

Parametric analysis (II part)

Data description: The “ovarian” dataset contains information on survival in a randomized trial comparing two treatments group for ovarian cancer

1. Working on the pooled sample:

```
. cd "working directory path"
. use ovarian.dta,clear
. des
```

Contains data from ovarian.dta

```
obs:          26
vars:          7                      7 May 2015 15:22
size:          390 (99.9% of memory free)
```

variable name	storage type	display format	value label	variable label
id	byte	%8.0g		identity number
futime	int	%8.0g		survival or censoring time
fustat	byte	%8.0g		censoring status
age	float	%9.0g		age (in years)
residdds	byte	%8.0g		residual disease present (1=no, 2=yes)
rx	byte	%8.0g		treatment group
ecogps	byte	%8.0g		ECOG performance status (1 if better)

```
. recode rx (2=0)
```

2. **Preliminary analysis** Describe data looking at frequency distributions of variables and at thier cross-tabulations

```
. tab rx
```

treatment group	Freq.	Percent	Cum.
1	13	50.00	50.00
2	13	50.00	100.00
Total	26	100.00	

```
. tab residdds
```

residual disease present (1=no, 2=yes)	Freq.	Percent	Cum.
1	11	42.31	42.31
2	15	57.69	100.00
Total	26	100.00	

```
. tab ecogps
      ECOG |
performance |
      status (1 |
if better) |      Freq.      Percent      Cum.
-----+-----
          1 |          14          53.85          53.85
          2 |          12          46.15         100.00
-----+-----
        Total |          26         100.00
```

```
. sum age,d
              age (in years)
-----+-----
      Percentiles      Smallest
  1%          38.8932          38.8932
  5%          39.2712          39.2712
 10%          43.1233          43.1233      Obs          26
 25%          50.1096          43.137      Sum of Wgt.      26

 50%          56.8466                      Mean          56.16544
                      Largest          Std. Dev.          10.10036
 75%          63.2192          66.4658
 90%          72.3315          72.3315      Variance          102.0173
 95%          74.4932          74.4932      Skewness           .0434026
 99%          74.5041          74.5041      Kurtosis           2.343786
```

```
. ttest age,by(rx)
Two-sample t test with equal variances
-----+-----
      Group |      Obs      Mean      Std. Err.      Std. Dev.      [95% Conf. Interval]
-----+-----
          1 |       13      55.73235      3.749421      13.51873      47.56307      63.90164
          2 |       13      56.59853      1.503165      5.419739      53.32342      59.87365
-----+-----
combined |       26      56.16544      1.980844      10.10036      52.08582      60.24507
-----+-----
      diff |          -.8661781      4.039513          -9.203323      7.470966
-----+-----

      diff = mean(1) - mean(2)                      t = -0.2144
Ho: diff = 0                      degrees of freedom =          24

      Ha: diff < 0                      Ha: diff != 0                      Ha: diff > 0
Pr(T < t) = 0.4160          Pr(|T| > |t|) = 0.8320          Pr(T > t) = 0.5840
```

```
. tab residdds rx,col row chi2
residual |
disease |
```

present	treatment group		
(1=no,	1	2	Total
2=yes)			
-----+-----+-----			
1	5	6	11
	45.45	54.55	100.00
	38.46	46.15	42.31
-----+-----+-----			
2	8	7	15
	53.33	46.67	100.00
	61.54	53.85	57.69
-----+-----+-----			
Total	13	13	26
	50.00	50.00	100.00
	100.00	100.00	100.00

Pearson chi2(1) = 0.1576 Pr = 0.691

```
. tab ecogps rx, col row chi2
```

ECOG	treatment group		
performanc	1	2	Total
e status			
(1 if			
better)			
-----+-----+-----			
1	7	7	14
	50.00	50.00	100.00
	53.85	53.85	53.85
-----+-----+-----			
2	6	6	12
	50.00	50.00	100.00
	46.15	46.15	46.15
-----+-----+-----			
Total	13	13	26
	50.00	50.00	100.00
	100.00	100.00	100.00

Pearson chi2(1) = 0.0000 Pr = 1.000

```
. tab residdds ecogps, col row chi2
```

residual	ECOG performance		
disease	status (1 if better)		
present	1	2	Total
(1=no,			
2=yes)			
-----+-----+-----			
1	5	6	11
	45.45	54.55	100.00
	35.71	50.00	42.31

2	9	6	15
	60.00	40.00	100.00
	64.29	50.00	57.69
Total	14	12	26
	53.85	46.15	100.00
	100.00	100.00	100.00

Pearson chi2(1) = 0.5403 Pr = 0.462

. ttest age,by(resids)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
1	11	52.7721	2.280056	7.562091	47.69182	57.85238
2	15	58.65389	2.894191	11.20915	52.44647	64.86132
combined	26	56.16544	1.980844	10.10036	52.08582	60.24507
diff		-5.881793	3.912005		-13.95577	2.192188

diff = mean(1) - mean(2)

t = -1.5035

Ho: diff = 0

degrees of freedom = 24

Ha: diff < 0

Ha: diff != 0

Ha: diff > 0

Pr(T < t) = 0.0729

Pr(|T| > |t|) = 0.1458

Pr(T > t) = 0.9271

. ttest age,by(ecogps)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
1	14	54.98004	2.900216	10.85161	48.71451	61.24558
2	12	57.54841	2.720518	9.424152	51.56059	63.53623
combined	26	56.16544	1.980844	10.10036	52.08582	60.24507
diff		-2.568365	4.021366		-10.86806	5.731326

diff = mean(1) - mean(2)

t = -0.6387

Ho: diff = 0

degrees of freedom = 24

Ha: diff < 0

Ha: diff != 0

Ha: diff > 0

Pr(T < t) = 0.2645

Pr(|T| > |t|) = 0.5291

Pr(T > t) = 0.7355

3. Declare data to be survival-data in order to take into account the follow-up

. stset futime,fail(fustat)id(id)

```

            id: id
        failure event: fustat != 0 & fustat < .
obs. time interval: (fuptime[_n-1], fuptime]
exit on or before: failure

```

```

-----
26 total obs.
0 exclusions
-----

```

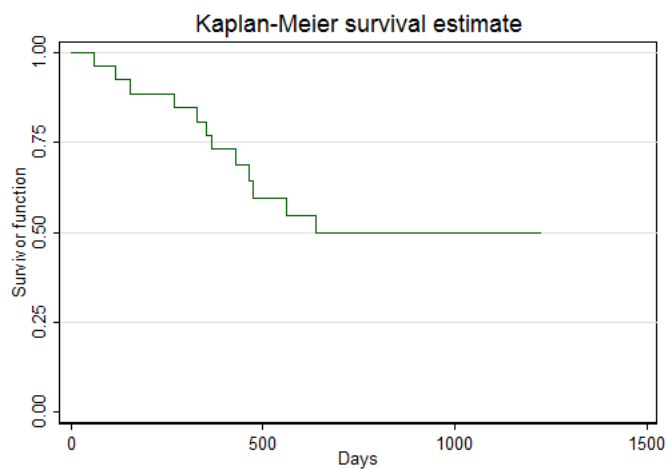
```

26 obs. remaining, representing
26 subjects
12 failures in single failure-per-subject data
15588 total analysis time at risk, at risk from t = 0
        earliest observed entry t = 0
        last observed exit t = 1227

```

4. By using non-parametric methods represent the total survival curve and the survival curves by variables by Kaplan-Meier estimator and test their difference

```
. sts graph
```



```
. sts graph,by(rx)
```

```
. sts test rx
```

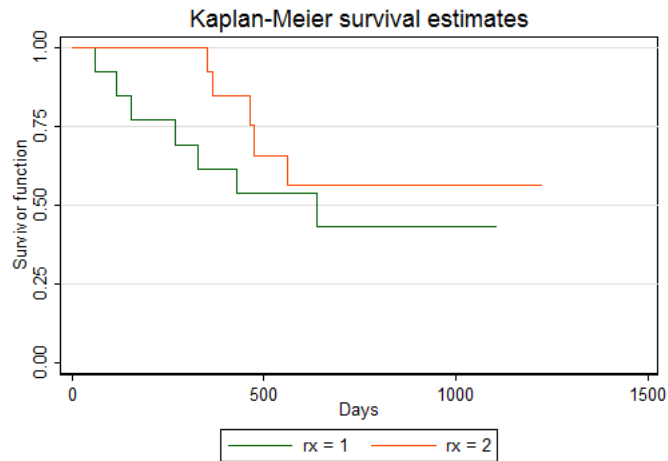
Log-rank test for equality of survivor functions

rx	Events	
	observed	expected
1	7	5.23
2	5	6.77

Total | 12 12.00

chi2(1) = 1.06

Pr>chi2 = 0.3026



```
. sts graph,by(residds)
```

```
. sts test residds
```

Log-rank test for equality of survivor functions

	Events	Events
residds	observed	expected
1	3	6.26
2	9	5.74
Total	12	12.00

chi2(1) = 3.62

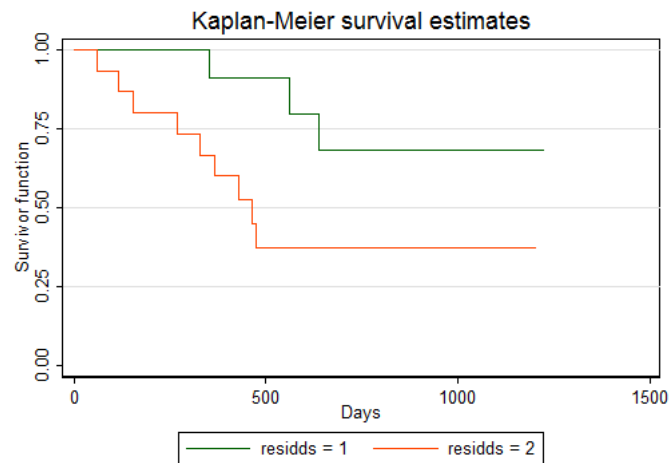
Pr>chi2 = 0.0570

```
. sts graph,by(ecogps)
```

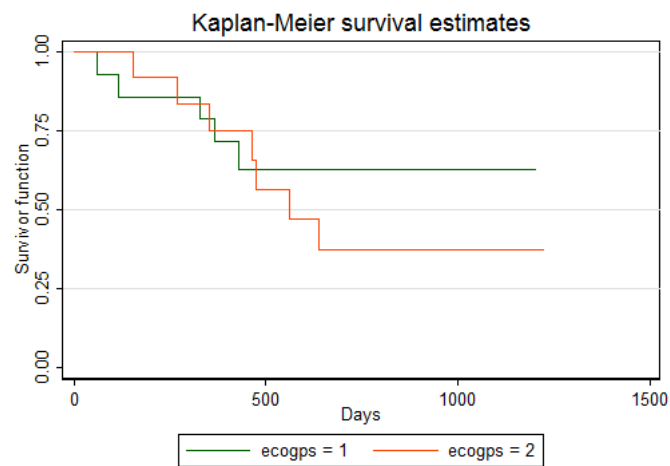
```
. sts test ecogps
```

Log-rank test for equality of survivor functions

	Events	Events
ecogps	observed	expected
1	5	6.18
2	7	5.82
Total	12	12.00



```
chi2(1) =      0.47
Pr>chi2 =      0.4941
```

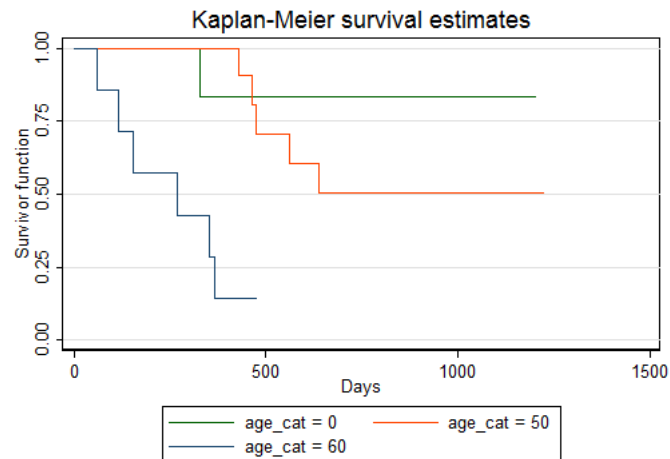


```
. egen age_cat=cut(age),at(0,50(10)60,100)
. sts graph,by(age_cat)
. sts test age_cat
Log-rank test for equality of survivor functions
```

	Events	Events
age_cat	observed	expected
0	1	3.60
50	5	6.93
60	6	1.47

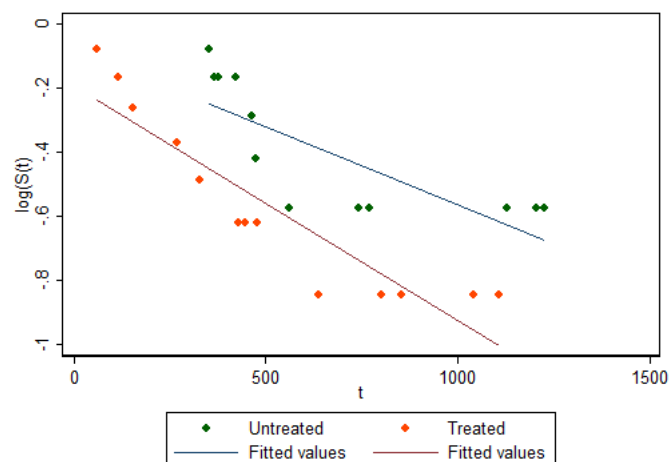
```
-----+-----
Total   |          12          12.00
```

```
chi2(2) =      17.68
Pr>chi2 =      0.0001
```



5. Before fitting the regression models, check for the assumptions underlying the exponential, weibull, log-logistic and log-normal models respectively by graphical methods

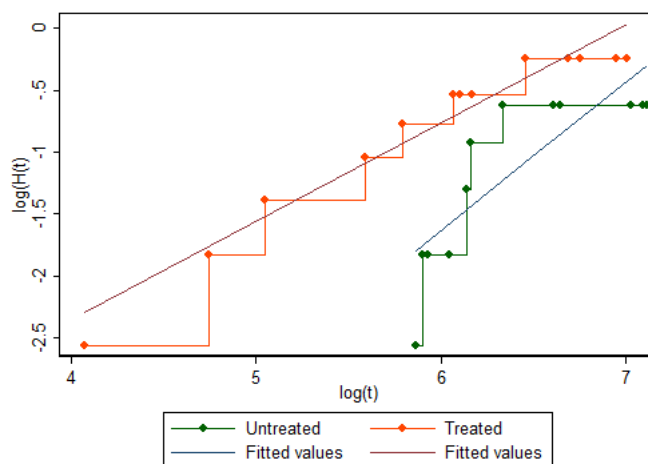
```
. *exponential model
. *survival function
. sts gen km= s,by(rx)
. gen logkm_exp=log(km)
. twoway (scatter logkm_exp _t if rx==0,sort)(scatter logkm_exp _t if rx==1,sort) ///
> (lfit logkm_exp _t if rx==0,sort) (lfit logkm_exp _t if rx==1,sort), ///
> legend(lab(1 "Untreated") lab(2 "Treated")) ytitle("log(S(t))" xtitle("t")
```



```

. *cumulative hazard function
. sts gen cumhaz_exp=na,by(rx)
. gen cumhaz_exp1=cumhaz_exp
. gen cumhaz_exp0=cumhaz_exp
. gen logcumhaz_exp=log(cumhaz_exp)
. gen logtime=log(_t)
. twoway (scatter logcumhaz_exp logtime if rx==0,connect(J)sort) ///
> (scatter logcumhaz_exp logtime if rx==1,connect(J)sort) ///
> (lfit logcumhaz_exp logtime if rx==0) (lfit logcumhaz_exp logtime if rx==1), ///
> legend(lab(1 "Untreated") lab(2 "Treated")) ytitle("log(H(t)") xtitle("log(t)")

```



```

. *observed vs predicted
. xi:streg rx,dist(e) nohr

```

	_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	rx	.6125256	.58554	1.05	0.296	-.5351118	1.760163
	_cons	-7.480203	.4472136	-16.73	0.000	-8.356725	-6.60368

```

*observed vs predicted
. xi:streg rx,dist(e) nohr
. predict surv_exp,surv
. twoway (line km _t if rx==0,sort)(line km _t if rx==1,sort) ///
> (scatter surv_exp _t if rx==0,sort) (scatter surv_exp _t if rx==1,sort), ///
> legend(lab(1 "Untreated observed") lab(2 "Treated observed") lab(3 "Untreated predicted")
> lab(4 "Treated predicted")) ytitle("log(S(t)") xtitle("_t")

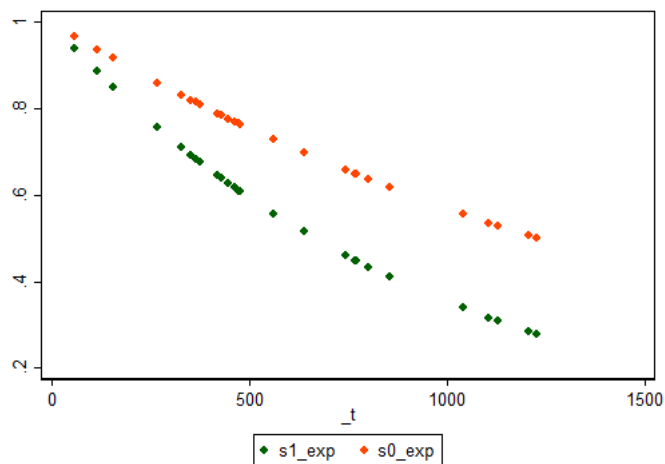
```

```

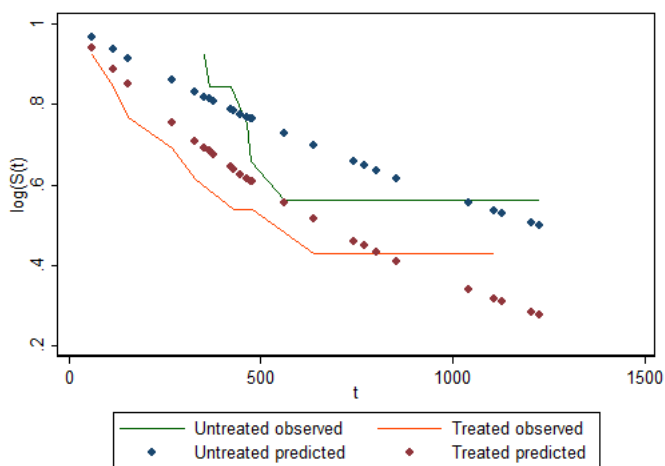
*alternatively
. gen s1_exp=exp(-exp(_b[_cons]+_b[rx])*_t)
. gen s0_exp=exp(-exp(_b[_cons])*_t)
. sort _t
. twoway (scatter s1_exp _t) (scatter s0_exp _t)

```

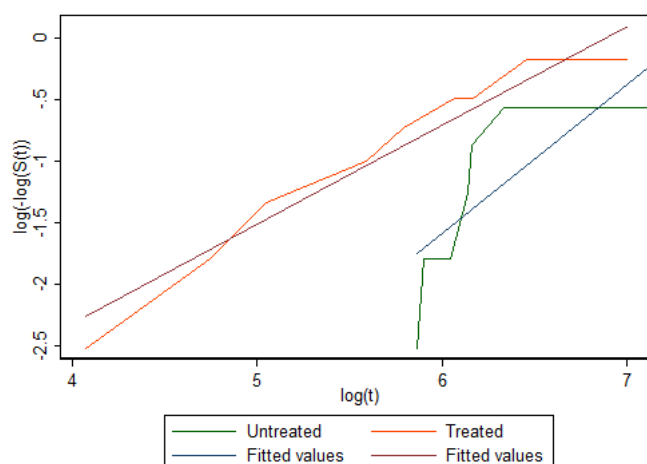
```
. stcurve, surv at(rx=0) at1(rx=1)
. twoway (line km _t if rx==0,sort)(line km _t if rx==1,sort) ///
> (scatter s0_exp _t,sort) (scatter s1_exp _t,sort), ///
> legend(lab(1 "Untreated observed") lab(2 "Treated observed") lab(3 "Untreated predicted")
```



```
. twoway (line km _t if rx==0,sort)(line km _t if rx==1,sort) ///
> (scatter s0_exp _t,sort) (scatter s1_exp _t,sort), legend(lab(1 "Untreated observed") ///
> lab(2 "Treated observed") lab(3 "Untreated predicted") lab(4 "Treated predicted")) ytitle
```



```
. *weibull model
. *survival function
. gen logkm_w=log(-log(km))
. gen logt=log(_t)
. twoway (line logkm_w logt if rx==0,sort)(line logkm_w logt if rx==1,sort) ///
> (lfit logkm_w logt if rx==0,sort) (lfit logkm_w logt if rx==1,sort), ///
> legend(lab(1 "Untreated") lab(2 "Treated")) ytitle("log(-log(S(t)))" xtitle("log(t)")
```



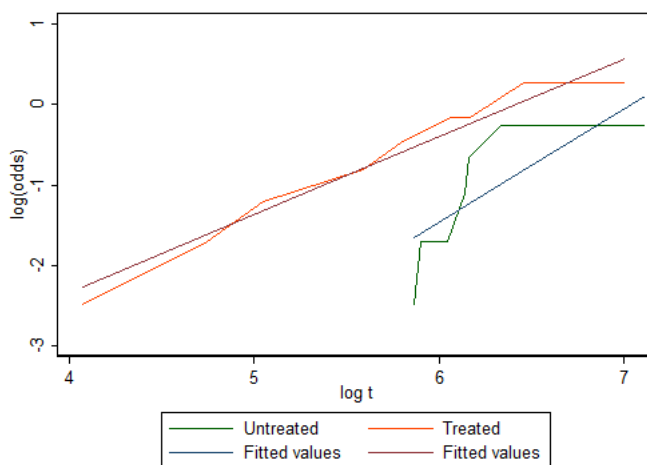
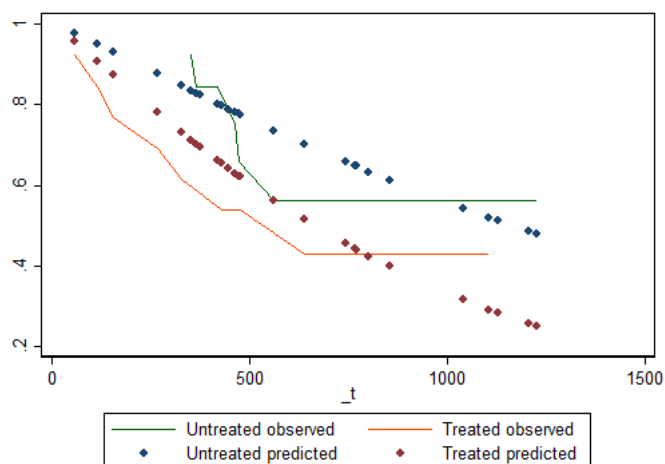
```
. *observed vs predicted
. xi:streg rx,dist(w) nohr
```

	_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
rx		.6314782	.5868807	1.08	0.282	-.518787	1.781743
_cons		-8.337065	1.940367	-4.30	0.000	-12.14011	-4.534017
/ln_p		.1214247	.2514893	0.48	0.629	-.3714853	.6143346
p		1.129104	.2839576			.6897092	1.848426
1/p		.8856578	.2227334			.5410008	1.449886

```
. gen s_w0=exp(-exp(_b[_cons])*_t^1.129104)
. gen s_w1=exp(-exp(_b[_cons]+_b[rx])*_t^1.129104)
. twoway (scatter s_w0 _t) (scatter s_w1 _t)
. stcurve, surv at(rx=0) at1(rx=1)
. twoway (line km _t if rx==0,sort)(line km _t if rx==1,sort) ///
> (scatter s_w0 _t,sort) (scatter s_w1 _t,sort), ///
> legend(lab(1 "Untreated observed") lab(2 "Treated observed") ///
> lab(3 "Untreated predicted") lab(4 "Treated predicted"))

. *log-logistic model
. gen odds_km=(1-km)/km
. gen logodds=log(odds_km)
. twoway (line logodds logt if rx==0,sort)(line logodds logt if rx==1,sort) ///
> (lfit logodds logt if rx==0,sort) (lfit logodds logt if rx==1,sort), ///
> legend(lab(1 "Untreated") lab(2 "Treated")) ytitle("log(odds)") xtitle("log t")
```

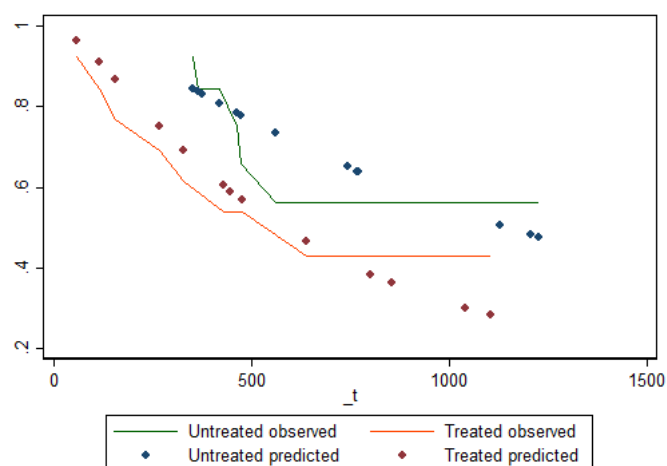
```
. *observed vs predicted
. xi:streg rx,dist(loglog)
```



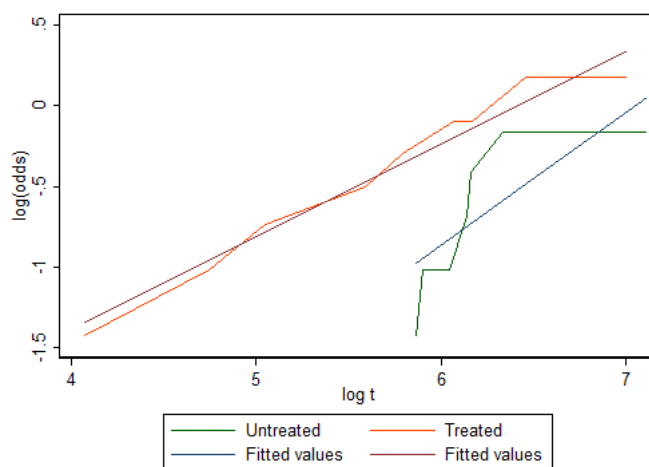
_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
rx	-.6851239	.5365499	-1.28	0.202	-1.736742	.3664945
_cons	7.046231	.4118578	17.11	0.000	6.239005	7.853458
/ln_gam	-.3594936	.2466853	-1.46	0.145	-.8429879	.1240007
gamma	.6980297	.1721937			.4304225	1.132017

```
. predict surv_loglog,surv
. twoway (line km _t if rx==0,sort)(line km _t if rx==1,sort) ///
> (scatter surv_loglog _t if rx==0,sort) (scatter surv_loglog _t if rx==1,sort), ///
> legend(lab(1 "Untreated observed") lab(2 "Treated observed") ///
lab(3 "Untreated predicted") lab(4 "Treated predicted"))
```

```
. *log-normal model
```



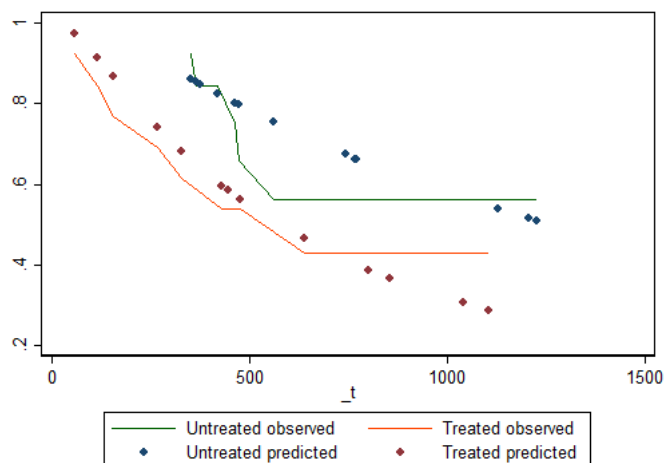
```
. gen inv_km=invnormal(1-km)
. twoway (line inv_km logt if rx==0,sort)(line inv_km logt if rx==1,sort) ///
> (lfit inv_km logt if rx==0,sort) (lfit inv_km logt if rx==1,sort), ///
> legend(lab(1 "Untreated") lab(2 "Treated")) ytitle("log(odds)") xtitle("log t")
```



```
. *observed vs predicted
. xi:streg rx,dist(ln)
```

_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
rx	-.7864423	.5269101	-1.49	0.136	-1.819167	.2462826
_cons	7.140936	.4239738	16.84	0.000	6.309963	7.971909
/ln_sig	.1600969	.2267412	0.71	0.480	-.2843076	.6045015
sigma	1.173625	.266109			.7525351	1.830339

```
-----
. predict surv_ln,surv
```



6. By using the exponential model evaluate the treatment effect (crude and adjusted estimates) on mortality

```
. xi:streg i.rx,dist(e)
Log likelihood = -29.288092          Prob > chi2      = 0.2910
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
_Irx_1	1.845086	1.080371	1.05	0.296	.5856038	5.813385

```
. xi:streg i.rx age,dist(e)
Log likelihood = -23.022327          Prob > chi2      = 0.0011
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
_Irx_1	1.93685	1.192845	1.07	0.283	.5792526	6.476258
age	1.110781	.0354083	3.30	0.001	1.043506	1.182394

```
. xi:streg i.rx i.resids age i.ecogps,dist(e)
Log likelihood = -22.457651          Prob > chi2      = 0.0052
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
_Irx_1	1.87184	1.153513	1.02	0.309	.5593961	6.263511
_Iresids_2	2.15096	1.594008	1.03	0.301	.5033083	9.192438

age	1.091417	.0368664	2.59	0.010	1.0215	1.166119
_Iecogps_2	1.287025	.7800968	0.42	0.677	.3923316	4.222024

```
. xi:streg i.rx*i.ecogps i.residdds age,dist(e)
Log likelihood = -21.612988 Prob > chi2 = 0.0056
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
_Irx_1	5.318805	6.053983	1.47	0.142	.5714224 49.50748
_Iecogps_2	3.638872	4.113217	1.14	0.253	.3970149 33.35238
_IrxXeco_1_2	.1914859	.2593592	-1.22	0.222	.0134655 2.723027
_Iresiddds_2	2.135929	1.589535	1.02	0.308	.4967432 9.184209
age	1.090369	.037685	2.50	0.012	1.018954 1.16679

```
. xi:streg i.rx*i.residdds age i.ecogps,dist(e)
Log likelihood = -22.321705 Prob > chi2 = 0.0102
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
_Irx_1	1.085295	1.355736	0.07	0.948	.0938076 12.55619
_Iresiddds_2	1.595025	1.460327	0.51	0.610	.265124 9.595908
_IrxXres_1_2	2.179131	3.296521	0.51	0.607	.1123602 42.26241
age	1.083755	.0372925	2.34	0.019	1.013074 1.159369
_Iecogps_2	1.339869	.8126651	0.48	0.630	.4081191 4.398838

```
. xi:streg i.rx i.residdds age i.ecogps,dist(e)
Log likelihood = -22.457651 Prob > chi2 = 0.0052
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
_Irx_1	1.87184	1.153513	1.02	0.309	.5593961 6.263511
_Iresiddds_2	2.15096	1.594008	1.03	0.301	.5033083 9.192438
age	1.091417	.0368664	2.59	0.010	1.0215 1.166119
_Iecogps_2	1.287025	.7800968	0.42	0.677	.3923316 4.222024

```
. est store a
. xi:streg i.rx i.residdds i.age_cat i.ecogps,dist(e)
Log likelihood = -22.151354 Prob > chi2 = 0.0088
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
_Irx_1	1.866375	1.312079	0.89	0.375	.4705396 7.402894

```

    _Iresidss_2 |      2.3475    1.724604      1.16    0.245      .5562446    9.907075
    _Iage_cat_50 |     4.503212    5.283785      1.28    0.200      .4516232    44.9023
    _Iage_cat_60 |     14.74248     16.171      2.45    0.014      1.717469    126.5471
    _Iecogps_2 |     1.773817    1.067332      0.95    0.341      .5454163    5.768852

```

```

-----
. est store b
. lrtest a b
Likelihood-ratio test                                LR chi2(1) =      0.61
(Assumption: a nested in b)                        Prob > chi2 =      0.4338
. xi:streg i.rx i.residss age i.ecogps,dist(e)
Log likelihood = -22.457651                          Prob > chi2      =      0.0052

```

```

-----
      _t | Haz. Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
      _Irx_1 |      1.87184    1.153513     1.02   0.309     .5593961    6.263511
    _Iresidss_2 |      2.15096    1.594008     1.03   0.301     .5033083    9.192438
      age |      1.091417   .0368664     2.59   0.010     1.0215     1.166119
    _Iecogps_2 |      1.287025   .7800968     0.42   0.677     .3923316    4.222024

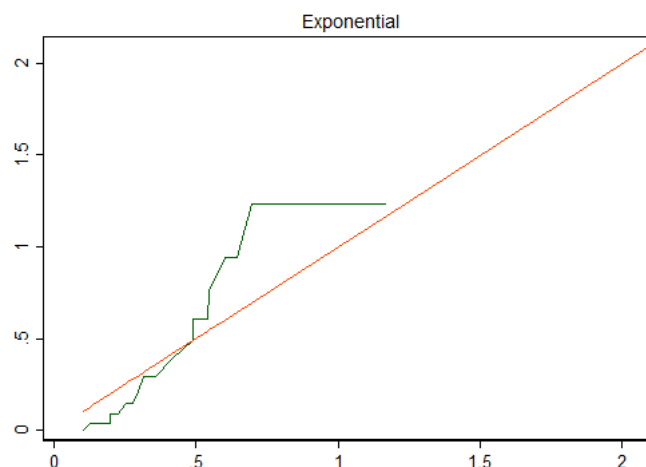
```

7. Calculate Cox-Snell residuals for the exponential model (post-estimation)

```

. predict residual, csnell
. stset residual, failure(fustat)id(id)
. sts gen km = s
. generate cumhaz = -ln(km)
. line cumhaz residual residual, sort name("exponential, replace") ///
> legend(off) t1(Exponential) xtitle("")
. drop cumhaz residual k

```



8. Calculate standardized residuals for the exponential model (post-estimation)

```

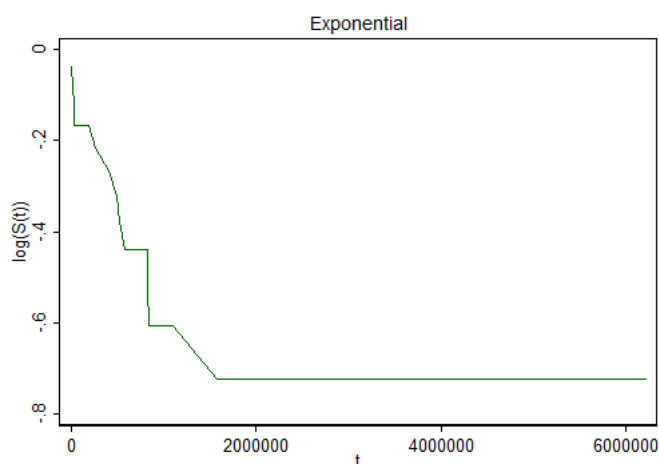
. stset futime,fail(fustat)id(id)

```

```

. xi:streg i.rx i.resids age i.ecogps,dist(e)
. predict pred, xb
. gen res_stand=(log(_t)-pred)
. replace res_stand=exp(res_stand)
. stset res_stand, failure(fustat)id(id)
. sts gen km = s
. gen logkm=log(km)
. line logkm _t,sort legend(off) t1(Exponential) xtitle("t") ytitle("log(S(t))")
. drop km logkm res_stand pred

```



9. Calculate deviance residuals for the exponential model (post-estimation)

```

. stset futime,fail(fustat)id(id)
. xi:streg i.rx i.resids age i.ecogps,dist(e)
. predict deviance, deviance
. graph twoway scatter deviance id, legend(off) t1(Exponential) xtitle("id")
. drop deviance

```

10. After fitting the Weibull model, calculate Cox-Snell residuals (post-estimation)

```

. stset futime,fail(fustat)id(id)
. xi:streg i.rx i.resids age i.ecogps,dist(w)
. predict residual, csnell
. stset residual, failure(fustat)id(id)
. sts gen km = s
. generate cumhaz = -ln(km)
. line cumhaz residual residual, sort name("Weibull, replace") ///
> legend(off) t1(Weibull) xtitle("")
. drop cumhaz residual k

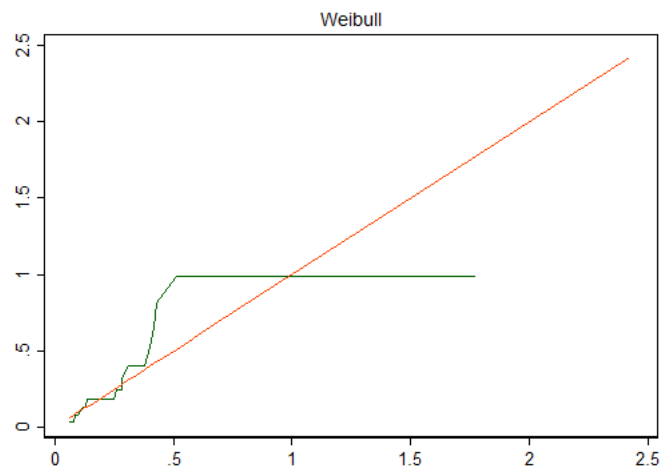
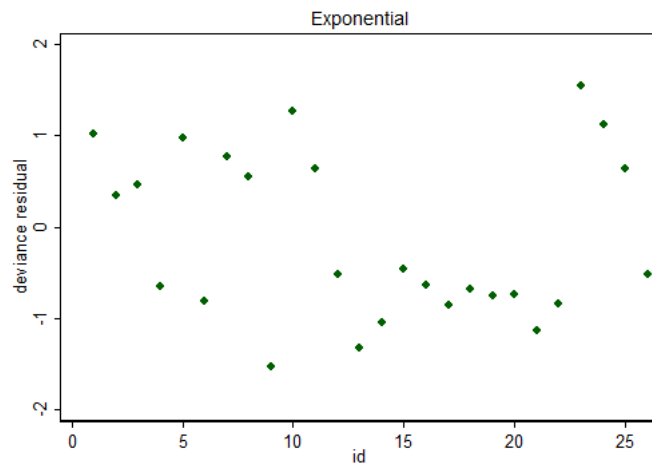
```

11. Calculate standardized residuals for the weibull model (post-estimation)

```

. stset futime,fail(fustat)id(id)
. xi:streg i.rx i.resids age i.ecogps,dist(w)

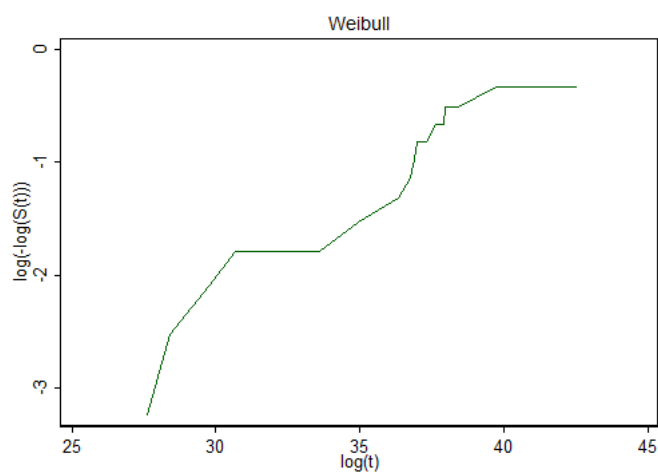
```



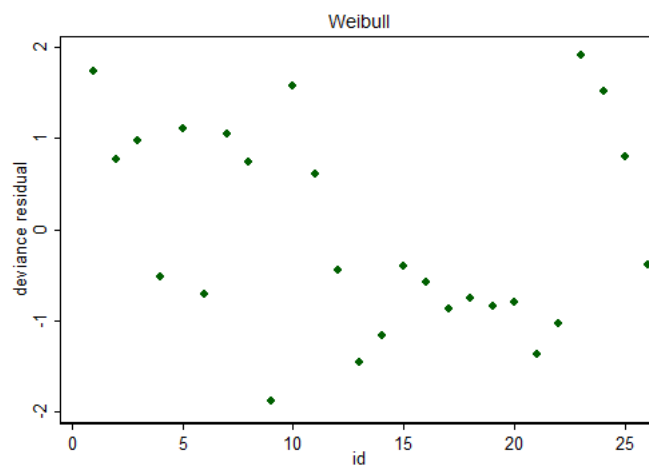
```
. predict pred, xb
. gen sigma= 0.5195874
. gen res_stand=(log(_t)-pred)/sigma
. replace res_stand=exp(res_stand)
. stset res_stand, failure(fustat)id(id)
. sts gen km = s
. gen gen tras_km=log(-log(km))
. gen log_time=log(_t)
. line tras_km log_time ,sort legend(off) t1(Weibull) xtitle("log(t)") ///
> ytitle("log(-log(S(t)))")
. drop km tras_km res_stand sigma pred log_time
```

12. Calculate deviance residuals for the weibull model (post-estimation)

```
. stset futime,fail(fustat)id(id)
. xi:streg i.rx i.residdds age i.ecogps,dist(w)
. predict deviance, deviance
```



```
. graph twoway scatter deviance id, legend(off) t1(Weibull) xtitle("id")
. drop deviance
```

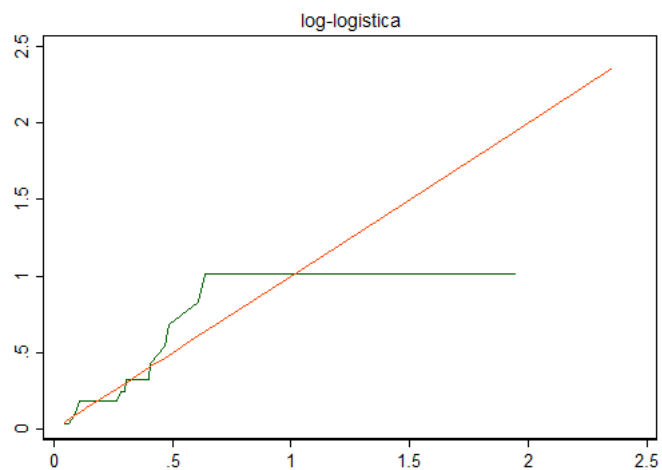


13. After fitting the log-logistic model, calculate Cox-Snell residuals (post-estimation)

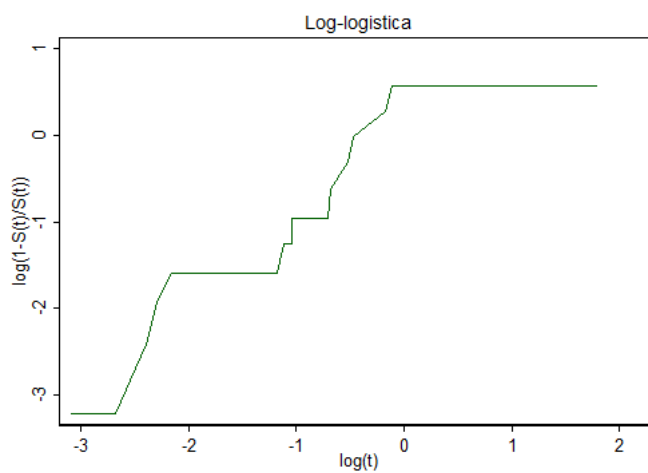
```
. stset futime, fail(fustat) id(id)
. xi:streg i.rx i.resids age i.ecogps, dist(loglog)
. predict residual, csnell
. stset residual, failure(fustat) id(id)
. sts gen km = s
. line cumhaz residual residual, sort name("loglogistica", replace) ///
. legend(off) t1(log-logistica) xtitle("")
. drop cumhaz residual km
```

14. Calculate standardized residuals for the log-logistic model (post-estimation)

```
. stset futime, fail(fustat) id(id)
```



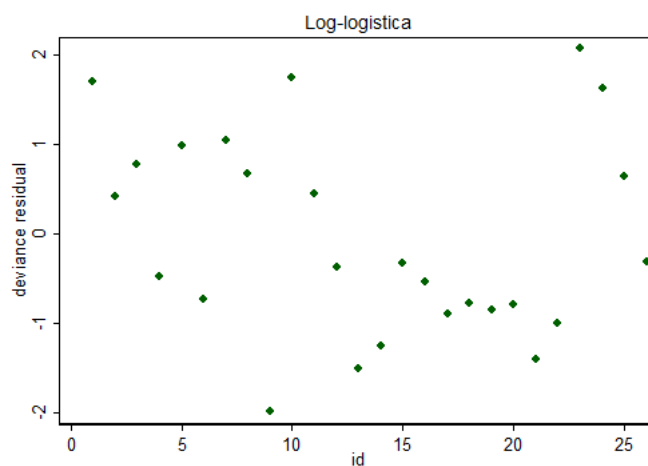
```
. xi:streg i.rx i.resids age i.ecogps,dist(loglog)
. predict pred, xb
. gen sigma=.3727281
. gen res_stand=(log(_t)-pred)/sigma
. replace res_stand=exp(res_stand)
. stset res_stand, failure(fustat)id(id)
. sts gen km = s
. gen tras_km=log((1-km)/km)
. gen log_time=log(_t)
. line tras_km log_time,sort legend(off) t1(Log-logistica) xtitle("log(t)") ///
> ytitle("log(1-S(t)/S(t))")
. drop km tras_km res_stand sigma pred log_time
```



15. Calculate deviance residuals for the log-logistic model (post-estimation)

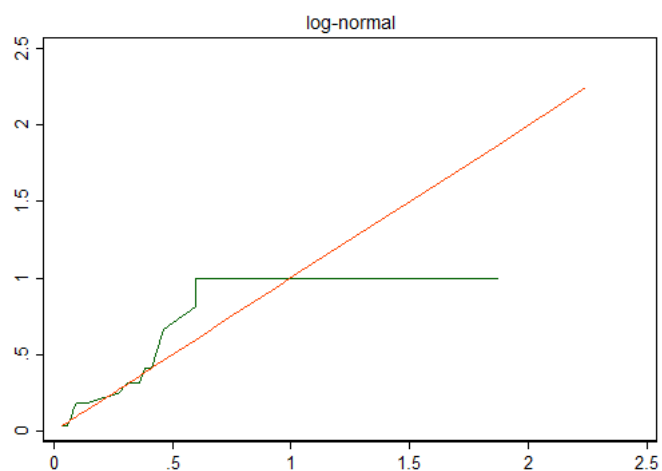
```
. stset futime,fail(fustat)id(id)
. xi:streg i.rx i.resids age i.ecogps,dist(loglog)
```

```
. predict deviance, deviance
. graph twoway scatter deviance id, legend(off) t1(Log-logistica) xtitle("id")
. drop deviance
```



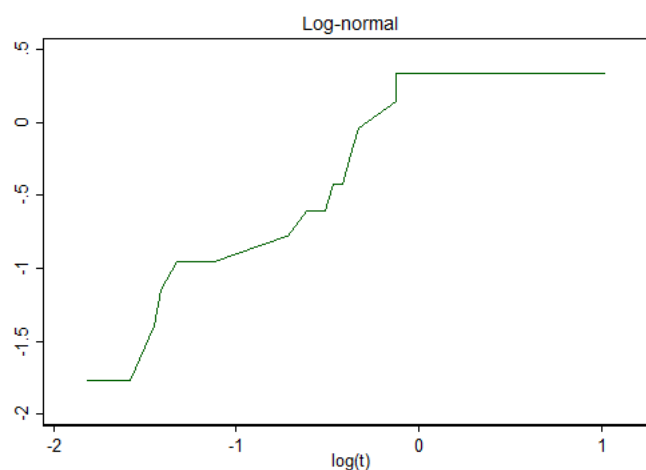
16. After fitting the log-normal model, calculate Cox-Snell residuals (post-estimation)

```
. stset futime, fail(fustat) id(id)
. xi:streg i.rx i.residss age i.ecogps, dist(ln)
. predict residual, csnell
. stset residual, fail(fustat) id(id)
. sts gen km = s
. generate cumhaz = -ln(km)
. line cumhaz residual residual, sort name("lognormale, replace") ///
. legend(off) t1(log-normal) xtitle("")
. drop cumhaz residual km
```



17. Calculate standardized residuals for the log-normal model (post-estimation)

```
.stset futime,fail(fustat)id(id)
.xi:streg i.rx i.residss age i.ecogps,dist(ln)
.predict pred, xb
.gen sigma=.6319339
.gen res_stand=(log(_t)-pred)/sigma
.replace res_stand=exp(res_stand)
.stset res_stand, fail(fustat)id(id)
.sts gen km = s
.gen tras_km=invnormal(1-km)
.gen log_time=log(_t)
.line tras_km log_time,sort legend(off) t1(Log-normal) xtitle("log(t)") ///
.ytitle("")
.drop km tras_km res_stand sigma pred log_time
```



18. Calculate deviance residuals for the log-normal model (post-estimation)

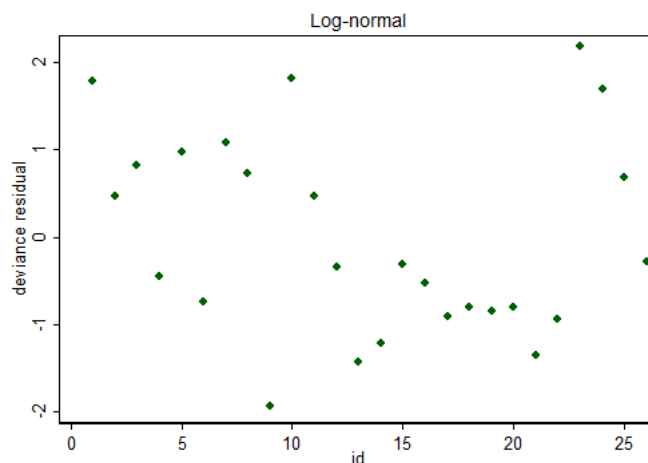
```
stset futime,fail(fustat)id(id)
.xi:streg i.rx i.residss age i.ecogps,dist(ln)
.predict deviance, deviance
.graph twoway scatter deviance id, legend(off) t1(Log-normal) xtitle("id")
.drop deviance
```

19. Compare the models by AIC (Akaike Criterion Information= $2k-2\ln(L)$)

```
stset futime,fail(fustat)id(id)
.xi:streg i.rx i.residss age i.ecogps,dist(e)
.estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
-----+-----						
.	26	-29.84554	-22.45765	5	54.9153	61.20579

```
xi:streg i.rx i.residss age i.ecogps,dist(w)
.estat ic
```



Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	26	-29.76724	-19.66043	6	51.32085	58.86943

```
xi:streg i.rx i.residss age i.ecogps,dist(loglog)
estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	26	-29.16804	-18.63665	6	49.2733	56.82188

```
xi:streg i.rx i.residss age i.ecogps,dist(ln)
estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	26	-28.93508	-18.22165	6	48.4433	55.99188

20. After selecting the best model fitting data, interpret the estimates obtained by that model

```
xi:streg i.rx i.residss age i.ecogps,dist(ln)
```

```
Log likelihood = -18.22165 Prob > chi2 = 0.0003
```

_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
_Irx_1	-.6058923	.3026846	-2.00	0.045	-1.199143 -.0126413
_Iresidss_2	-.5331212	.3493213	-1.53	0.127	-1.217778 .151536
age	-.0676027	.0162841	-4.15	0.000	-.0995189 -.0356865
_Iecogps_2	-.0205032	.3203281	-0.06	0.949	-.6483347 .6073282
_cons	10.99911	.9817151	11.20	0.000	9.074987 12.92324

