

# Systems Dynamics

Course ID: 267MI – Fall 2023

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**267MI –Fall 2023**

**Lecture 8**

**Dynamic Models of Stationary Discrete-Time  
Stochastic Processes**

## 8. Dynamic Models of Stationary Discrete-Time Stochastic Processes

8.1 Signal energy & power

8.2 Spectral representation of stationary stochastic processes

8.3 Analysis of dynamic systems driven by input stationary stochastic processes

8.3.1 White noise

8.3.2 Moving average process (MA)

8.3.3 Auto-regressive process (AR)

8.3.4 Auto-regressive and Moving average process (ARMA)

8.3.5 Spectral factorization

# Signal energy & power

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# Signal Energy and Power

- The terms **signal energy** and **signal power** are used to characterize a continuous-time or discrete-time signal.
- They are not actually energy and power measurements, even though the energy and power signal definitions are inspired by expressions used to evaluate energy or power of electrical signals.
- Indeed, the definition of signal energy and power refers to any signal, including signals that take on complex values.

# Energy and power for continuous-time signals

- Consider a generic *deterministic* continuous-time signal  $x(t)$
- Let's define
  - **instantaneous power**

$$P(t) = |x(t)|^2$$

- **energy** over the time period  $[t_0, t_1]$

$$E(t_0, t_1) = \int_{t_0}^{t_1} |x(t)|^2 dt$$

- **average power** over the time period  $[t_0, t_1]$

$$P(t_0, t_1) = \frac{1}{t_1 - t_0} \int_{t_0}^{t_1} |x(t)|^2 dt$$

## Analogy

Consider a current signal  $i(t)$  flowing through a transmission line represented by resistance  $R = 1\Omega$ . Evaluate the energy loss in the line, the instantaneous and average power loss in the line and compare the results with the definitions above.

## Energy and power for continuous-time signals (cont.)

- Extending the time period, let's define
  - **energy** of a deterministic signal  $x(t)$

$$E_{\infty} = \int_{-\infty}^{+\infty} |x(t)|^2 dt$$

- **power** of a deterministic signal  $x(t)$

$$P_{\infty} = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^{+T} |x(t)|^2 dt$$

### Energy and power for a continuous-time deterministic signal

- the deterministic continuous-time signal  $x(t)$  is called an **energy signal** if  $0 < E_{\infty} < \infty$
- the deterministic continuous-time signal  $x(t)$  is called a **power signal** if  $0 < P_{\infty} < \infty$

# Energy and power for discrete-time signals

- Consider a generic discrete-time *deterministic* signal  $\{x(k)\}$
- Let's define
  - **instantaneous power**

$$P(k) = |x(k)|^2$$

- **energy** over the time period  $[k_0, k_1]$

$$E(k_0, k_1) = \sum_{k=k_0}^{k_1} |x(k)|^2$$

- **average power** over the time period  $[k_0, k_1]$

$$P(k_0, k_1) = \frac{1}{k_1 - k_0 + 1} \sum_{k=k_0}^{k_1} |x(k)|^2$$



## Energy and power for discrete-time signals (cont.)

Analogously to the continuous-time case, let's define

- **energy** of a discrete-time signal  $x(k)$

$$E_{\infty} = \sum_{-\infty}^{+\infty} |x(k)|^2$$

- **power** of a discrete-time signal  $x(k)$

$$P_{\infty} = \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{-N}^{+N} |x(k)|^2$$

### Energy and power for a discrete-time deterministic signal

- the discrete-time signal  $x(k)$  is called an **energy signal** if  $0 < E_{\infty} < \infty$
- the discrete-time signal  $x(k)$  is called a **power signal** if  $0 < P_{\infty} < \infty$

# Energy and power for stationary stochastic processes

- Consider now a discrete-time **stationary stochastic process (in weak sense)**. Let's define

- instantaneous power**

$$P(k) = \mathbb{E} \left[ |x(k)|^2 \right]$$

- power** of a stationary (in weak sense) stochastic process  $x(k)$

$$P_\infty = \mathbb{E} \left\{ \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{-N}^{+N} |x(k)|^2 \right\} = \mathbb{E} \left\{ |x(k)|^2 \right\}$$

## Energy and power for a discrete-time stationary (in weak sense) stochastic process

- $E_\infty \rightarrow \infty$
- $P_\infty = \text{var}(x) + (\mathbb{E}[x(k)])^2$

## Remarks

- The power for an energy signal is zero

$$E_{\infty} < \infty \implies P_{\infty} = 0$$

- The energy for a power signal is infinite

$$P_{\infty} < \infty \implies E_{\infty} \rightarrow \infty$$

- Some signals are neither energy nor power signals.
- A signal can't be both an energy signal and a power signal.

- Pure deterministic signals

$$x(k) = A, k = 1, 2, \dots \implies E_{\infty} \rightarrow \infty, P_{\infty} = A^2$$

$$v(k) = A e^{-k}, k = 1, 2, \dots \implies E_{\infty} = \sum_{k=1}^{+\infty} A^2 e^{-2k}, P_{\infty} = 0$$

- Pure stochastic signal

$$\eta(k) = \text{WN}(0, \lambda^2) \implies E_{\infty} \rightarrow \infty, P_{\infty} = \lambda^2$$

# **Spectral representation of stationary stochastic processes**

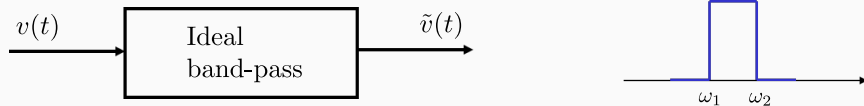
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# Description of stationary stochastic processes

- We described a zero-mean stationary process  $v(t)$  by the **correlation function**:

$$\gamma(\tau) = E[v(t)v(t + \tau)]$$

- However, it is of customary importance to devise a **frequency-based** description of stationary stochastic processes
- Consider the ideal conceptual scheme:



$\text{var} [v(t)] = \lambda^2$   
Average power of the process

$\text{var} [\tilde{v}(t)] = \tilde{\lambda}^2$   
Power contribution in the  
frequency-interval  $[\omega_1, \omega_2]$

# Spectral power density (spectrum)

- Thus, we introduce the **spectral power density (spectrum)** as the Fourier transform of the correlation function:

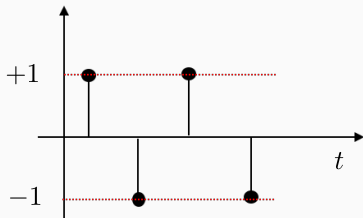
$$\Gamma(\omega) = \mathcal{F}\{\gamma(\tau)\} = \sum_{\tau=-\infty}^{+\infty} \gamma(\tau)e^{-j\omega\tau} \quad (\star)$$

In order the Fourier transform  $(\star)$  to converge, it is necessary that  $\gamma(\tau) \rightarrow 0$  as  $\tau \rightarrow \infty$  “quickly enough”, that is:

$$\sum_{\tau=-\infty}^{+\infty} |\gamma(\tau)| < \infty \quad \longrightarrow \quad \begin{array}{l} \Gamma(\omega) \text{ exists } \forall \omega \\ \Gamma(\omega) \text{ is continuous} \end{array}$$

## Spectral power density (spectrum) (cont.)

- Consider the **discrete-time** periodic sequence:



This sequence is characterized by the **maximum possible** frequency of variation (sign-change at every time-instant).

The smallest period is  $T = 2$  and hence the maximum frequency is  $1/2$ .

Consequently the maximum angular frequency is  $\frac{2\pi}{T} = \pi$ .

$\Gamma$  is evaluated between  $-\pi$  and  $\pi$



- $\Gamma(\omega) \in \mathbb{R}$

In fact  $\gamma(\tau) = \gamma(-\tau)$  (correlation function is even)

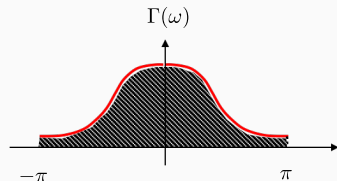
$$e^{-j\omega(-\tau)}\gamma(-\tau) + e^{-j\omega\tau}\gamma(\tau) = \gamma(\tau) [e^{j\omega\tau} + e^{-j\omega\tau}] \in \mathbb{R}$$

- $\Gamma(\omega)$  is even, that is  $\Gamma(\omega) = \Gamma(-\omega)$
- $\Gamma(\omega) \geq 0$
- $\Gamma(\omega)$  is periodic of period  $2\pi$
- $\gamma(\tau) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma(\omega) e^{j\omega\tau} d\omega$

## Properties (cont.)

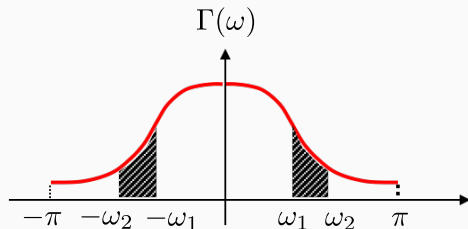
- Thus:

$$\text{var}[v(t)] = \gamma(0) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma(\omega) d\omega$$



Area “below”  $\Gamma(\omega)$  in  $[-\pi, \pi]$  is proportional to the process power and hence to its “variability”.

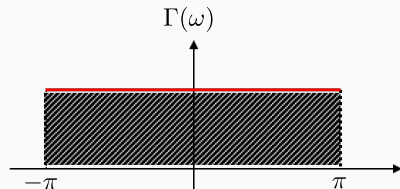
Area below  $\Gamma(\omega)$  in  $[-\omega_2, -\omega_1]$  and  $[\omega_1, \omega_2]$  represents the distribution of the “process variability” in the angular frequency range  $[\omega_1, \omega_2]$ .



## Example

$$v(\cdot) \sim WN(0, \lambda^2)$$

$$\gamma(\tau) = \begin{cases} \lambda^2, & \tau = 0 \\ 0, & \tau \neq 0 \end{cases}$$



$$\Gamma(\omega) = \sum_{\tau=-\infty}^{+\infty} \gamma(\tau) e^{-j\omega\tau} = \lambda^2, \quad \forall \omega \in [-\pi, \pi]$$

$$\gamma(0) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma(\omega) d\omega = \frac{1}{2\pi} \lambda^2 \times 2\pi = \lambda^2$$


Then, in a white process, all angular frequencies contribute in the same way to the overall “process variability”

# Dynamic representations of stationary stochastic processes

- The representations seen up to now are **static**: the process is considered in its entirety (sequence of infinite r.v. from  $-\infty$  to  $+\infty$  ).
- These representations, though correct, are not useful for the solution of the prediction problem.
- We need to devise **dynamic** representations in which it is possible **to relate the future evolution of the process with its past**.
- Special care has to be exercised

## Example 1

Consider a r.v.  $v$  and define the process  $v(t) = v$

  $v(t, s) = v(s)$  (all realizations have constant behavior)

By observing  $v(t, s)$  at some time-instant **the uncertainty disappears** in the sense that the value that the process will take on in the future will not change.

Only **a priori** uncertainty:  
BEFORE the “observation” of the process.

## Example 2

$$v(\cdot) \sim WN(0, \lambda^2)$$

Opposite situation with respect to Example 1: the observation of past values of the process does not help in predicting future values (past and future are not correlated).

The best prediction is the expected value (= 0)

## Dynamic representations of stochastic processes (cont.)

- The two opposite (and kind of “extreme”) examples show that, in the context of the solution of the prediction problem, a more peculiar description of the stochastic process is needed.
- Define a vector space  $\mathcal{G}$  and assume to represent the r.v.  $v(t), v(t-1), v(t-2), \dots$  as vectors in  $\mathcal{G}$ .
- Define

- $\mathcal{H}_t[v]$

Subspace of the past with respect to  $t$  hyper-plane generated by vectors associated with observations  $v(t), v(t-1), v(t-2), \dots$

- $\tilde{\mathcal{H}}_t[v] = \bigcap_t \mathcal{H}_t[v]$

Subspace of the “remote” past

# Dynamic representations of stochastic processes (cont.)

- $\hat{\mathcal{H}}_t[v] = \mathcal{H}_t[v] \setminus \tilde{\mathcal{H}}_t[v]$  such that  $\hat{\mathcal{H}}_t[v] \perp \tilde{\mathcal{H}}_t[v]$

Orthogonal  
complement

Hence, the vectors in  $\hat{\mathcal{H}}_t[v]$  are orthogonal to all vectors in  $\tilde{\mathcal{H}}_t[v]$

- $\tilde{v}(t)$  projection of  $v(t)$  on  $\tilde{\mathcal{H}}_t[v]$
- $\hat{v}(t)$  projection of  $v(t)$  on  $\hat{\mathcal{H}}_t[v]$

Purely deterministic component

Purely non-deterministic  
component

## Wold decomposition

$$\tilde{v}(t) \perp \hat{v}(t)$$

Once  $v$  is decomposed on the two components  $\tilde{v}$ ,  $\hat{v}$ , only the purely non-deterministic component is useful for solving the prediction problem as **the purely deterministic component is perfectly predictable from the past.**



## Dynamic representations of stochastic processes (cont.)

- Then, from now on, let us refer to **dynamic representation of purely non-deterministic processes**
- Consider the hyper-plane  $\hat{\mathcal{H}}_t[v]$  and consider a basis of orthogonal vectors having the same norm

$$\eta(t), \eta(t-1), \eta(t-2), \dots$$

- Such basis is chosen in such a way that:
  - $\eta(t-1), \eta(t-2), \eta(t-3), \dots$  basis for  $\hat{\mathcal{H}}_{t-1}[v]$
  - $\eta(t-2), \eta(t-3), \eta(t-4), \dots$  basis for  $\hat{\mathcal{H}}_{t-2}[v]$
  - ...

vectors  $\eta(t), \eta(t-1), \eta(t-2), \dots$  correspond to r.v. mutually uncorrelated with the same norm

$\eta(\cdot)$  is a white process

## Dynamic representations of stochastic processes (cont.)

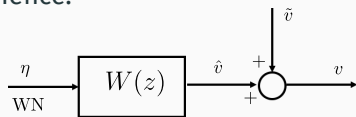
- Project  $v(t)$  on  $\eta(t), \eta(t-1), \eta(t-2), \dots$ 
  - $w(0), w(1), w(2), \dots$  projections components
  - Write the projection  $\hat{v}(t)$  as
$$\hat{v}(t) = w(0)\eta(t) + w(1)\eta(t-1) + \dots$$

$$= \sum_{i=-\infty}^t w(t-i)\eta(i)$$

convolution of  $w(\cdot), \eta(\cdot)$

time-invariance

Hence:



$$W(z) = \mathcal{Z}[w(t)] = \sum_{i=0}^{\infty} w(i)z^{-i}$$

where  $w(t)$  is the **impulse response** of the system

### Therefore

The purely non deterministic component of a **stationary** stochastic process can be seen as the output of a discrete-time dynamic systems driven by a white input process.

### Remark

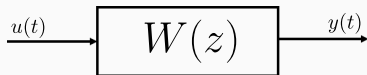
$W(z)$  is not necessarily a rational function of polynomials in  $z$ .

**Analysis of dynamic systems  
driven by input stationary  
stochastic processes**

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## Properties of LTI discrete-time systems

- Consider a linear time-invariant dynamic system with  $W(z)$  as transfer function:



where

$$W(z) = \frac{\gamma_0 z^n + \gamma_1 z^{n-1} + \cdots + \gamma_n}{\alpha_0 z^n + \alpha_1 z^{n-1} + \cdots + \alpha_n}$$

- Hence

$$\begin{aligned} \alpha_0 y(t) + \alpha_1 y(t-1) + \cdots + \alpha_n y(t-n) &= \\ = \gamma_0 u(t) + \gamma_1 u(t-1) + \cdots + \gamma_n u(t-n) \end{aligned}$$

## Properties of LTI discrete-time systems (cont.)

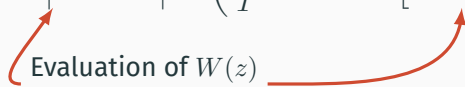
- Recall that, if all **poles** of  $W(z)$  are located strictly inside the unit circle and if

$$u(t) = A \sin \left( \frac{2\pi}{T} t + \beta \right), \quad t = 0, 1, \dots$$

then a **sinusoidal regime** takes place after transient

$$y(t) \simeq A \left| W(e^{j\bar{\omega}}) \right| \sin \left( \frac{2\pi}{T} t + \beta + \arg \left[ W(e^{j\bar{\omega}}) \right] \right)$$

Evaluation of  $W(z)$   
on the unit circle



where

$$\bar{\omega} = \frac{2\pi}{T}$$

## Properties of LTI discrete-time systems (cont.)

- Due to linearity, it follows that

$$y(t) = y_F(t) + y_L(t)$$

Forced response  
(zero initial conditions)

Free response  
(zero input)

- Then

$$y_F(t) = \sum_{j=-t_0}^t w(t-j)u(j)$$

where  $w(t)$  is the impulse response of the system

- Due to the assumed asymptotic stability, we have

$$\lim_{t_0 \rightarrow -\infty} y_L(t) = 0$$

hence

$$\lim_{t_0 \rightarrow -\infty} y(t) = \sum_{j=-\infty}^t w(t-j)u(j)$$

# Properties of LTI discrete-time systems (cont.)

## Then

If all poles of  $W(z)$  are located strictly inside the unit circle, irrespective of the initial conditions, for  $t_0 \rightarrow -\infty$  we have

$$y(t) = \sum_{j=-\infty}^t w(t-j)u(j) = \sum_{i=0}^{+\infty} w(i)u(t-i)$$

**Convolution formula**



## A stochastic process as input of an LTI system

- Now, assume that  $u(\cdot)$  is a stochastic process with zero expected value (that is  $E(u) = 0$ )
- Thus

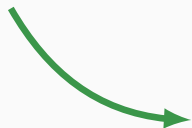
$$\begin{aligned} E[y(t)] &= E \left[ \sum_{i=0}^{\infty} w(i)u(t-i) \right] \\ &= \sum_{i=0}^{\infty} w(i) E[u(t-i)] \\ &= \sum_{i=0}^{\infty} w(i) E[u(t)] \\ &= 0 \end{aligned}$$

stationary

## A stochastic process as input of an LTI system (cont.)

- As we established that  $y(\cdot)$  has zero expected value the covariance functions coincide with the correlation functions.
- Consider input and output values at different time instant  $t_1$  and  $t_2$ :

$$y(t_2) = \sum_{i=0}^{\infty} w(i)u(t_2 - i)$$




$$u(t_1)y(t_2) = \sum_{i=0}^{\infty} w(i)u(t_1)u(t_2 - i)$$

$$y(t_1)y(t_2) = \sum_{i=0}^{\infty} w(i)y(t_1)u(t_2 - i)$$


## A stochastic process as input of an LTI system (cont.)

- Hence, evaluating the expected values

$$E[u(t_1)y(t_2)] = \sum_{i=0}^{\infty} w(i) E[u(t_1)u(t_2 - i)]$$


$$\gamma_{uy}(t_1, t_2) = \sum_{i=0}^{\infty} w(i) \gamma_{uu}(t_1, t_2 - i)$$

$$E[y(t_1)y(t_2)] = \sum_{i=0}^{\infty} w(i) E[y(t_1)u(t_2 - i)]$$


$$\gamma_{yy}(t_1, t_2) = \sum_{i=0}^{\infty} w(i) \gamma_{yu}(t_1, t_2 - i)$$

## A stochastic process as input of an LTI system (cont.)

- As **we assumed** that the process  $u(\cdot)$  is stationary, we obtain:

$$\gamma_{uy}(\tau) = \sum_{i=0}^{\infty} w(i)\gamma_{uu}(\tau - i)$$

$$\gamma_{yy}(\tau) = \sum_{i=0}^{\infty} w(i)\gamma_{yu}(\tau - i)$$

- Then, the correlation function between the input and output processes is the convolution of the impulse response with the auto-correlation function of the input.
- Analogously, the auto-correlation function of the output process is given by the convolution of the impulse response with the input-output correlation function.

# Correlation functions and complex spectra

- Now, introduce the two-sided Z transform of all the possible correlation functions:

$$\Phi_{uu}(z) = \sum_{\tau=-\infty}^{+\infty} \gamma_{uu}(\tau) z^{-\tau}$$

$$\Phi_{yy}(z) = \sum_{\tau=-\infty}^{+\infty} \gamma_{yy}(\tau) z^{-\tau}$$

$$\Phi_{uy}(z) = \sum_{\tau=-\infty}^{+\infty} \gamma_{uy}(\tau) z^{-\tau}$$

$$\Phi_{yu}(z) = \sum_{\tau=-\infty}^{+\infty} \gamma_{yu}(\tau) z^{-\tau}$$

- Recalling the definitions of spectral power density, we have:


$$\Gamma_{uu}(\omega) = \Phi_{uu}(z)|_{z=e^{j\omega}} = \Phi_{uu}(e^{j\omega})$$

$$\Gamma_{yy}(\omega) = \Phi_{yy}(z)|_{z=e^{j\omega}} = \Phi_{yy}(e^{j\omega})$$

where  $\Phi_{uu}(z)$  and  $\Phi_{yy}(z)$  take on the name of **complex spectra** of the input and the output.

# Properties

- $\Gamma_{uu}(\omega), \Gamma_{yy}(\omega) \in \mathbb{R}$ 
  - though in general  $\Gamma_{uy}(\omega), \Gamma_{yu}(\omega) \in \mathbb{C}$
- $\gamma_{uy}(\tau) = \gamma_{yu}(-\tau)$


$$\Phi_{yu}(z) = \Phi_{uy}(z^{-1})$$

$$\Gamma_{yu}(\omega) = \Gamma_{uy}^*(\omega)$$

## Properties (cont.)

- Moreover

$$\gamma_{uy}(\tau) = \sum_{i=0}^{\infty} w(i)\gamma_{uu}(\tau - i) \qquad \Phi_{uy}(z) = W(z)\Phi_{uu}(z)$$



$$\gamma_{yy}(\tau) = \sum_{i=0}^{\infty} w(i)\gamma_{yu}(\tau - i) \qquad \Phi_{yy}(z) = W(z)\Phi_{yu}(z)$$

But

$$\Phi_{yu}(z) = \Phi_{uy}(z^{-1}) = W(z^{-1})\Phi_{uu}(z^{-1})$$

$$\Phi_{yy}(z) = W(z)W(z^{-1})\Phi_{uu}(z^{-1})$$

$$\Phi_{yy}(z) = W(z)W(z^{-1})\Phi_{uu}(z)$$

This is a very important result: we are able to compute the complex spectrum of the output process of a LTI asymptotically stable system driven by a stationary input process.



- Going back to the frequency domain:

$$\begin{aligned}\Gamma_{yy}(\omega) &= \Phi_{yy}(e^{j\omega}) = W(e^{j\omega})W(e^{-j\omega})\Phi_{uu}(e^{j\omega}) \\ &= W(e^{j\omega})W(e^{-j\omega})\Gamma_{uu}(\omega)\end{aligned}$$

$$\Gamma_{yy}(\omega) = |W(e^{j\omega})|^2 \Gamma_{uu}(\omega)$$

This is a very important result as well: we relate spectral power densities with the frequency response.

## More common stationary processes with rational spectra

- White noise
- Moving average process (**MA**)
- Auto-regressive process (**AR**)
- Auto-regressive and Moving average process (**ARMA**)

Common characteristic: as we will see, these processes are generated from a filtered white process.

# **Analysis of dynamic systems driven by input stationary stochastic processes**

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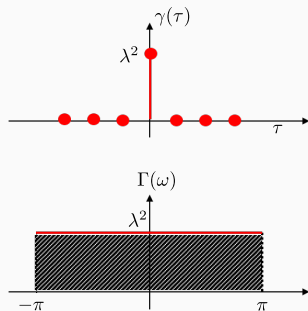
**White noise**

# White noise

$$v(\cdot) \sim WN(0, \lambda^2)$$

$$\gamma(\tau) = \begin{cases} \lambda^2, & \tau = 0 \\ 0, & \tau \neq 0 \end{cases}$$

$$\begin{aligned} \Gamma(\omega) &= \sum_{\tau=-\infty}^{+\infty} \gamma(\tau) e^{-j\omega\tau} \\ &= \gamma(0) = \lambda^2 \end{aligned}$$



As can be seen from the correlation function, in a white process the values of the process in different time-instants do not have any mutual relation, that is, the knowledge of  $v(t)$  is of no help to estimate  $v(\bar{t})$ ,  $\bar{t} \neq t$ .

# White Noise Process – Correlation Function and Spectrum from Data

## Matlab live script

Given a realization of a white noise stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrate how to perform the estimation.



Steps to retrieve the live script:

- Download as a ZIP archive the whole contents of the folder named "**Lecture8**," available in the "**Class Materials**" file area of **the MS Teams course team**, and uncompress it in a preferred folder.
- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_WhiteNoise_CorrelationFunction_Spectrum.mlx');
```

- Explore the live script and run it.

# **Analysis of dynamic systems driven by input stationary stochastic processes**

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**Moving average process (MA)**

- Given a white process  $\eta(\cdot) \sim WN(0, \lambda^2)$
- A **MA process** of order  $n$  (denoted as  $MA(n)$ ) is the process

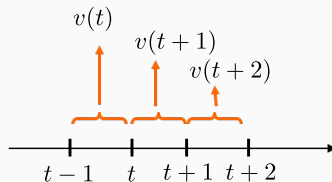
$$v(t) = c_0\eta(t) + c_1\eta(t-1) + c_2\eta(t-2) + \cdots + c_n\eta(t-n)$$

- Hence  $v(t)$  is a linear combination (**average**) of the values taken on by the white process in the time-window from  $t-n$  to  $t$ . When  $t$  increases, this time-window shifts (**moving**).

# Example 1

- Consider the process MA(1) :

$$v(t) = c_0\eta(t) + c_1\eta(t - 1)$$



- Expected value:

$$\begin{aligned} E[v(t)] &= E[c_0\eta(t) + c_1\eta(t - 1)] \\ &= c_0 E[\eta(t)] + c_1 E[\eta(t - 1)] = 0 \end{aligned}$$



## Example 1 (cont.)

- Correlation function (= covariance due to the zero expected value):

$$\gamma(t_1, t_2) = E[v(t_1)v(t_2)]$$

A priori could be non-stationary

- $t_1 = t_2 = t$

$$\begin{aligned}\gamma(t, t) &= E\{v(t)^2\} = E\{[c_0\eta(t) + c_1\eta(t-1)]^2\} \\ &= c_0^2 E\{\eta(t)^2\} + c_1^2 E\{\eta(t-1)^2\} + 2c_0c_1 E\{\eta(t)\eta(t-1)\} \\ &= (c_0^2 + c_1^2) \lambda^2 + 0\end{aligned}$$

## Example 1 (cont.)

- Correlation function (cont.)

- $t_2 = t_1 \pm 1$

$$\begin{aligned}\gamma(t, t \pm 1) &= E \{v(t)v(t \pm 1)\} \\ &= E \{[c_0\eta(t) + c_1\eta(t - 1)][c_0\eta(t \pm 1) + c_1\eta(t - 1 \pm 1)]\}\end{aligned}$$

$$= c_0^2 E \{ \eta(t)\eta(t \pm 1) \} + c_1^2 E \{ \eta(t - 1)\eta(t - 1 \pm 1) \}$$

$$+ c_0c_1 E \{ \eta(t)\eta(t - 1 \pm 1) \} + c_0c_1 E \{ \eta(t - 1)\eta(t \pm 1) \}$$

$0$   $\lambda^2$   $0$   $\lambda^2$

$t_2 = t_1 - 1$   $t_2 = t_1 - 1$   $t_2 = t_1 + 1$   $t_2 = t_1 + 1$

$$= c_0c_1\lambda^2 = \gamma(t, t \pm 1)$$

## Example 1 (cont.)

- Correlation function (cont.)

- $t_2 = t_1 \pm 2$

$$\gamma(t, t \pm 2) = E \{v(t)v(t \pm 2)\} = 0$$

- In general:

$$\gamma(t, t \pm k) = E \{v(t)v(t \pm k)\} = 0, \quad k \geq 2$$

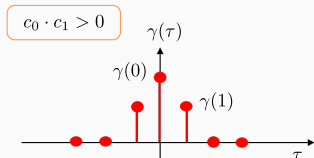
- Note that  $\gamma(t_1, t_2)$  actually is always only a function of  $t_2 - t_1$  and hence the process MA(1) is stationary.

## Example 1 (cont.)

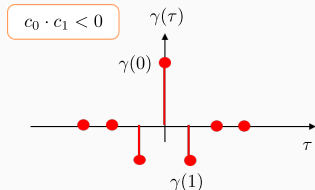
- Thus:

$$MA(1) : \begin{cases} \gamma(0) = (c_0^2 + c_1^2) \lambda^2 \\ \gamma(1) = c_0 c_1 \lambda^2 \\ \gamma(k) = 0, \forall k > 1 \end{cases}$$

Two possible cases:



tendency not to change sign  
in consecutive time-instants

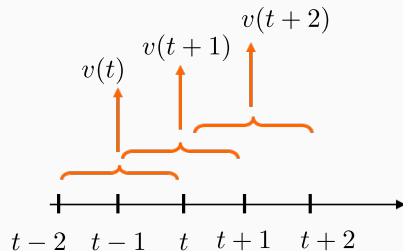


tendency to change sign  
in consecutive time-instants

## Example 2

- Consider the process MA(2) :

$$v(t) = c_0\eta(t) + c_1\eta(t - 1) + c_2\eta(t - 2)$$



- Expected value:

$$\begin{aligned} E[v(t)] &= E[c_0\eta(t) + c_1\eta(t - 1) + c_2\eta(t - 2)] \\ &= c_0 E[\eta(t)] + c_1 E[\eta(t - 1)] + c_2 E[\eta(t - 2)] = 0 \end{aligned}$$

## Example 2 (cont.)

- Correlation function (= covariance due to the zero expected value):  
By a simple algebra analogous to the previous case, we have:

$$\gamma(t, t) = (c_0^2 + c_1^2 + c_2^2) \lambda^2$$

$$\gamma(t, t \pm 1) = (c_0 c_1 + c_1 c_2) \lambda^2$$

$$\gamma(t, t \pm 2) = (c_0 c_2) \lambda^2$$

$$\gamma(t, t \pm k) = 0, \quad \forall k > 2$$

- Here again we notice that  $\gamma(t_1, t_2)$  is always only a function of  $t_2 - t_1$  and hence the process MA(2) is stationary.

- In general, for a process MA( $n$ ) we can prove that:

$$MA(n) : \quad \gamma(\tau) = \begin{cases} 0 & \text{if } |\tau| > n \\ (c_0c_{|\tau|} + c_1c_{|\tau|+1} + \cdots + c_{n-|\tau|}c_n) \lambda^2 & \text{if } n \geq |\tau| \geq 0 \end{cases}$$

- The stationarity should not be surprising: the process is just a **linear combination of values taken on by a stationary process**.
- Clearly:

$$\text{var}[v(t)] = \gamma(0) = (c_0^2 + c_1^2 + \cdots + c_n^2) \lambda^2$$

## MA( $n$ ) process (cont.)

- Let us consider the general expression of a process MA( $n$ ):

$$v(t) = c_0\eta(t) + c_1\eta(t-1) + c_2\eta(t-2) + \dots + c_n\eta(t-n)$$

- By using the unity delay operator  $z^{-1}$  we have:

$$\begin{aligned}v(t) &= c_0\eta(t) + c_1z^{-1}\eta(t) + c_2z^{-2}\eta(t) + \dots + c_nz^{-n}\eta(t) \\ &= (c_0 + c_1z^{-1} + c_2z^{-2} + \dots + c_nz^{-n})\eta(t) \\ &= C(z)\eta(t)\end{aligned}$$

where we set

$$C(z) = c_0 + c_1z^{-1} + c_2z^{-2} + \dots + c_nz^{-n}$$

- Then, the transfer function turns out to be

$$W(z) = \frac{c_0z^n + c_1z^{n-1} + c_2z^{n-2} + \dots + c_n}{z^n}$$

which is **asymptotically stable** ( $n$  poles in the origin).



## Important remark

- In general, stationary processes are characterized via the expected value and the covariance function
- In the case of the process  $MA(n)$  the **expected value** is zero and the **covariance function** is given by

$$MA(n) : \quad \gamma(\tau) = \begin{cases} 0 & \text{if } |\tau| > n \\ (c_0c_{|\tau|} + c_1c_{|\tau|+1} + \dots + c_{n-|\tau|}c_n) \lambda^2 & \text{if } n \geq |\tau| \geq 0 \end{cases}$$

The process is fully characterized  
by the **parameters**

$$c_0, c_1, \dots, c_n, \lambda^2$$

## Important remark (cont.)

- However: the characterization by the parameters  $c_0, c_1, \dots, c_n, \lambda^2$  is **redundant**.

In fact: let

$$\tilde{c}_0 = \alpha c_0, \tilde{c}_1 = \alpha c_1, \dots, \tilde{c}_n = \alpha c_n$$

and consider the process

$$\tilde{v}(t) = \tilde{c}_0 \tilde{\eta}(t) + \tilde{c}_1 \tilde{\eta}(t-1) + \dots + \tilde{c}_n \tilde{\eta}(t-n)$$

where  $\tilde{\eta}(\cdot) \sim \text{WN}(0, \tilde{\lambda}^2)$ .

If

$$\tilde{\lambda}^2 = \frac{\lambda^2}{\alpha^2} \implies \tilde{\gamma}(\tau) = \gamma(\tau)$$

and hence the two processes  $v(t)$   $\tilde{v}(t)$  are not **distinguishable**.

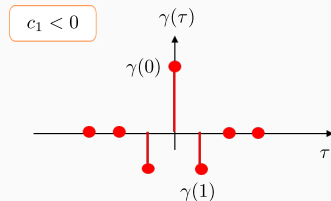
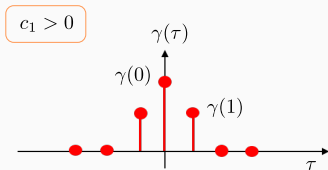
- The redundancy is eliminated assigning one of the parameters. The typical choice is  $c_0 = 1$  and then the process  $\text{MA}(n)$  is written as

$$v(t) = \eta(t) + c_1 \eta(t-1) + \dots + c_n \eta(t-n)$$

## Example 1 (continued)

$$MA(1) : v(t) = \eta(t) + c\eta(t-1)$$

$$MA(1) : \begin{cases} \gamma(0) = (1 + c^2) \lambda^2 \\ \gamma(1) = c \lambda^2 \\ \gamma(k) = 0, \forall k > 1 \end{cases}$$



## Example 1 (cont.)

- Let us determine the spectrum:  
a)

$$\begin{aligned}\Gamma(\omega) &= \sum_{\tau=-\infty}^{+\infty} \gamma(\tau)e^{-j\omega\tau} = \gamma(0) + \gamma(1)e^{-j\omega} + \gamma(-1)e^{j\omega} \\ &= \gamma(0) + \gamma(1)(e^{-j\omega} + e^{j\omega}) = \gamma(0) + 2\gamma(1)\cos\omega \\ &= (1 + c^2)\lambda^2 + 2c\lambda^2\cos\omega \\ &= (1 + c^2 + 2c\cos\omega)\lambda^2\end{aligned}$$

## Example 1 (cont.)

- **b)**

$$v(t) = \eta(t) + c\eta(t-1) \implies W(z) = (1 + cz^{-1})$$

$$\begin{aligned}\Phi(z) &= W(z)W(z^{-1})\lambda^2 = (1 + cz^{-1})(1 + cz)\lambda^2 \\ &= [1 + c^2 + c(z^{-1} + z)]\lambda^2\end{aligned}$$

But  $\Gamma(\omega) = \Phi(e^{j\omega})$

$$\begin{aligned}\Gamma(\omega) &= [1 + c^2 + c(e^{j\omega} + e^{-j\omega})]\lambda^2 \\ &= (1 + c^2 + 2c\cos\omega)\lambda^2\end{aligned}$$

## Example 1 (cont.)

- c)

$$v(t) = \eta(t) + c\eta(t-1) \quad \Longrightarrow \quad \begin{aligned} W(z) &= (1 + cz^{-1}) \\ &= \frac{z+c}{z} \end{aligned}$$

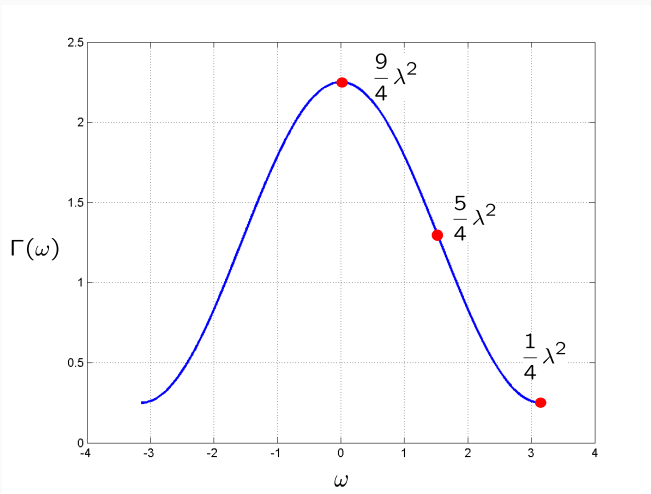
But  $\Gamma_{vv}(\omega) = |W(e^{j\omega})|^2 \Gamma_{\eta\eta}(\omega)$

$$\begin{aligned} \Gamma(\omega) &= \frac{|e^{j\omega} + c|^2}{|e^{j\omega}|^2} \lambda^2 = \frac{|\cos \omega + j \sin \omega + c|^2}{1} \lambda^2 \\ &= [(\cos \omega + c)^2 + (\sin \omega)^2] \lambda^2 \\ &= [(\cos \omega)^2 + c^2 + 2c \cos \omega + (\sin \omega)^2] \lambda^2 \\ &= (1 + c^2 + 2c \cos \omega) \lambda^2 \end{aligned}$$

## Example 1 (cont.)

- if  $c = \frac{1}{2}$

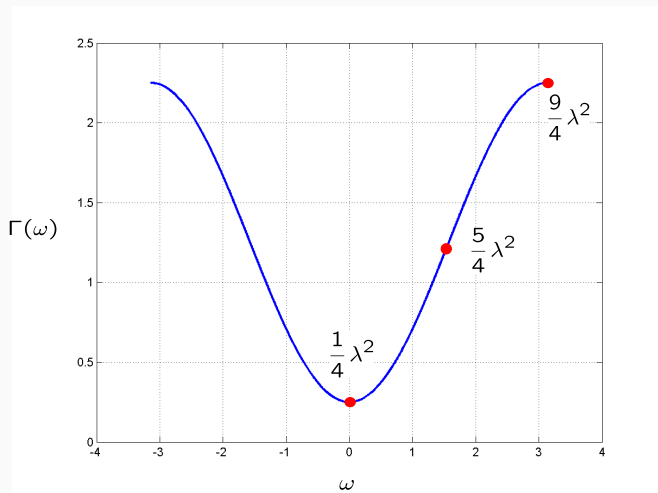
The “variability” of the process is concentrated at lower frequencies



## Example 1 (cont.)

- if  $c = -\frac{1}{2}$

The “variability” of the process is concentrated at higher frequencies





## Matlab live script

Given a realization of a  $MA(1)$  stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrates how to perform the estimation.



Steps to retrieve the live script:

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- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_MA_Processes_AutocorrelationFunction_Spectrum.mlx');
```

- Explore the live script and run it.

## Process $MA(\infty)$

- It is a **mathematical abstraction** of great conceptual importance that will be used in the sequel
- Consider  $\eta(\cdot) \sim WN(0, \lambda^2)$

$$v(t) = \sum_{i=0}^{\infty} c_i \eta(t-i) \quad \text{where } c_0 = 1 \quad (\star)$$

- In order  $(\star)$  to represent a well-defined stationary process it is necessary that

$$\text{var}[v(t)] = \left( \sum_{i=0}^{\infty} c_i^2 \right) \lambda^2 < \infty$$

- It is possible to prove that

$$\sum_{i=0}^{\infty} c_i^2 < \infty \implies \gamma(\tau) < \infty, \quad \forall \tau$$

**Analysis of dynamic systems  
driven by input stationary  
stochastic processes**

---

**Auto-regressive process (AR)**

- Given the white process  $\eta(\cdot) \sim \text{WN}(0, \lambda^2)$
- The AR process of order  $n$  (denoted with  $\text{AR}(n)$ ) is defined as

$$v(t) = a_1v(t-1) + a_2v(t-2) + \dots + a_nv(t-n) + \eta(t)$$

- Hence  $v(t)$  is a linear combination of the values taken on by  $v(t)$  itself in the time-window from  $t-n$  to  $t-1$  plus a white process.

## Example: process AR(1)

- Consider the process AR(1):

$$v(t) = a v(t-1) + \eta(t)$$




$$\begin{aligned}v(t) &= a v(t-1) + \eta(t) \\&= a [a v(t-2) + \eta(t-1)] + \eta(t) \\&= a^2 v(t-2) + a \eta(t-1) + \eta(t) \\&= a^3 v(t-3) + a^2 \eta(t-2) \\&\quad + a \eta(t-1) + \eta(t) \\&= \dots\end{aligned}$$

## Process AR(1) (cont.)


- Then, in general, for a process AR(1) we can write:

$$v(t) = a v(t-1) + \eta(t)$$


$$v(t) = \sum_{i=t_0}^{t-1} a^{t-1-i} \eta(i+1) + a^{t-t_0} v(t_0) \quad (\star)$$

- $(\star)$  is consistent with standard linear systems theory according to which the output response can be decomposed in free response (depending only on initial conditions) and forced response (depending only on the input)
- Define

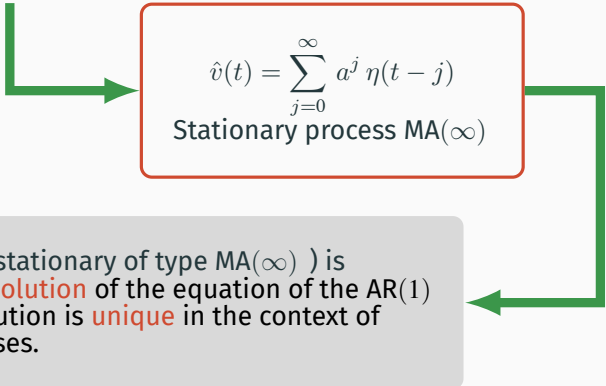
$$\hat{v}(t) = \lim_{t_0 \rightarrow -\infty} v(t)$$

- if  $|a| < 1$    $\hat{v}(t) = \sum_{i=-\infty}^{t-1} a^{t-1-i} \eta(i+1) = \sum_{j=0}^{\infty} a^j \eta(t-j)$

## Process AR(1) (cont.)

- But:

$$|a| < 1 \rightarrow \sum_{j=0}^{\infty} (a^j)^2 < \infty$$


$$\hat{v}(t) = \sum_{j=0}^{\infty} a^j \eta(t-j)$$

Stationary process MA( $\infty$ )

The process  $\hat{v}(t)$  (stationary of type MA( $\infty$ )) is the **steady-state solution** of the equation of the AR(1) process. Such solution is **unique** in the context of stationary processes.

# Analysis of the process $\hat{v}(t)$

- Expected value**

$$v(t) = a v(t-1) + \eta(t)$$


$$E[v(t)] = a E[v(t-1)] + E[\eta(t)]$$

The process  
is stationary


$$\bar{v} = a\bar{v}$$


$$\bar{v} = 0$$


- Variance**

$$E\{[v(t)]^2\} = E\{a^2 [v(t-1)]^2 + [\eta(t)]^2 + 2a v(t-1)\eta(t)\}$$

$$= E\{a^2 [v(t-1)]^2\} + E\{[\eta(t)]^2\} + 2a E[v(t-1)\eta(t)]$$

$$= a^2 E\{[v(t)]^2\} + E\{[\eta(t)]^2\}$$


$$(1 - a^2) E\{[v(t)]^2\} = \lambda^2$$


$$E\{[v(t)]^2\} = \frac{\lambda^2}{(1 - a^2)} \quad (|a| < 1)$$



- **Correlation function**

- Consider  $\tau \geq 0$  ( $\gamma(\tau)$  is even).
- At steady-state the process is MA( $\infty$ ) and hence we can use the general formula

$$\begin{aligned}\gamma(\tau) &= \lambda^2 \sum_{i=0}^{\infty} c_i c_{i+\tau} = \lambda^2 \sum_{i=0}^{\infty} a^i a^{i+\tau} = \lambda^2 a^\tau \sum_{i=0}^{\infty} a^{2i} \\ &= \lambda^2 a^\tau \frac{1}{(1-a^2)} \quad (|a| < 1)\end{aligned}$$

## Analysis of the process $\hat{v}(t)$ (cont.)

- An alternative **algebraic** technique to determine the correlation function is:

$$v(t) = a v(t-1) + \eta(t)$$

$$v(t)v(t-\tau) = a v(t-1)v(t-\tau) + \eta(t)v(t-\tau)$$

$$E[v(t)v(t-\tau)] = a E[v(t-1)v(t-\tau)] + E[\eta(t)v(t-\tau)]$$

Hence:

$$\begin{cases} \lambda^2 & \tau = 0 \\ 0 & \tau > 0 \end{cases}$$

- $\tau > 0 \implies \gamma(\tau) = a \gamma(\tau-1) \implies \gamma(\tau) = a^\tau \gamma(0)$

$$\gamma(0) = \frac{1}{a} \gamma(1)$$

- $\gamma(0) = a \gamma(-1) + \lambda^2 = a \gamma(1) + \lambda^2$

$$\frac{\gamma(1)}{a} = a \gamma(1) + \lambda^2 \implies \gamma(1) = \frac{a}{1-a^2} \lambda^2$$

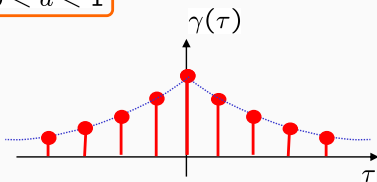
$$\gamma(\tau) = \lambda^2 a^\tau \frac{1}{(1-a^2)} \quad \tau \geq 0$$

## Process AR(1) (cont.)

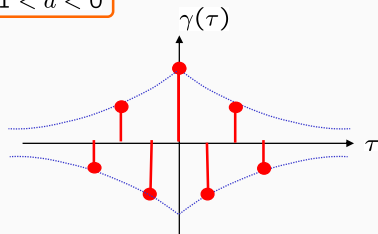
- AR(1) :  $v(t) = a v(t - 1) + \eta(t)$

$$\gamma(\tau) = \lambda^2 a^\tau \frac{1}{(1 - a^2)} \quad \tau \geq 0$$

$$0 < a < 1$$



$$-1 < a < 0$$



- Compared to the case of MA processes, the correlation function vanishes asymptotically (and hence a AR process is “slower” than a MA process).

## Process AR(1) (cont.)

- Let us determine the spectrum:

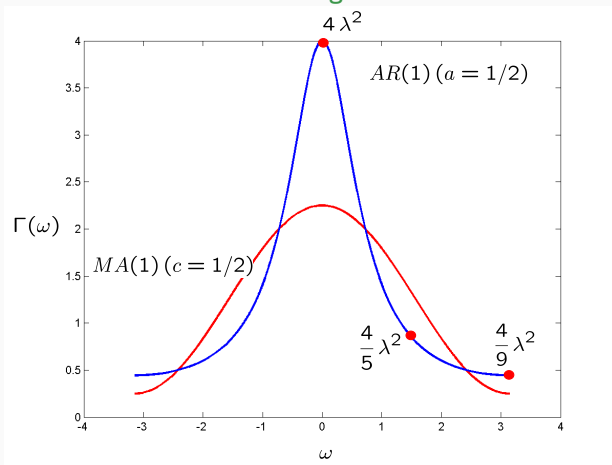
$$v(t) = a v(t-1) + \eta(t) \quad \Longrightarrow \quad W(z) = \frac{1}{1 - az^{-1}} = \frac{z}{z - a}$$

But  $\Gamma_{vv}(\omega) = |W(e^{j\omega})|^2 \Gamma_{\eta\eta}(\omega)$

$$\begin{aligned} \Gamma(\omega) &= \frac{|e^{j\omega}|^2}{|e^{j\omega} - a|^2} \lambda^2 = \frac{1}{|\cos \omega + j \sin \omega - a|^2} \lambda^2 \\ &= \frac{\lambda^2}{[(\cos \omega - a)^2 + (\sin \omega)^2]} \\ &= \frac{\lambda^2}{[(\cos \omega)^2 + a^2 - 2a \cos \omega + (\sin \omega)^2]} \\ &= \frac{\lambda^2}{(1 + a^2 - 2a \cos \omega)} \end{aligned}$$

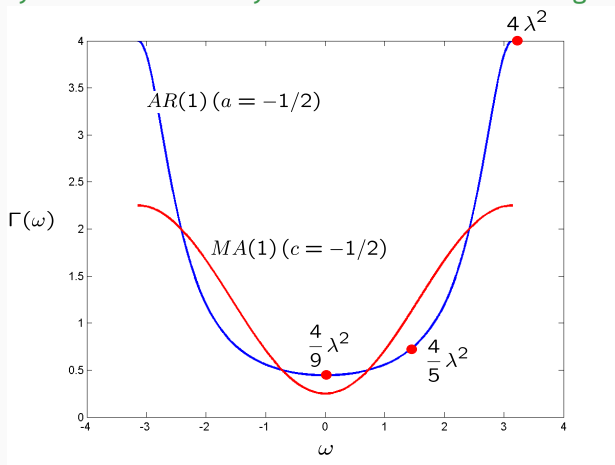
## Process AR(1) (cont.)

- $a = \frac{1}{2}$  The “variability” of the process is more concentrated at low-frequencies and the process is “slower” than the “analogous” MA



## Process AR(1) (cont.)

- $a = -\frac{1}{2}$  The “variability” of the process is more concentrated at high frequency and the frequency distribution is very different from the “analogous” MA



# $AR(1)$ Process – Correlation Function and Spectrum from Data

## Matlab live script

Given a realization of an  $AR(1)$  stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrates how to perform the estimation.



Steps to retrieve the live script:

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- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_AR_1_Processes_AutocorrelationFunction_Spectrum.mlx');
```

- Explore the live script and run it.

## Example: process AR(2)

- Consider the process AR(2)

$$v(t) = a_1 v(t-1) + a_2 v(t-2) + \eta(t)$$



$$v(t)v(t-\tau) = a_1 v(t-1)v(t-\tau) \\ + a_2 v(t-2)v(t-\tau) + \eta(t)v(t-\tau)$$

Hence:

$$\begin{aligned} \gamma(\tau) &= \mathbb{E}[v(t)v(t-\tau)] \\ &= a_1 \mathbb{E}[v(t-1)v(t-\tau)] \\ &\quad + a_2 \mathbb{E}[v(t-2)v(t-\tau)] + \mathbb{E}[\eta(t)v(t-\tau)] \end{aligned}$$

$$\begin{cases} \lambda^2 & \tau = 0 \\ 0 & \tau > 0 \end{cases}$$

- Then

- $\tau > 0 \implies \gamma(\tau) = a_1 \gamma(\tau-1) + a_2 \gamma(\tau-2)$
- $\gamma(0) = a_1 \gamma(-1) + a_2 \gamma(-2) + \lambda^2 = a_1 \gamma(1) + a_2 \gamma(2) + \lambda^2$



# Yule-Walker equations

- The useful equations are:

$$\gamma(2) = a_1 \gamma(1) + a_2 \gamma(0)$$

$$\gamma(1) = a_1 \gamma(0) + a_2 \gamma(1)$$

$$\gamma(0) = a_1 \gamma(1) + a_2 \gamma(2) + \lambda^2$$

- These equations can be organized in matrix form:

$$\begin{bmatrix} a_2 & a_1 & -1 \\ a_1 & a_2 - 1 & 0 \\ 1 & -a_1 & -a_2 \end{bmatrix} \begin{bmatrix} \gamma(0) \\ \gamma(1) \\ \gamma(2) \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \lambda^2 \end{bmatrix} \quad (*)$$

- Then, for given  $a_1$ ,  $a_2$ ,  $\lambda^2$  it is possible to compute  $\gamma(0)$ ,  $\gamma(1)$ ,  $\gamma(2)$  and afterwards proceed in a **recursive** way.
- Equations (\*) are the well-known **Yule-Walker equations** and can be written for any generic AR process.

## $AR(2)$ Process – Correlation Function and Spectrum from Data

### Matlab live script

Given a realization of an  $AR(2)$  stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrates how to perform the estimation.



Steps to retrieve the live script:

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- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_AR_2_Process_AutocorrelationFunction_Spectrum.mlx');
```

- Explore the live script and run it.

## $AR(2)$ Process – The Yule-Walker equations

### Matlab live script

Given the description of an  $AR(2)$  stationary stochastic process, the Yule-Walker equations allow to determine the values of the autocorrelation function. A **Matlab live script** propose an exercise on this topic.



Steps to retrieve the live script:

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- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_AR_2_Process_YuleWalkerEquations.mlx');
```

- Explore the live script and solve the proposed exercise.

## Analysis in case of non-zero mean white noise

$$v(t) = a v(t-1) + \eta(t) \text{ with } \eta(\cdot) \sim WN(\bar{\eta}, \lambda^2)$$

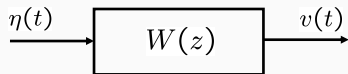
- Consider the process

$$\begin{aligned} E[v(t)] &= a E[v(t-1)] + E[\eta(t)] \\ &= a E[v(t-1)] + \bar{\eta} \end{aligned}$$

- Recalling that  $|a| < 1$  and setting  $E[v(t)] = \bar{v}$

$$\bar{v} = a \bar{v} + \bar{\eta} \implies \bar{v} = \frac{\bar{\eta}}{1-a}$$

- Notice that



$$W(z) = \frac{z}{z-a}$$

$$\bar{v} = W(1) \bar{\eta}$$

Static gain

## Analysis in case of non-zero mean white noise (cont.)

- Let us determine the covariance function:

$$\gamma(\tau) = E \{ [v(t) - \bar{v}][v(t - \tau) - \bar{v}] \}$$

- Introduce the process

$$\tilde{v}(t) = v(t) - \bar{v} \quad \implies \quad v(t) = \tilde{v}(t) + \bar{v}$$

and hence

$$\begin{aligned} \gamma(\tau) &= E \{ [v(t) - \bar{v}][v(t - \tau) - \bar{v}] \} \\ &= E \left\{ [\tilde{v}(t) + \cancel{\bar{v}} - \cancel{\bar{v}}][\tilde{v}(t - \tau) + \cancel{\bar{v}} - \cancel{\bar{v}}] \right\} \end{aligned}$$

The correlation function of the zero-mean process  $\tilde{v}(t)$  coincides with the covariance of the original process.

## AR(1) Process – The Case of Non-zero Mean White Noise

### Matlab live script

Given a realization of a  $MA(1)$  stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrates how to perform the estimation.



Steps to retrieve the live script:

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- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

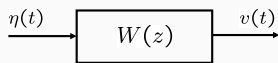
```
open( 'L8_AR_1_Process_NOZeroMean.mlx' );
```

- Explore the live script and run it.

- Consider the process

$$AR(n) : v(t) = a_1 v(t-1) + \dots + a_n v(t-n) + \eta(t)$$

$$A(z)v(t) = \eta(t) \text{ with } A(z) = 1 - a_1 z^{-1} - \dots - a_n z^{-n}$$



$$\begin{aligned} W(z) &= \frac{1}{A(z)} = \frac{1}{1 - a_1 z^{-1} - \dots - a_n z^{-n}} \\ &= \frac{z^n}{z^n - a_1 z^{n-1} - \dots - a_n} \end{aligned}$$

- If the roots of  $A(z)$  (i.e., the poles of  $W(z)$ ) are all strictly located inside the unit-circle, then **in steady-state** we obtain a stationary process equivalent to a process  $MA(\infty)$ .

# **Analysis of dynamic systems driven by input stationary stochastic processes**

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**Auto-regressive and Moving average  
process (ARMA)**



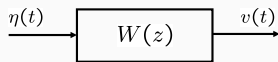
# ARMA processes

- Given a white process  $\eta(\cdot) \sim WN(0, \lambda^2)$
- An ARMA process of order  $n_a, n_c$  (and denoted  $ARMA(n_a, n_c)$ ) is given by

$$v(t) = a_1 v(t-1) + a_2 v(t-2) + \dots + a_n v(t-n) \\ + \eta(t) + c_1 \eta(t-1) + c_2 \eta(t-2) + \dots + c_n \eta(t-n)$$



$$A(z)v(t) = C(z)\eta(t) \quad \text{with} \quad A(z) = 1 - a_1 z^{-1} - \dots - a_{n_a} z^{-n_a} \\ C(z) = 1 + c_1 z^{-1} + \dots + c_{n_c} z^{-n_c}$$



$$W(z) = \frac{C(z)}{A(z)} = \frac{1 + c_1 z^{-1} + \dots + c_{n_c} z^{-n_c}}{1 - a_1 z^{-1} - \dots - a_{n_a} z^{-n_a}}$$

If  $n = \max(n_a, n_c)$  **maximum input/output delay**

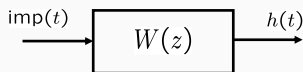
$$W(z) = \frac{z^n + c_1 z^{n-1} + \dots + c_{n_c} z^{n-n_c}}{z^n - a_1 z^{n-1} - \dots - a_{n_a} z^{n-n_a}}$$

## ARMA processes (cont.)

- If the **stability condition is satisfied** (all roots of  $A(z)$  are strictly inside the unity circle) then in **steady-state** we obtain a stationary process equivalent to a process  $MA(\infty)$

$$W(z) = \frac{C(z)}{A(z)} = w_0 + w_1 z^{-1} + \dots + w_i z^{-i} + \dots$$

where  $h(t) = w_t$ ,  $t = 0, 1, \dots$  is the impulse response of the dynamic system

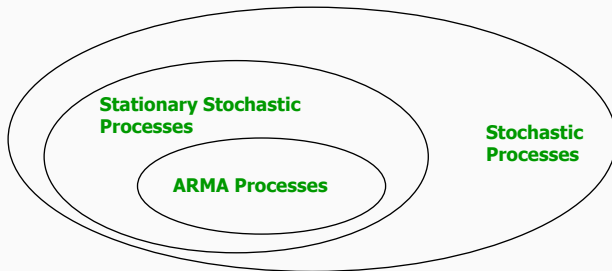


- If the stability condition is satisfied  $\longrightarrow \sum_{j=0}^{\infty} (w_j)^2 < \infty$   $\longleftarrow$  process variance of  $MA(\infty)$  is finite

- With reference to a generic stationary stochastic process:

Difference-equation  $\iff$  Process model  
Stationary solution of difference-equation  $\iff$  Process

- In general:



## Example

- Consider the process  $ARMA(1, 1)$

$$v(t) = \frac{1}{2} v(t-1) + \eta(t) + \frac{1}{3} \eta(t-1), \quad \eta(\cdot) \sim WN(0, 1)$$

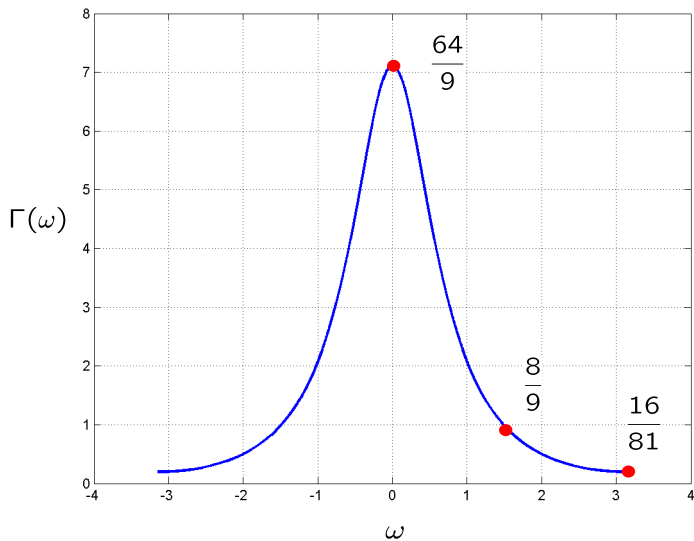
$$\left(1 - \frac{1}{2}z^{-1}\right) v(t) = \left(1 + \frac{1}{3}z^{-1}\right) \eta(t)$$

$$v(t) = \frac{1 + \frac{1}{3}z^{-1}}{1 - \frac{1}{2}z^{-1}} \eta(t) = \frac{z + \frac{1}{3}}{z - \frac{1}{2}} \eta(t)$$

- Let us determine the spectrum:

$$\begin{aligned} \Gamma(\omega) &= |W(e^{j\omega})|^2 \lambda^2 = \frac{|e^{j\omega} + \frac{1}{3}|^2}{|e^{j\omega} - \frac{1}{2}|^2} = \frac{|\cos \omega + j \sin \omega + \frac{1}{3}|^2}{|\cos \omega + j \sin \omega - \frac{1}{2}|^2} \\ &= \frac{(\cos \omega)^2 + \frac{1}{9} + \frac{2}{3} \cos \omega + (\sin \omega)^2}{(\cos \omega)^2 + \frac{1}{4} - \cos \omega + (\sin \omega)^2} = \frac{\frac{10}{9} + \frac{2}{3} \cos \omega}{\frac{5}{4} - \cos \omega} \end{aligned}$$

## Example (cont.)



# ARMA(1, 1) Process – Correlation Function and Spectrum from Data

## Matlab live script

Given a realization of a  $ARMA(1, 1)$  stationary stochastic process, the autocorrelation function and the spectrum can be estimated from the data. A **Matlab live script** illustrates how to perform the estimation.



Steps to retrieve the live script:

- Download as a ZIP archive the whole contents of the folder named "**Lecture8**," available in the "**Class Materials**" file area of **the MS Teams course team**, and uncompress it in a preferred folder.
- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_ARMA1_1_Autocorrelation_Spectrum.mlx');
```

- Explore the live script and run it.

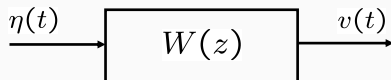
# **Analysis of dynamic systems driven by input stationary stochastic processes**

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**Spectral factorization**

# Spectral factorization

- A few families of stationary stochastic process with rational spectra have been described. The prediction problem will be addressed in the context of these families.
- However: some further discussions are necessary on the **representation of st. processes with rational spectra.**
- Consider



with  $\eta(\cdot) \sim \text{WN}(0, \lambda^2)$  and  $W(z) = \frac{N(z)}{D(z)}$  rational



### Fundamental question:

Does  $\tilde{W}(z)$  exist such that, for a suitable input process,  $\tilde{\eta}(\cdot) \sim \text{WN}(0, \tilde{\lambda}^2)$  we have

$$\gamma_{vv}(\tau) = \gamma_{\tilde{v}\tilde{v}}(\tau), \forall \tau \quad \text{that is} \quad \Gamma(\omega) = \tilde{\Gamma}(\omega), \forall \omega$$

In qualitative terms, does another transfer function exist yielding the same correlation function/spectrum?

- The question is important: in case of existence of such other transfer function  $\tilde{W}(z)$ , this would imply the existence of *more than one* rational representation of the same stochastic process
- Trying to estimate a transfer function (for example, a predictor) on the basis of experimental data would be a ill-posed problem

## Spectral factorization (cont.)

- Then, let us try to understand the mutual relation between transfer functions of **equivalent representations of the same process**.
- More details are needed on the **representation of processes with rational spectra**.
- Recall that  $\Phi(z) = W(z)W(z^{-1})\lambda^2$  where  $\Phi(z)$  is the complex spectrum

### Spectral factorization problem

Given a complex spectrum  $\Phi(z)$ , determine all pairs  $[W(z), \lambda^2]$  such that  $W(z)W(z^{-1})\lambda^2 = \Phi(z)$

## Spectral factorization (cont.)

- Let us now analyze the ways to modify  $W(z)$  without modifying  $\Phi(z)$

**(a)**  $\alpha \cdot W(z)$  with  $\alpha \neq 0 \implies \Phi(z) = \alpha W(z) \alpha W(z^{-1}) \tilde{\lambda}^2$

Choosing  $\tilde{\lambda}^2 = \frac{\lambda^2}{\alpha^2}$  we have that the pairs

$[W(z), \lambda^2]$  and  $[\alpha W(z), \frac{\lambda^2}{\alpha^2}]$  with  $\alpha \neq 0$

have the same  $\Phi(z)$

This result is not surprising: the variance of a stationary process can be changed either by acting on the static gain of the transfer function and on the variance of the input process.

## Spectral factorization (cont.)

(b)  $z^{-k} W(z)$  with  $k \neq 0$   
$$\Phi(z) = z^{-k} W(z) z^k W(z^{-1}) \lambda^2 = W(z) W(z^{-1}) \lambda^2$$

The pairs  $[W(z), \lambda^2]$  and  $[z^{-k} W(z), \lambda^2]$   
have the same  $\Phi(z)$

Also this result is not surprising: multiplying by  $z^{-k}$  means considering realizations delayed by time-steps and this clearly does not alter the probabilistic features of the stochastic process.

### (c) trivial case

$$\frac{(z+a)^n}{(z+a)^n} W(z) \quad \text{with } a \in \mathbb{C}, n \geq 1$$

The pairs  $[W(z), \lambda^2]$  and  $\left[\frac{(z+a)^n}{(z+a)^n} W(z), \lambda^2\right]$

have the same  $\Phi(z)$

## Spectral factorization (cont.)

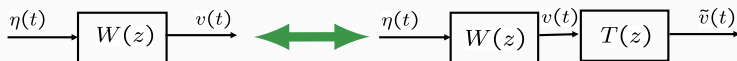
(c) non-trivial case  $\frac{1}{a} \frac{z+a}{z+\frac{1}{a}} W(z)$  with  $a \in \mathbb{C}$

$$\begin{aligned} T(z)T(z^{-1}) &= \varrho \frac{z+a}{z+\frac{1}{a}} \varrho \frac{z^{-1}+a}{z^{-1}+\frac{1}{a}} = \varrho^2 \frac{(z+a)(z^{-1}+a)}{(z+\frac{1}{a})(z^{-1}+\frac{1}{a})} \\ &= \varrho^2 \frac{1+a^2+a(z+z^{-1})}{1+\frac{1}{a^2}+\frac{1}{a}(z+z^{-1})} = \varrho^2 a^2 \frac{1+a^2+a(z+z^{-1})}{1+a^2+a(z+z^{-1})} = \varrho^2 a^2 \end{aligned}$$

Choosing  $\varrho = \frac{1}{a}$  the pairs  $[W(z), \lambda^2]$  and  $\left[ \frac{1}{a} \frac{z+a}{z+\frac{1}{a}} W(z), \lambda^2 \right]$  have the same  $\Phi(z)$

## An “all-pass” filter

- Then, if  $T(z) = \frac{1}{a} \frac{z + a}{z + \frac{1}{a}}$



in the sense that the spectra of  $v(t)$  and of  $\tilde{v}(t)$  coincide.

Therefore, canceling a pole (zero) with a reciprocal zero (pole) leaves the spectrum unchanged (except for a possible multiplicative constant)

## Spectral factorization (cont.)

- The problem we are addressing is the one of **guaranteeing the uniqueness of the representation of the stationary stochastic process** that is, that there exists a **unique** transfer function  $W(z)$  for which the process can be represented as the output of a linear system with transfer function  $W(z)$  and a white process as input.
- Clearly, there are many ways to constrain the representation to be unique. The ways we are considering are the ones that will be useful in the context of the solution of the prediction problem.



## Spectral factorization (cont.)

- With reference to cases (a), (b), (c) previously addressed:
  - given  $W(z) = \frac{N(z)}{D(z)}$
  - (a): it is sufficient to set some parameter. We impose that  $N(z)$  and  $D(z)$  are **monic polynomials**
  - (b): we impose that  $N(z)$  and  $D(z)$  have the **degree a-priori set** (for example, the same degree)
  - (c): we impose that  $N(z)$  and  $D(z)$  are **co-prime** and that **all zeros and poles are inside the unit circle**

## Spectral factorization theorem

Given a process with rational spectrum  $\Phi(z)$ , there exists one and only one representation of the process as the output of a linear system driven by a white process and with transfer function  $W(z) = \frac{N(z)}{D(z)}$  if the following conditions are imposed on  $W(z)$  :

- $N(z)$  and  $D(z)$  monic, co-prime and of the same degree
- all roots of  $N(z)$  (zeros of  $W(z)$ ) have  $|\cdot| \leq 1$
- all roots of  $D(z)$  (poles of  $W(z)$ ) have  $|\cdot| < 1$

# The Spectral Factorization Theorem

## Matlab live script

Given the spectrum  $\Phi(z)$  of a stationary stochastic process, a **Matlab live script** illustrates how to perform the spectral factorization.



Steps to retrieve the live script:

- Download as a ZIP archive the whole contents of the folder named **"Lecture8,"** available in the **"Class Materials"** file area of **the MS Teams course team**, and uncompress it in a preferred folder.
- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_SpectralFactorization.mlx');
```

- Explore the live script and run it.

# The Spectral Factorization Theorem: an Exercise

## Matlab live script

Given the spectrum  $\Phi(z)$  of a stationary stochastic process, perform the spectral factorization.

A **Matlab live script** propose an exercise on solving the spectral factorization problem using MATLAB.

Steps to retrieve the live script:

- Download as a ZIP archive the whole contents of the folder named "**Lecture8**," available in the "**Class Materials**" file area of **the MS Teams course team**, and uncompress it in a preferred folder.
- Add the chosen folder and subfolders to the Matlab path.
- Open the live script using the Matlab command:

```
open('L8_HandsON_SpectralFact.mlx');
```

- Explore the live script and solve the spectral factorization problem.



**267MI –Fall 2023**

**Lecture 8**

**Dynamic Models of Stationary Discrete-Time  
Stochastic Processes**

**END**