

# homework 1 solutions

## Exercise 1

$Y$  is a random variable taking values in  $\{1, \dots, N\}$  and for each value  $P(Y = j) = \frac{1}{N}$ .

$$E[Y] = \sum_{j=1}^N j \frac{1}{N} = \frac{1}{N} \sum_{j=1}^N j = \frac{1}{N} \frac{N(N+1)}{2} = \frac{N+1}{2}$$

## Exercise 2

From the definition of univariate normal distribution:

$$p_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}},$$

where  $\mu \in \mathbb{R}$ ,  $\sigma \in \mathbb{R}_0$ .

From the definition of expectation of continuous random variable:

$$\begin{aligned} E[X] &= \int_{-\infty}^{+\infty} \frac{x}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{+\infty} x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx. \end{aligned}$$

Let  $t := \frac{x-\mu}{\sqrt{2\sigma^2}}$ , hence  $dt = \frac{dx}{\sqrt{2\sigma^2}}$ . Then:

$$\begin{aligned} E[X] &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{+\infty} x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{+\infty} (\sqrt{2\sigma^2}t + \mu) e^{-t^2} \sqrt{2\sigma^2} dt \\ &= \frac{\sqrt{2\sigma^2}}{\sqrt{2\pi\sigma^2}} \left[ \int_{-\infty}^{+\infty} (\sqrt{2\sigma^2}t) e^{-t^2} dt + \int_{-\infty}^{+\infty} \mu e^{-t^2} dt \right]. \end{aligned}$$

Applying the Fundamental Theorem of Calculus we can conclude the proof,

$$\begin{aligned} E[X] &= \frac{1}{\sqrt{\pi}} \left[ \sqrt{2\sigma^2} \left[ -\frac{1}{2} e^{-t^2} \right]_{-\infty}^{+\infty} + \mu \sqrt{\pi} \right] \\ &= \frac{1}{\sqrt{\pi}} [0 + \mu \sqrt{\pi}] = \mu. \end{aligned}$$

## Exercise 3

$X$  and  $Y$  have discrete joint distribution:

$$p(x, y) = \begin{cases} \frac{1}{30}(x + y) & \text{for } x = 0, 1, 2 \text{ and } y = 0, 1, 2, 3 \\ 0 & \text{otherwise} \end{cases}.$$

calculating  $X$  and  $Y$  marginal distributions we obtain

$$p(x) = \sum_{y=0}^3 \frac{1}{30}(x + y) = \frac{1}{30} \left[ \sum_{y=0}^3 x + \sum_{y=0}^3 y \right] = \frac{1}{15}(2x + 3)$$

$$p(y) = \sum_{x=0}^2 \frac{1}{30}(x + y) = \frac{1}{30} \left[ \sum_{x=0}^2 x + \sum_{x=0}^2 y \right] = \frac{1}{10}(y + 1)$$

multiplying them together we have

$$p(x)p(y) = \frac{1}{150}(2x + 3)(y + 1) = \frac{1}{150}(2xy + 2x + 3y + 3)$$

which is different from  $p(x, y)$

#### Exercise 4

From the definition of marginal density for continuous random variables:

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{+\infty} 24xy \, dx \\ &= \int_{-\infty}^0 24xy \, dx + \int_0^{1-y} 24xy \, dx + \int_{1-y}^{+\infty} 24xy \, dx \\ &= 0 + \int_0^{1-y} 24xy \, dx + 0 \\ &= y \int_0^{1-y} 24x \, dx. \end{aligned}$$

Applying the Fundamental Theorem of Calculus we can conclude that:

$$f_Y(y) = \begin{cases} 12y(1 - y)^2 & \text{for } y \in (0, 1) \\ 0 & \text{otherwise} \end{cases}$$

From the definition of conditional density:

$$f_{X|Y=\frac{1}{2}}(x) = \begin{cases} \frac{f_{XY}(x, \frac{1}{2})}{f_Y(\frac{1}{2})} & \text{for } x \in (0, 1), x + \frac{1}{2} < 1 \\ 0 & \text{otherwise} \end{cases}$$

From the solution (1) we can conclude that:

$$f_{X|Y=\frac{1}{2}}(x) = \begin{cases} 8x & \text{for } x \in (0, \frac{1}{2}) \\ 0 & \text{otherwise} \end{cases}$$

## Exercise 5

The joint distribution of  $X$  and  $Y$  is:

$$f(x, y) = \frac{\Gamma(\alpha_1 + \alpha_2 + \alpha_3)}{\Gamma(\alpha_1)\Gamma(\alpha_2)\Gamma(\alpha_3)} x^{\alpha_1-1} y^{\alpha_2-1} (1 - x - y)^{\alpha_3-1}$$

with  $0 < x < 1, 0 < y < 1, x + y < 1$ .

$$f(x) = \int_0^{1-x} f(x, y) dy$$

$$f(x) = \frac{\Gamma(\alpha_1 + \alpha_2 + \alpha_3)}{\Gamma(\alpha_1)\Gamma(\alpha_2)\Gamma(\alpha_3)} x^{\alpha_1-1} \int_0^{1-x} y^{\alpha_2-1} (1 - x - y)^{\alpha_3-1} dy$$

performing the change of variables  $y=(1-x)t$  we obtain:

$$f(x) = \frac{\Gamma(\alpha_1 + \alpha_2 + \alpha_3)}{\Gamma(\alpha_1)\Gamma(\alpha_2)\Gamma(\alpha_3)} x^{\alpha_1-1} (1 - x)^{\alpha_2+\alpha_3-1} \int_0^1 t^{\alpha_2-1} (1 - t)^{\alpha_3-1} dt$$

now since the integral is a Beta function not normalized we have

$$\int_0^1 t^{\alpha_2-1} (1 - t)^{\alpha_3-1} dt = \frac{\Gamma(\alpha_2)\Gamma(\alpha_3)}{\Gamma(\alpha_2 + \alpha_3)}$$

therefore we prove

$$f(x) = \frac{\Gamma(\alpha_1 + \alpha_2 + \alpha_3)}{\Gamma(\alpha_1)\Gamma(\alpha_2 + \alpha_3)} x^{\alpha_1-1} (1 - x)^{\alpha_2+\alpha_3-1} = \text{Beta}(\alpha_1, \alpha_2 + \alpha_3)$$