

CLASSIFICATION

dataset $D = (x_i, y_i), i=1, \dots, N$

- DISCRIMINANT FUNCTIONS

$$f(x) \in \{1, \dots, K\}$$

- DISCRIMINATIVE APPROACH

$$P(C_k | x) = \underbrace{f}_{h(x)}(\omega^T \phi(x))$$

- GENERATIVE APPROACH

$$P(C_k | x) = \frac{P(x | C_k) P(C_k)}{P(x)}$$

ACTIVATION FUNCTION

f^{-1} LINK FUNCTION

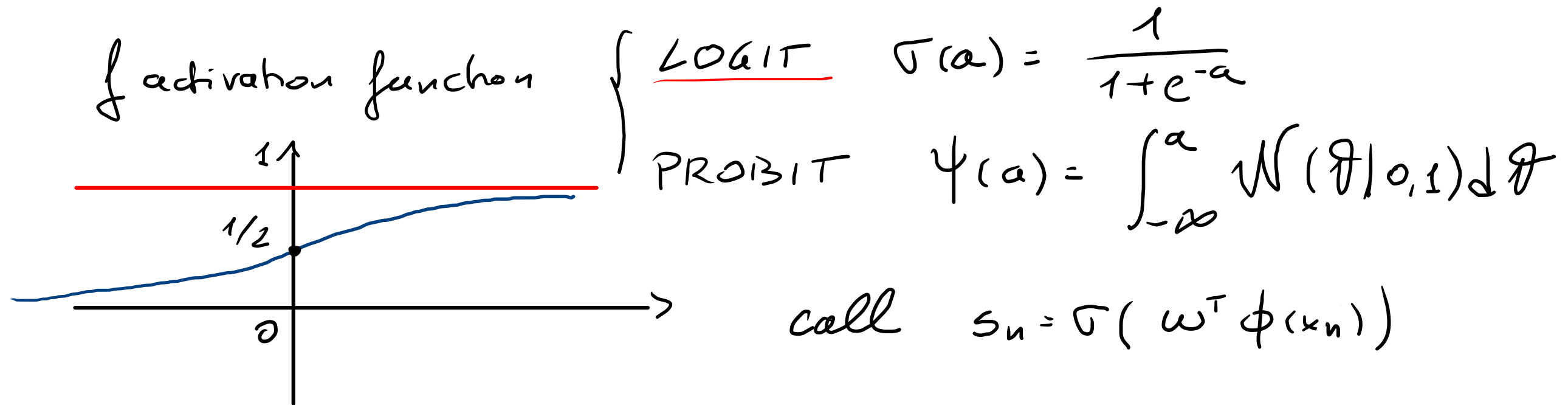
2 class problem $y_n \in \{0, 1\}$

$K \geq 2$ class problem $y_n = (y_{nj})_{j=1, \dots, K}$ ONE HOT ENCODING $y_{nj} \in \{0, 1\}, \sum_j y_{nj} = 1$

LOGISTIC REGRESSION

(x_n, y_n) , $n=1, \dots, N$; $\phi(x) = (\phi_0(x), \dots, \phi_{m-1}(x))$ basis functions

$$P(C_1 | x) = f(\omega^T \phi(x))$$



• NOISE: BERNOLLI $P(y_n | x_n) = s_n^{y_n} (1-s_n)^{1-y_n}$

• LIKELIHOOD $P(\underline{y} | \underline{x}) = \prod_{n=1}^N s_n^{y_n} (1-s_n)^{1-y_n}$

$$E(\omega) = -\frac{1}{N} \sum_{i=1}^N \log p(\underline{y} | \underline{x}) = -\frac{1}{N} \sum_{i=1}^N y_i \log s_i + (1-y_i) \log (1-s_i)$$

$$\nabla E(\omega) = \frac{1}{N} \sum_{i=1}^N (s_i - y_i) \phi(x_i) \quad \left(\nabla E(\omega) = 0 \text{ has no analytic solution} \right)$$

$$\omega_{ML} = \arg \min_{\omega} E(\omega) \quad (\text{convex})$$

• GRADIENT DESCENT: $\omega_{n+1} = \omega_n - \eta_n \nabla E(\omega_n)$

• SEPARATING HYPERPLANE: $\omega_{ML}^T \phi(x) = 0$

for Bayes decision rule: class 1 iff $\sigma(\omega_{ML}^T \phi(x)) \geq \frac{1}{2}$

MULTICLASS PROBLEM: (x_n, y_n) , with $y_n = (y_{n1}, \dots, y_{nK})$ K classes

$$p(C_k | x) = \sigma_k(W^T \phi(x))$$

W^T is a $K \times M$ matrix

where $\sigma_k(\underline{a}) = \frac{\exp(a_k)}{\sum_j \exp(a_j)}$ is the SOFTMAX

$$\text{Now } \bar{E}(w) = -\frac{1}{N} \sum_{n=1}^N \sum_{j=1}^K y_{nj} \log s_{nj}, \quad s_{nj} = \sigma_j(W^T \phi(x_n))$$

$$\underbrace{\nabla_{w_{\cdot j}} \bar{E}(w)}_{\text{GRADIENT FOR CLASS } j \text{ WEIGHTS}} = \frac{1}{N} \sum_{n=1}^N (s_{nj} - y_{nj}) \phi(x_n)$$

GRADIENT FOR
CLASS j WEIGHTS

LAPLACE APPROXIMATION

1. 1D case

$$p(z) = \frac{1}{Z} f(z) \quad , \quad Z = \int f(z) dz$$

- FIND A MODE OF $f(z)$, say z_0 .
- MATCH THE CURVATURE OF f AT z_0 WITH A NORMAL DISTRIBUTION.

Let z_0 a point s.t. $\frac{d}{dz} f(z_0) = 0$, and z_0 maximum.

We TAYLOR EXPAND $\log f(z)$ around z_0 :

$$\log f(z) \approx \log f(z_0) - \frac{1}{2} A (z - z_0)^2$$

$$A = - \frac{d^2}{dz^2} \log f(z_0), \quad A > 0 :$$

$$f(z) \approx f(z_0) \cdot \exp\left(-\frac{1}{2} A (z - z_0)^2\right)$$

APPROX p by a Gaussian $q(z)$
 $q(z) \sim \mathcal{N}(z|z_0, A^{-1}) :$
 $\left(\frac{A}{2\pi}\right)^{1/2} \exp\left(-\frac{1}{2} A (z - z_0)^2\right)$

$$p(z) = \frac{1}{Z} f(z) \approx \frac{1}{Z} f(z_0) \exp\left(-\frac{1}{2} A (z - z_0)^2\right)$$

$$q(z) = \left(\frac{A}{2\pi}\right)^{1/2} \exp\left(-\frac{1}{2} A (z - z_0)^2\right)$$

$$\Rightarrow Z \approx f(z_0) \left(\frac{A}{2\pi}\right)^{-1/2}$$

n-dim case

$$p(z) = \frac{1}{Z} f(z), \quad z_0 \text{ is a mode of } f$$

$$\log f(z) \approx \log f(z_0) - \frac{1}{2} (z - z_0)^T A (z - z_0), \quad A = -\nabla \nabla \log f(z_0)$$

$$\Rightarrow q(z) = \mathcal{N}(z | z_0, A^{-1}) \Rightarrow Z = \frac{(2\pi)^{n/2}}{|A|^{1/2}} f(z_0)$$

$\mathcal{D}$, M model, depending on θ , $p(\theta)$ PRIOR

$$p(\theta | \mathcal{D}) = \frac{p(\mathcal{D} | \theta) \cdot p(\theta)}{p(\mathcal{D})}$$

* computable

← this is hard

$$p(\mathcal{D}) = \int p(\mathcal{D} | \theta) p(\theta) d\theta$$

$$f(\theta) = p(\mathcal{D} | \theta) \cdot p(\theta) \quad Z = p(\mathcal{D}) \Rightarrow \text{we can use Laplace.}$$

1) θ_{MP}

$$M = |\theta|$$

$$\Rightarrow \log p(\mathcal{D}) \approx \log p(\mathcal{D} | \theta_{\text{MP}}) + \log p(\theta_{\text{MP}}) + \frac{M}{2} \log(2\pi) - \frac{1}{2} \log |A|$$

$$A = -\nabla \nabla^T [\log p(\mathcal{D} | \theta_{\text{MP}}) + \log p(\theta_{\text{MP}})]$$

BA YESIAN INFORMATION CONTENT BIC

$$\log p(\mathcal{D}) \approx \log p(\mathcal{D} | \theta_{\text{MP}}) - \frac{1}{2} M \log N$$

BAYESIAN LOGISTIC REGRESSIONS

$(x_n, y_n)_{n=1..N}$ DATA.

$\phi(x) = \phi_0^{(k)} \rightarrow \phi_{k-1}^{(x)}$ $\sigma(\omega^T \phi(x))$ LOGIT

$$p(\omega) = \mathcal{N}(\omega | m_0, S_0)$$

↑
PRIOR

$$p(\underline{y} | \omega, \underline{x}) = \prod_{i=1}^N s_i^{y_i} (1-s_i)^{1-y_i}$$

$$p(\omega | \underline{y}, \underline{x}) = \frac{p(\underline{y} | \omega, \underline{x}) \cdot p(\omega)}{p(\underline{y} | \underline{x})} \propto p(\underline{y} | \omega, \underline{x}) p(\omega)$$

$$s_i = \sigma(\omega^T \phi(x_i))$$

$s_i = s_i(\omega)$

$$\log p(\omega | \underline{y}, \underline{x}) = \underbrace{-\frac{1}{2} (\omega - m_0)^T S_0^{-1} (\omega - m_0)}_{\propto \log p(\omega)} + \underbrace{\sum_{i=1}^N \left[y_i \cdot s_i(\omega) + (1-y_i)(1-s_i(\omega)) \right]}_{\log p(\underline{y} | \omega)} + C$$

↑
NOT ANALYTICALLY TRACTABLE

LAPLACE APPROXIMATION $\Rightarrow p(\omega) = \frac{1}{Z} f(\omega)$, $f(\omega) = p(\omega) p(\underline{y} | \omega)$

1) FIND $\omega_{MAP} = \arg \max_{\omega} \log p(\omega | \underline{y}) = \arg \max_{\omega} \log p(\omega) + \log p(\underline{y} | \omega)$

2) COMPUTE COVARIANCE AT ω_{MAP} . $S_N^{-1} = S_0^{-1} + \sum_{n=1}^N s_n(\omega_{MAP})(1-s_n(\omega_{MAP})) \phi(x_n) \phi^T(x_n)$

$$p(\omega | \underline{y}) \approx q(\omega) \quad \underbrace{q(\omega) = \mathcal{N}(\omega | \omega_{\text{map}}, S_N)}$$

PREDICTIVE DISTRIBUTION

$$p(C_1 | x^*, \underline{y}, \underline{x}) = \int p(C_1 | x^*, \omega, \underline{x}, \underline{y}) p(\omega | \underline{y}, \underline{x}) d\omega$$

$$\approx \int \sigma(\underbrace{\omega^T \phi(x^*)}) q(\omega) d\omega$$

M-tim integral!

$a = \omega^T \phi(x^*)$ is
1-D Gaussian

$$q(a) = \mathcal{N}(a | \mu_a, \sigma_a^2)$$

$$\left. \begin{aligned} \mu_a &= \omega_{\text{map}}^T \phi(x^*) \\ \sigma_a^2 &= \phi^T(x^*) S_N \phi(x^*) \end{aligned} \right\}$$

$$\Rightarrow p(C_1 | x^*, \underline{y}, \underline{x}) = \int \sigma(a) q(a) da$$

PROBIT APPROXIMATION $\sigma(a) \approx \Psi(\lambda a)$

$$\lambda^2 = \frac{\pi}{8} \text{ MATCHING } \sigma'(0) = \Psi'(0)$$

$$\int \Psi(\lambda a) q(a) da = \Psi\left(\frac{\mu_a}{(\lambda^2 + \sigma_a^2)^{1/2}}\right) \Rightarrow \boxed{p(C_1 | x^*, \underline{y}, \underline{x}) \approx \sigma(K(\sigma_a^2) \mu_a)}$$

$$\boxed{K(\sigma_a^2) = (1 + \pi \sigma_a^2 / 8)^{-1/2}}$$