#  <br> 993SM - Laboratory of Computational Physics lecture 6 - part 1 April 15, 2020 

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## Numerical integration - II

- other deterministic methods
- more on errors in deterministic methods
- comparison of efficiency between deterministic and Monte Carlo methods
- Central Limit Theorem and related algorithms
M. Peressi - UniTS - Laurea Magistrale in Physics Laboratory of Computational Physics - UnitVI


## Other deterministic methods

## Numerical integration; other deterministic methods:

- in the simplest equally-spaced-point methods, we choose weights to calculate the average of the function:

$$
\begin{gathered}
\int_{a}^{b} f(x) d x \approx F_{N}=\sum_{i=1}^{N} v_{i} f\left(x_{i}\right) \\
\text { rectangular rule: } \quad x_{i}=a+\frac{b-a}{N} i, \quad v_{i}=\frac{b-a}{N} \quad \forall i=1, \ldots N-1 \\
\text { trapezoidal rule: } \quad x_{i}=a+\frac{b-a}{N} i, \quad v_{i}=\frac{b-a}{N} \quad \forall i \neq 1, N ; \quad v_{1}=v_{N}=\frac{b-a}{2 N}
\end{gathered}
$$

(at variance with these methods, in MC methods such as the 'importance sampling', we choose only points, not weights)

## Numerical integration; other deterministic methods:

- in the simplest equally-spaced-point methods, we choose weights to calculate the average of the function:

$$
\int_{a}^{b} f(x) d x \approx F_{N}=\sum_{i=1}^{N} v_{i} f\left(x_{i}\right)
$$

idea: choose not only weights but also points: more degrees of freedom!

$$
x_{i}=?, \quad v_{i}=?
$$

## Another deterministic method:

## Gaussian quadrature - I

Consider $\int_{a}^{b} f(x) d x$ and a function $W(x)$ defined on $[a, b]$ We can always formally write:

$$
\begin{gathered}
\int f(x) d x=\int W(x) F(x) d x \approx \sum_{j=1}^{N} w_{j} F\left(x_{j}\right)=\sum_{j=1}^{N} w_{j} \frac{f\left(x_{j}\right)}{W\left(x_{j}\right)}=\sum_{j=1}^{N} v_{j} f\left(x_{j}\right) \\
\text { with } v_{j}=w_{j} / W\left(x_{j}\right)
\end{gathered}
$$

(This will be convenient in particular if the resulting $F(x)$ is smooth, but not necessarily)
$\left.{ }^{*}\right)$ in general: $w_{j} \neq W\left(x_{j}\right)$

## Another deterministic method:

## Gaussian quadrature - II

Consider $\int_{a}^{b} f(x) d x$ and a function $W(x)$ defined on $[a, b]$
$\int f(x) d x=\int W(x) F(x) d x \approx{ }^{(*} \approx \sum_{j=1}^{N} w_{j} F\left(x_{j}\right)=\sum_{j=1}^{N} w_{j} \frac{f\left(x_{j}\right)}{W\left(x_{j}\right)}=\sum_{j=1}^{N} v_{j} f\left(x_{j}\right)$
For a given $W(x)$, the N points and weights $\left\{x_{j}\right\},\left\{w_{j}\right\}$ can be chosen to make the approximate relationship $\left(^{*}\right.$ ) an exact equality if $F(x)$ is a 2 N -I degree polynomial.

## Another deterministic method:

## Gaussian quadrature - III

Consider

$$
\int_{x_{1}}^{x_{2}} W(x) F(x) d x=\sum_{j=1}^{N} w_{j} F\left(x_{j}\right)
$$

$F(x)$ a $2 \mathrm{~N}-\mathrm{I}$ degree polynomial. Which are the $\mathbf{N}\left\{x_{j}\right\},\left\{w_{j}\right\}$ ?

If there is a set of polynomials $\left\{p_{N}(x)\right\}$ which are orthogonal in the same interval and for the same weight function $W(x)$ :

$$
\left\langle p_{N} \mid p_{N^{\prime}}\right\rangle_{W}=\delta_{N, N^{\prime}} \text {, i.e., } \int_{x_{1}}^{x_{2}} W(x) p_{N}(x) p_{N^{\prime}}(x) d x=\delta_{N, N^{\prime}}
$$

the points $\left\{x_{i}\right\}$ are exactly the roots of the $p_{N}(x)$ polynomials. The weights $\left\{w_{j}\right\}$ are related to them, but in general $w_{j} \neq W\left(x_{j}\right)$.

## Gauss-Legendre quadrature

Consider $\int_{x_{1}}^{x_{2}} W(x) F(x) d x=\sum_{j=1}^{N} w_{j} F\left(x_{j}\right)$
with $F(x)$ a 2 N - I degree polynomial.
If: $W(x)=1$ and $x_{1}=-1, x_{2}=1$,
the Legendre polynomials $\left\{P_{N}(x)\right\}$ defined by:

$$
(j+1) P_{j+1}=(2 j+1) x P_{j}-j P_{j-1}
$$

are orthogonal in $[-1, \mathrm{I}]$ with $W(x)=1$;
$\left\{x_{j}\right\},\left\{w_{j}\right\}$ are such that $P_{N}\left(x_{i}\right)=0$ and

$$
w_{i}=\frac{2}{\left(1-x_{i}^{2}\right)\left[P_{N}^{\prime}\left(x_{i}\right)\right]^{2}}
$$

The first few Legendre polynomials are:

| $n$ | $P_{n}(x)$ |
| ---: | ---: |
| 0 | 1 |
| 1 | $\frac{1}{2}\left(3 x^{2}-1\right)$ |
| 2 | $\frac{1}{2}\left(5 x^{3}-3 x\right)$ |
| 3 |  |
| 4 | $\frac{1}{8}\left(35 x^{4}-30 x^{2}+3\right)$ |
| 5 | legendre polynomials |


polynomials are odd or even in $x \Rightarrow$ roots are even

Legendre polynomials in Physics: examples of applications
I) For a polynomial expansion of a gravitational or coulombic potential:

$$
\frac{1}{\left|\mathbf{x}-\mathbf{x}^{\prime}\right|}=\frac{1}{\sqrt{r^{2}+r^{\prime 2}-2 r r^{\prime} \cos \gamma}}=\sum_{\ell=0}^{\infty} \frac{r^{\ell}}{r^{\ell+1}} P_{\ell}(\cos \gamma)
$$

where $r$ and $r^{\prime}$ are the lengths of the vectors $\mathbf{x}$ and $\mathbf{x}^{\prime}$ respectively and $\gamma$ is the angle between those two vectors.
2) solution of Laplace's equation of the static potential, $\nabla^{2} \Phi(\mathbf{x})=0$, in a charge-free region of space, if the boundary conditions have axial symmetry :
$\theta$ is the angle between the position of the observer and the $\hat{\mathbf{z}}$ axis (the zenith angle); the solution for the potential will be

$$
\Phi(r, \theta)=\sum_{\ell=0}^{\infty}\left[A_{\ell} r^{\ell}+B_{\ell} r^{-(\ell+1)}\right] P_{\ell}(\cos \theta)
$$

3) solving Schrödinger equation in three dimensions for a central force :
the associated Legendre polynomials are derivatives of ordinary Legendre polynomials ( $m \geq 0$ )

$$
P_{\ell}^{m}(x)=(-1)^{m}\left(1-x^{2}\right)^{m / 2} \frac{d^{m}}{d x^{m}}\left(P_{\ell}(x)\right)
$$

The Legendre polynomials are closely related to the spherical harmonics

$$
Y_{\ell, m}(\theta, \phi)=\sqrt{\frac{(2 \ell+1)(\ell-m)!}{4 \pi(\ell+m)!}} P_{\ell}^{m}(\cos \theta) e^{i m \phi} \quad-\ell \leq m \leq \ell
$$

## Gaussian quadrature

In practice, we choose $\mathrm{W}(\mathrm{x})$ and N and use the set of N points and weights $\left\{x_{j}\right\},\left\{w_{j}\right\}$ for the approximate integration:

$$
\int f(x) d x=\int W(x) F(x) d x \approx \sum_{j=1}^{N} w_{j} F\left(x_{j}\right)=\sum_{j=1}^{N} w_{j} \frac{f\left(x_{j}\right)}{W\left(x_{j}\right)}=\sum_{j=1}^{N} v_{j} f\left(x_{j}\right)
$$

## Gauss-Legendre quadrature

For: $x_{1}=-1, \quad x_{2}=1$

| $\bar{N}$ | $\bar{i}$ | $x_{i}$ | $w_{i}$ | $\overline{\text { degree }}$ | degree of the polynomial exactly integrable |
| :---: | :---: | :---: | :---: | :---: | :---: |
| - | - |  |  |  |  |
| 1 | 1 | 0 | 2 | 1 |  |
| $\overline{2}$ | 1 | $\overline{-0.577350269189626}$ | 1 | 3 |  |
|  | 2 | 0.577350269189626 | 1 |  |  |
| $\overline{3}$ | - | $\overline{-0.774596669241483}$ | $\overline{0.555555555555556}$ | 5 |  |
|  | 2 | -0.745064 | 0.8888888888888889 |  |  |
|  | 3 | 0.774596669241483 | 0.555555555555556 |  |  |
| - | 1 | $-0.861136311594053$ | 0.347854845137454 | 7 |  |
|  | 2 | -0.339981043584856 | 0.652145154862546 |  |  |
|  | 3 | 0.339981043584856 | 0.652145154862546 |  |  |
|  | 4 | 0.861136311594053 | 0.347854845137454 |  |  |
|  |  |  |  |  |  |

The integration in an interval $[\mathrm{a}, \mathrm{b}]$ different from $[-\mathrm{I}, \mathrm{l}]$ ("old") can be easily done performing the scaling:

$$
x_{n e w}=\frac{b-a}{2} x_{o l d}+\frac{b+a}{2} \quad \text { and } \quad w_{n e w}=\frac{b-a}{2} w_{o l d}
$$

## Gauss-Legendre quadrature

In case of classical, well known, orthogonal polynomials, ready-to-use subroutines exist for the computation of Abscissas and Weights $\left\{x_{j}\right\},\left\{w_{j}\right\}$
e.g. $\operatorname{GAULEG}\left(x_{1}, x_{2}, x, w, n\right)$ of Numerical Recipes which, given $x_{1}, x_{2}, n$, provides as output the arrays $x(n), w(n)$

## Some programs:

on
\$/home/peressi/comp-phys/VI-integr [do: \$cp/home/peressi/.../VI-integr/* .] and https://moodle2.units.it/

## gauleg-Ilorder.f90

gauleg-others.f90 (generation of points up to 15 points in $[-1, I]$ using GAULEG adapted from "Numerical Recipes" (selfcontained) and some tests for easy-to-integrate functions)

## In the subdirectory: gauss-nr90/

 find the original routine from "Numerical Recipes" and related external routines/modules/interfaces and a main program for test (see following slide)
## Gauss-Legendre from Numerical Recipes

## Use of GAULEG:

In order to use the routines of Numerical Recipes, you have to compile and link the main program with:

- the subroutine gauleg.f90 which gives points and abscissas
- nrtype.f90 containing type declarations;
- nrutil.f90 containing moduli and utilities;
- nr. $\mathbf{f 9 0}$ containing (through the interfaces) the conventions
to call the subroutines with the main program
i) You must first compile these files with the option "-c": this produces .mod and .o (the objects).
ii) In a second step compile the main program.
iii) Finally you link all the files *.o and produce the executable:
g95 -c nrtype.f90 nrutil.f90 nr.f90 gauleg.f90
g95 -c gauleg_nr_test.f90
g95 -o a.out gauleg_nr_test.o nrtype.o nrutil.o nr.o gauleg.o
gauleg.f90 from Numerical Recipes

```
SUBROUTINE gauleg(x1,x2,x,w)
    USE nrtype; USE nrutil, ONLY : arth,assert_eq,nrerror
    IMPLICIT NONE
    REAL(SP), INTENT(IN) :: x1,x2
    REAL(SP), DIMENSION(:), INTENT(OUT) :: x,w
    REAL(DP), PARAMETER :: EPS=3.0e-14_dp
    INTEGER(I4B) :: its,j,m,n
    INTEGER(I4B), PARAMETER :: MAXIT=10
    REAL(DP) :: xl,xm
    REAL(DP), DIMENSION((size(x)+1)/2) :: p1,p2,p3,pp,z,z1
    LOGICAL(LGT), DIMENSION((size(x)+1)/2) :: unfinished
    n=assert_eq(size(x),size(w),'gauleg')
    m=(n+1)/2
    xm=0.5_dp*(x2+x1)
    xl=0.5_dp*(x2-x1)
    z=cos(PI_D*(arth(1,1,m)-0.25_dp)/(n+0.5_dp))
        ...
        ...
    x(1:m)=xm-xl*z
    x(n:n-m+1:-1)=xm+xl*z
    w(1:m)=2.0_dp*xl/((1.0_dp-z**2)*pp**2)
    w (n:n-m+1:-1)=w (1:m)
END SUBROUTINE gauleg
```

nrtype.f90 from Numerical Recipes

```
MODULE nrtype
    INTEGER, PARAMETER :: I4B = SELECTED_INT_KIND(9)
    INTEGER, PARAMETER :: I2B = SELECTED_INT_KIND(4)
    REAL(SP), PARAMETER :: PI=3.141592653589793238462643383279502884197_sp
            ..
            ...
    TYPE sprs2_sp
        INTEGER(I4B) :: n,len
        REAL(SP), DIMENSION(:), POINTER :: val
        INTEGER(I4B), DIMENSION(:), POINTER :: irow
        INTEGER(I4B), DIMENSION(:), POINTER :: jcol
    END TYPE sprs2_sp
    TYPE sprs2_dp
        INTEGER(I4B) :: n,len
        REAL(DP), DIMENSION(:), POINTER :: val
        INTEGER(I4B), DIMENSION(:), POINTER :: irow
        INTEGER(I4B), DIMENSION(:), POINTER :: jcol
    END TYPE sprs2_dp
END MODULE nrtype
```

nr.f90 from Numerical Recipes
MODULE nr INTERFACE

SUBROUTINE gauleg(x1,x2,x,w)
USE nrtype
REAL(SP), INTENT(IN) :: x1,x2
REAL(SP), DIMENSION(:), INTENT(OUT) :: x,w
END SUBROUTINE gauleg
END INTERFACE
! ... the original file contains several other INTERFACES ...
END MODULE nr
nrutil.f90 (Here only for: array_copy, arth, assert_eq, nrerror)
MODULE nrutil
USE nrtype
IMPLICIT NONE
INTEGER(I4B), PARAMETER : : NPAR_ARTH=16,NPAR2_ARTH=8

INTERFACE array_copy
MODULE PROCEDURE array_copy_r, array_copy_d, array_copy_i
END INTERFACE
! ... l'originale contiene ancora molte altre INTERFACEs.... CONTAINS

SUBROUTINE array_copy_r(src,dest,n_copied,n_not_copied) REAL(SP), DIMENSION(:), INTENT(IN) :: src
! .... and many other FUNCTIONs and SUBROUTINEs ....
END MODULE nrutil

## Summary of numerical integration (MC and deterministic) methods

## MC sample mean

$\int_{a}^{b} f(x) d x=(b-a)<f>\approx(b-a) \frac{1}{N} \sum_{i=1}^{N}\left(x_{i}\right)$ with $\left\{x_{i}\right\}$ randomly uniformly distributed in $[a, b]$
$\left(\right.$ it be considered as Importance sampling with $p(x)=\frac{1}{b-a}$ in $\left.[a, b]\right)$

## MC importance sampling

$\int_{a}^{b} f(x) d x=\int_{a}^{b} \frac{f(x)}{p(x)} p(x) d x=<\frac{f(x)}{p(x)}>\int_{a}^{b} p(x) d x \approx \frac{1}{N} \sum_{i=1}^{N} \frac{f\left(x_{i}\right)}{p\left(x_{i}\right)} \int_{a}^{b} p(x) d x$ with $\left\{x_{i}\right\}$ randomly distributed according $p(x)$

Deterministic, equispaced points
$\int_{a}^{b} f(x) d x \approx \sum_{i=1}^{N} v_{i} f\left(x_{i}\right) \quad$ with $x_{i}=a+\frac{b-a}{N} i, v_{i}$ to be determined
Deterministic, non equispaced points
$\int_{a}^{b} f(x) d x \approx \sum_{i=1}^{N} v_{i} f\left(x_{i}\right)$ with $\left\{x_{i}\right\},\left\{v_{i}\right\}$ to be determined

## Numerical integration, deterministic methods: comparison of errors in ID


(double precision needed to appreciate the convergence of Gauss-Legendre numerical estimate)

## Error estimate: comparison between deterministic and MC methods in d-dimension

## Error estimate for numerical integration

 with deterministic methods(Reminder from previous Lecture)

$$
\int f(x) d x=F_{n}+e r r o r
$$

How to evaluate the error? Consider the Taylor expansion of the integrand function and then integrate:

$$
\begin{gathered}
f(x)=f\left(x_{i}\right)+f^{\prime}\left(x_{i}\right)\left(x-x_{i}\right)+\frac{1}{2} f^{\prime \prime}\left(x_{i}\right)\left(x-x_{i}\right)^{2}+\ldots \\
\int_{x_{i}}^{x_{i+1}} f(x) d x=f\left(x_{i}\right) \Delta x+\frac{1}{2} f^{\prime}\left(x_{i}\right)(\Delta x)^{2}+\frac{1}{6} f^{\prime \prime}\left(x_{i}\right)(\Delta x)^{3}+\ldots\left(^{*}\right) \\
\left.\Delta x \equiv x_{25}\right)-x_{i}
\end{gathered}
$$

## Error estimate for numerical integration:

## Rectangular approximation

(Reminder from previous Lecture)

$$
\int_{x_{i}}^{x_{i+1}} f(x) d x \approx f\left(x_{i}\right) \Delta x
$$

Compare $\square$ with (*):
$\int_{x_{i}}^{x_{i+1}} f(x) d x=f\left(x_{i}\right) \Delta x+\frac{1}{2} f^{\prime}\left(x_{i}\right)(\Delta x)^{2}+\frac{1}{6} f^{\prime \prime}\left(x_{i}\right)(\Delta x)^{3}+\ldots$
error
(leading order in $\Delta x$ )

For $n$ intervals $(\Delta x=(b-a) / n):$ error is $n(\Delta x)^{2} \sim 1 / n$

# Numerical integration: multidimensional integrals <br> $$
F=\int_{R} f(x, y) d x d y
$$ 

The rectangular approximation gives $\Delta x \Delta y \sim(\Delta x)^{2} \sim 1 / n$, being $n$ the number of parts (or pairs of points) of the integration domain:

$$
\begin{equation*}
\int_{x_{i}}^{x_{i+1}} \int_{y_{i}}^{y_{i+1}} f(x, y) d x d y \approx f\left(x_{i}, y_{i}\right) \Delta x \Delta y \tag{*}
\end{equation*}
$$

The Taylor expansion of the integrand function gives:
$f(x, y)=f\left(x_{i}, y_{i}\right)+f_{x}^{\prime}\left(x_{i}, y_{i}\right)\left(x-x_{i}\right)+f_{y}^{\prime}\left(x_{i}, y_{i}\right)\left(y-y_{i}\right)+\ldots$
$\int_{x_{i}}^{x_{i+1}} \int_{y_{i}}^{y_{i+1}} f(x, y) d x d y=f\left(x_{i}, y_{i}\right) \Delta x \Delta y+f_{x}^{\prime}\left(x_{i}, y_{i}\right) \frac{(\Delta x)^{2}}{2} \Delta y+f_{y}^{\prime}\left(x_{i}, y_{i}\right) \Delta x \frac{(\Delta y)^{2}}{2}+\ldots\left(^{* *}\right)$
${ }^{(*)}$ against (**) => error
(leading order in $\Delta x$ )
For $n$ intervals: error is $n(\Delta x)^{3} \sim 1 / n^{1 / 2}$

## Numerical integration: multidimensional integrals

Therefore for rectangular approx.:

$$
\mathrm{d}=\mathrm{I}: \text { error } \sim 1 / n \quad \mathrm{~d}=2: \text { error } \sim 1 / n^{1 / 2}
$$

In general:
if the error decreases as $n^{-a}$ for $d=1$, then the error decreases as $n^{-a / d}$ in $d$ dimensions.

Classical formulas with equispaced points: slowly decreasing error for multidimensional integration !

# Numerical integration: error in MC methods 

$\sigma_{n} / \sqrt{n} \approx \sigma_{m} \approx \sigma_{s} / \sqrt{s}$
( $\sigma_{\mathrm{n}}$ is roughly constant; if points are uncorrelated, the variance of the averages goes like $\sim 1 / n^{1 / 2}$ )

- The average function value

$$
\langle f\rangle \equiv \frac{1}{N} \sum_{i=1}^{N} f\left(x_{i}\right)
$$

- The average squared function value

$$
\left\langle f^{2}\right\rangle \equiv \frac{1}{N} \sum_{i=1}^{N} f^{2}\left(x_{i}\right)
$$

- Estimate of the integrand (+/- standard error)

$$
\int f d V \approx V\langle f\rangle \pm V \sqrt{\frac{\left\langle f^{2}\right\rangle-\langle f\rangle^{2}}{N}}
$$

## Numerical integration: errors in multidimensional integrals

| $\bar{d}$ | $\overline{\text { Rect. }}$ |  |  |  |
| :---: | :--- | :--- | :--- | :--- |
| $\overline{T r a p .}$ | $\overline{\text { Simps. }}$ | $\overline{\mathrm{MC}}$ |  |  |
| $\overline{1}$ | $\overline{1 / n}$ | $\overline{1 / n^{2}}$ | $\overline{1 / n^{4}}$ | $1 / n^{1 / 2}$ |
| 2 | $1 / n^{1 / 2}$ | $1 / n$ | $1 / n^{2}$ | $1 / n^{1 / 2}$ |
| 4 | $1 / n^{1 / 4}$ | $1 / n^{1 / 2}$ | $1 / n$ | $1 / n^{1 / 2}$ |
| $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ |  |

if the error decreases as $n^{-a}$ for $d=1$, then the error decreases as $n^{-a / d}$ in $d$ dimensions.
the error for all Monte Carlo integration methods decreases as $n^{-1 / 2}$ independently of the integral.

Monte Carlo convenient for multidimensional integration !

## Summary: advantages of MC integration methods

- convergence as $\sim \mathrm{N}^{1 / 2}$ in any dimension regardless of the smoothness of the integrand
- simplicity: only two simple steps required (namely, producing a set of sampling points and evaluating the integrand function over such points)
- generality: sampling can be used even on domains that do not have a natural correspondence with the 'standard' domain $[0, \mathrm{I}]^{\text {d }}$ and thus are not well-suited to numerical quadrature
- better suited than quadrature for integrands with singularities (importance sampling can handle this problem)
- flexibility: easy to add more points as needed (in the Gaussian quadrature, increasing the accuracy implies doing calculations from scratch)


## The central limit theorem

## The central limit theorem

## the variance of the averages the Gaussian distribution

Consider a continuous random variable $x$ with probability density $f(x)$.
characterized by $\left\langle x^{m}\right\rangle=\int x^{m} f(x) d x$ and ${\sigma_{x}}^{2}=\left\langle x^{2}\right\rangle-\langle x\rangle^{2}$.
Consider $y$ s.t. $\quad y_{n}$ corresponding to the average of $n$ values of $x$ :

$$
y=y_{n}=\frac{1}{n}\left(x_{1}+x_{2}+\ldots+x_{n}\right)
$$

Suppose that we make many measurements of $y$. The variable $y$ is distributed according to a probability density $P(y) \neq f(x)$

## The central limit theorem

The random variable:

$$
y=y_{n}=\frac{1}{n}\left(x_{1}+x_{2}+\ldots+x_{n}\right)
$$

is distributed according to:

## $P(y)$ : gaussian distribution

with:

$$
<y>=<x>
$$

$$
\sigma_{y} \underset{\downarrow}{\approx} \sigma_{x} / \sqrt{n}
$$

(Therefore, the sample mean of a random sample is better than a single observation)
provided $\langle x\rangle$ and $\left\langle x^{2}\right\rangle$ exist (finite) and $n$ is large!

## The central limit theorem

Analogously, is instead of considering the new random variable as the average we consider just the sum:

$$
y=x_{1}+x_{2}+\ldots+x_{n}
$$

it also has a gaussian distribution but with:

$$
\langle y\rangle=n<x\rangle \quad \text { and } \quad \sigma_{y} \approx \sqrt{ } n \sigma_{x}
$$

provided $\langle x\rangle$ and $\left\langle x^{2}\right\rangle$ exist (finite) and $n$ is large!

## The central limit theorem

Note: large enough $n$ needed to obtain the gaussian distribution. Suppose that $\mathrm{f}(\mathrm{x})$ is uniform: e.g., playing dice:
$\mathrm{n}=2$ not enough
$\mathrm{mu}=7.034100$, sigma $^{2}=5.904537$

$\mathrm{n}=100$ OK


## The central limit theorem

The previous example was for UNIFORM distribution (dice) but the central limit theorem work also with random deviates x with NON UNIFORM distribution; e.g. with exponential distribution:

$$
f(x)= \begin{cases}e^{-x}, & \text { if } x \geq 0 \\ 0, i f & x<0\end{cases}
$$

(NOTE: the central limit theorem gives therefore another operative method to generate random numbers with a gaussian distribution)

## ...but sometimes it doesn't work:

## Cauchy-Lorentz

probability density function

$$
\begin{aligned}
& f\left(x ; x_{0}, \gamma\right)=\frac{1}{\pi \gamma\left[1+\left(\frac{x-x_{0}}{\gamma}\right)^{2}\right]} \\
& =\frac{1}{\pi}\left[\frac{\gamma}{\left(x-x_{0}\right)^{2}+\gamma^{2}}\right]
\end{aligned}
$$



The Cauchy-Lorentz distribution is an example of "fat-tailed" distribution. Fat-tailed distributions decay to infinity slower than exponentially. For instance, they can decay with a power law: $f(x) \sim x^{-(1+\alpha)}$ as $x \rightarrow+\infty$ In some cases the expression "fat-tailed" indicates distributions where $0<\alpha<2$.


## ...but sometimes it doesn't work:

## Cauchy-Lorentz

probability density function

$$
\begin{aligned}
& f\left(x ; x_{0}, \gamma\right)=\frac{1}{\pi \gamma\left[1+\left(\frac{x-x_{0}}{\gamma}\right)^{2}\right]} \\
& =\frac{1}{\pi}\left[\frac{\gamma}{\left(x-x_{0}\right)^{2}+\gamma^{2}}\right]
\end{aligned}
$$


$\begin{aligned} & \text { Cumulative } \\ & \text { distribution }:\end{aligned} \frac{1}{\pi} \arctan \left(\frac{x-x_{0}}{\gamma}\right)+\frac{1}{2}$


## .but sometimes it doesn't work:

## Cauchy-Lorentz

probability density function

$$
\begin{aligned}
& f\left(x ; x_{0}, \gamma\right)=\frac{1}{\pi \gamma\left[1+\left(\frac{x-x_{0}}{\gamma}\right)^{2}\right]} \\
& =\frac{1}{\pi}\left[\frac{\gamma}{\left(x-x_{0}\right)^{2}+\gamma^{2}}\right]
\end{aligned}
$$



Mean and variance are not defined
The mean: $\quad \int_{-\infty}^{\infty} x f(x) d x \quad$ which can be rewritten as: $\int_{0}^{\infty} x f(x) d x-\int_{-\infty}^{0}|x| f(x) d x$
is not defined since both terms are infinite; only the Cauchy principal value is defined:
$\lim _{a \rightarrow \infty} \int_{-a}^{a} x f(x) d x$
Without a defined mean, it is impossible to define the variance (but the second moment is defined and it is infinite). Some results in probability theory about expected values, such as the law of large numbers, do not work in such cases.
Also, the mean of a set of random variates drawn from a Cauchy distribution is no better than a single observation, because the chance of including extreme values is high.

## Statistical Properties of Price Returns

## Simulated Returns (Geometric Brownian Motion)



Real Returns (Financial Time-Series)


Cont, Empirical properties of asset returns, stylized facts and statistical issues, 2001

## Statistical Properties of Price Returns

Simulated Returns (Geometric Brownian Motion)


$1^{\text {st }}$ issue
Price returns are not normal

Cont, Empirical properties of asset returns, stylized facts and statistical issues, 2001

## Empirical Distribution of Price Returns

Empirical Distribution of Returns (superposition of all stocks) different time-scales - from 1 minute to 2 hours


Filiasi, PhD Thesis
Cont, Empirical properties of asset returns, stylized facts and statistical issues, 2001

## Empirical Distribution of Price Returns

empirical distribution of price returns

riliasi, PhD Thesis
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Maria Peressi

## END OF THE LECTURE

