVERSION METHOD

One can simulate unidimensional random variables by sampling a uniform r.v.  $U \in [0, 1]$ , and then finding  $t^*$  such that  $t^* = \inf_t \mathbb{P}(T \leq t) = U$ .

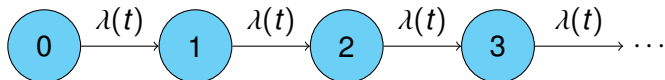
For a time-inhomogeneous  $\text{Exp}(\lambda(t))$ , one has to solve  $e^{-\Lambda(t)} = U$ , iff  $\Lambda(t) = -\log U = \xi$ , with  $\xi \sim \text{Exp}(1)$ .

If  $\lambda$  is constant, then  $\Lambda(t) = \lambda t$ , and one has  $t = -\frac{1}{\lambda} \log(U)$ .

In general, one can either integrate  $\lambda(t)$  or the equivalent ODE  $\frac{d\Lambda(t)}{dt} = \lambda(t)$ , and check for the root of  $\Lambda(t) + \log(U)$  along the solution.

## TIME-INHOMOGENEOUS POISSON PROCESS

A time-inhomogeneous Poisson process  $\mathcal{N}_\lambda(0, t)$ ,  $\lambda = \lambda(t)$ , is a Poisson process with time-varying rate.



It can be shown (same generating function argument as above) that the distribution of  $\mathcal{N}_\lambda(0, t)$  is *Poisson*( $\Lambda(t)$ ), i.e. it is the r.v.

$$\mathcal{Y}(\Lambda(t)) = \mathcal{Y}\left(\int_0^t \lambda(s) ds\right).$$

# TIME-INHOMOGENEOUS CTMC

## TIME-INHOMOGENEOUS CTMC

In general, if the rate matrix  $Q$  of a CTMC depends on time,  $Q = Q(t)$ , then the CTMC is time inhomogeneous.

The probability semigroup depends now also on the initial time:

$$P_{ij}(t_1, t_2) = \mathbb{P}\{X(t_2) = s_j \mid X(t_1) = s_i\}.$$

## FORWARD KOLMOGOROV EQUATION

$$\frac{\partial P(t_1, t_2)}{\partial t_2} = P(t_1, t_2)Q(t_2)$$

## BACKWARD KOLMOGOROV EQUATION

$$\frac{\partial P(t_1, t_2)}{\partial t_1} = -Q(t_1)P(t_1, t_2)$$